CMPE 58N - Lecture 1. Monte Carlo methods

Introduction to Monte Carlo method, Motivating Examples, Law of Large Numbers, Central Limit Theorem



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Outline

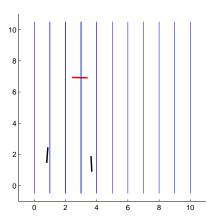
- Introduction to Monte Carlo method,
- Motivating Examples,
- ▶ Law of Large Numbers,
- Central Limit Theorem
- Example

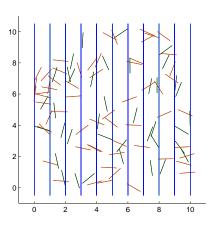
Monte Carlo Methods

- Represent the solution of a problem as a parameter of a hypothetical population,
- use a pseudo-random sequence of numbers to construct a sample of a population, from which statistical estimates of the parameter can be obtained
- Stochastic Simulation or Sampling methods

History of Monte Carlo methods

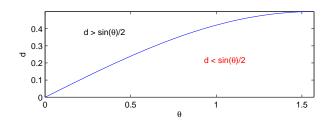
- 1733 Buffon's needle problem.
- 1812 Laplace suggests using Buffon's needle experiment to estimate π .
- 1946 ENIAC (Electronic Numerical Integrator And Computer) built.
- 1947 John von Neuman and Stanislaw Ulam propose a computer simulation to solve the problem of neutron diffusion in fissionable material.
- 1949 Metropolis and Ulam publish their results in the Journal of the American Statistical Association.
- 1984 Geman & Geman publish their paper on the Gibbs sampler ... continuously growing interest with increases in computational power





- d: Distance from the middle of the needle to the nearest line
- ightharpoonup heta : Acute angle between the parallel lines and the needle
- A needle touches a line iff

$$\frac{d}{\sin \theta} < \frac{1}{2}$$



The area of the rectangle is

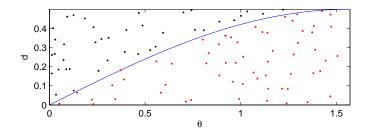
$$S = \frac{1}{2} \frac{\pi}{2}$$

▶ The area under the sin is

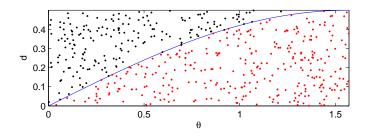
$$\int_0^{\pi/2} \sin(\theta)/2 = \frac{1}{2}$$

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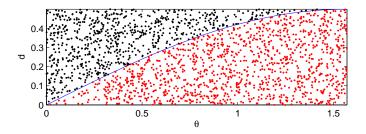
$$\Pr\{d < \sin(\theta)/2\} = \frac{1/2}{\pi/4} = \frac{2}{\pi}$$



$$\pi \approx 3.2787$$



 $\pi \approx 3.1949$



$$\pi \approx 3.1596$$

Indicator function

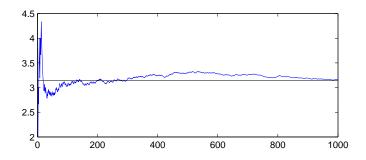
$$\mathbb{I}\{x\} = \begin{cases} 1 & x \text{ is true} \\ 0 & \text{otherwise} \end{cases}$$

Alternative notation: Iverson convention

$$[x] = \begin{cases} 1 & x \text{ is true} \\ 0 & \text{otherwise} \end{cases}$$

▶ Draw $(d^{(n)}, \theta^{(n)}) \sim U_S$ and estimate π via

$$\begin{array}{ll} \pi & = & \frac{2}{\Pr\{d < \sin(\theta)/2\}} \approx \frac{2\text{\# of all dots}}{\text{\# of red dots}} \\ & = & \frac{2N}{\sum_{n=1}^{N} \mathbb{I}\{d^{(n)} < \sin(\theta^{(n)})/2\}} \end{array}$$



Speed of convergence

- ▶ Monte Carlo integration: error behaves as $n^{-1/2}$.
- Numerical integration of a one-dimensional function by Riemann sums: error behaves as n^{-1} .
- For one-dimensional problems Riemann is better; however deteriorates with increasing dimension: curse of dimensionality.
- Order of convergence of Monte Carlo integration is independent of the dimension of the problem.
 - → Monte Carlo methods can be a good choice for high-dimensional integrals.

Convergence of random variables

(Liu, Appendix A.1.4.)

$$y_n \sim p_n(y_n)$$
 $F_n(y_n) = \int_{-\infty}^{y_n} p_n(\tau) d\tau$

1 Convergence in distribution

$$\lim_{n\to\infty} F_n(y_n) = F(y)$$

2 Convergence in probability

$$\lim_{n\to\infty} \Pr(|y_n - y| > \epsilon) = 0$$

3 Convergence almost surely

$$\Pr(\lim_{n\to\infty}|y_n-y|=0)=1$$

Convergence of Random variables

- Convergence of random variables is a delicate subject
- Important to get a deeper understanding
- Not get intimidated while reading the literature; remember the definitions and different modes of convergence
- See, e.g., Grimmet and Stirzaker, Ch. 7

Law of Large Numbers

$$X_1, \ldots, X_n, \ldots$$
 are i.i.d.

• Weak Law: $\langle X_i \rangle = \mu$

$$\frac{X_1+\cdots+X_n}{n} o \mu$$
 in probability

Strong Law: $\langle X_i \rangle = \mu$ and X_i with finite variance

$$\frac{X_1+\cdots+X_n}{n} o \mu$$
 a.s.

Central Limit Theorem

 X_i are i.i.d. with mean μ and variance σ^2

$$\bar{X}_n = \frac{X_1 + \cdots + X_n}{n}$$

$$\frac{\sqrt{n}(\bar{X}_n - \mu)}{\sigma} \to \mathcal{N}(0, 1)$$

We have approximately

$$\bar{X}_n \sim \mathcal{N}(\mu, \sigma^2/n)$$

- The famous letters between Pascal and Fermat (start of probability) mention a request for help from a French nobleman and gambler, Chevalier de Méré.
- Méré bets for:
 - in four rolls of a die, at least one six would turn up
- Later he bets for:
 - in 24 rolls of two dice, a pair of sixes would turn up. but he was not happy with the latter schema

Setup a computer simulation for a single die

```
K = 4; % Number of dice throws
N = 1000; % Number of games
for trial=1:10,
    D = ceil(rand(N,K)*6);
    disp(sum(sum(D==6, 2) > 0)/N)
end
```

Per game, Méré won

```
0.4950, 0.4950, 0.5090, 0.5210, 0.5460
0.5420, 0.5360, 0.5160, 0.5210, 0.5010
```

The analytical solution

$$Pr\{M\text{\'er\'e wins}\} = 1 - Pr\{M\text{\'er\'e loses}\}\$$

= $1 - (5/6)^4 = 0.5177$

Setup a computer simulation for a pair of dice

```
K = 24; % Number of dice throws
N = 1000; % Number of games
for trial=1:10,
    D = ceil(rand(N,K,2)*6);
    sum(sum(D(:,:,1)==6 & D(:,:,2)==6,2) > 0)/N
end
```

Per game, Mere wins

```
0.502, 0.486, 0.497, 0.533, 0.521
0.474, 0.451, 0.508, 0.470, 0.481 ...
```

 Accurate results by simulation require a large number of experiments

The analytical solution

$$Pr\{M\acute{e}r\acute{e}wins\} = 1 - (35/36)^{24} = 0.4914$$

Therefore, 24 times is not a good bet. But with 25 (Pascal)

$$Pr\{M\acute{e}r\acute{e}wins\} = 1 - (35/36)^{25} = 0.5055$$

- ▶ What is the distribution of the estimate for *N* games ?
- V_n the outcome that Méré wins the n'th game

$$V_n \sim \mathcal{BE}(V_n; p)$$

$$S_n = \frac{V_1 + \dots + V_n}{n}$$

▶ Evoke the law of large numbers $\langle V_n \rangle = p$

$$S_n \rightarrow p \qquad n \rightarrow \infty$$

Accuracy is given by the Central Limit Theorem

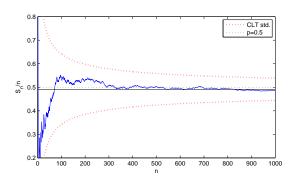
$$\langle V_n \rangle = p$$

$$Var\{V_n\} = p(1-p)$$

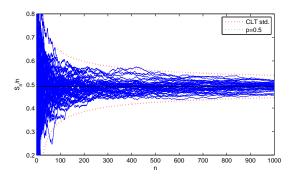
$$\sqrt{\frac{n}{p(1-p)}}(S_n - p) \rightarrow \mathcal{N}(0, 1)$$

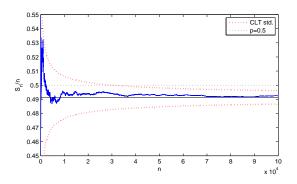
Approximately

$$S_n \sim \mathcal{N}(p, p(1-p)/n)$$



Chevalier de Méré (cont.)





▶ We need around 30000 games to say with about %99 confidence that the game with 24 throws is truly unfavorable.

Summary

- Law of large numbers: Consistency.
- CLT: Provides information about the rate of convergence
- ▶ If we can draw N independent and identically distributed samples from a distribution p(x), we can estimate expectations $E_p(\varphi(x))$ with an error $O(N^{1/2})$, **independent** of the dimensionality of x.