

Universitat de Lleida

Document downloaded from:

<http://hdl.handle.net/10459.1/65751>

The final publication is available at:

<https://doi.org/10.1016/j.scitotenv.2018.01.150>

Copyright

cc-by-nc-nd, (c) Elsevier, 2018

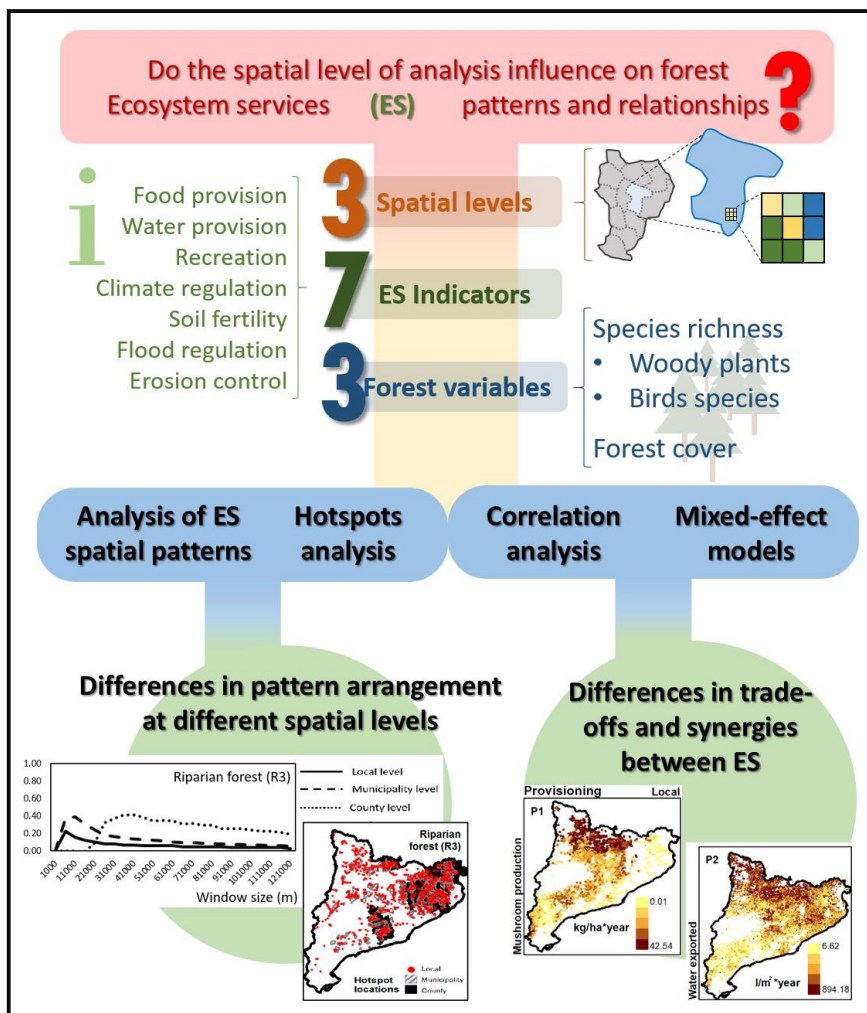


Està subjecte a una llicència de [Reconeixement-NoComercial-SenseObraDerivada 4.0 de Creative Commons](https://creativecommons.org/licenses/by-nc-nd/4.0/)

The spatial level of analysis affects the patterns of forest ecosystem services supply and their relationships

Highlights

- Scale is a relevant aspect in the analysis and of Ecosystem Services (ES)
- The effects of the spatial level of analysis on 7 ES indicators were assessed
- ES Indicators were estimated at local (1km²), municipality and county levels
- Averaging effects at higher spatial levels obscured local ES heterogeneity patterns
- Identification of hotspots and ES relationships depend on the level of analysis



14 **The spatial level of analysis affects the patterns of forest ecosystem services supply**
15 **and their relationships**

16

17 José V. Roces-Díaz^{1,2,*}, Jordi Vayreda¹, Mireia Banqué-Casanovas¹, Emilio Díaz-Varela³, Jose
18 A. Bonet^{4,5}, Lluís Brotons^{1,4,6}, Sergio de-Miguel⁵, Sergi Herrando^{1,7}, Jordi Martínez-Vilalta^{1,8}

19

20 1. CREAM, E08193 Bellaterra (Cerdanyola del Vallès), Catalonia, Spain

21 2. Agroscope, Reckenholzstr. 191, CH-8046 Zurich, Switzerland

22 3. ECOAGRASOC Research Group, Department of Agricultural and Forestry Engineering,
23 University of Santiago de Compostela, Escola Politécnica Superior, Campus Universitario, s/n.
24 27002 Lugo, Galicia, Spain

25 4. Forest Science Center of Catalonia (CTFC), 25280 Solsona, Catalonia, Spain

26 5. Department of Crop and Forest Sciences, Universitat de Lleida-Agrotecnio Center (UdL-
27 Agrotecnio), Av. Rovira Roure, 191, E-25198 Lleida, Catalonia, Spain

28 6. Consejo Superior de Investigaciones Científicas (CSIC), 08193 Cerdanyola del Valles,
29 Catalonia, Spain

30 7. Catalan Ornithological Institute, Natural History Museum of Barcelona, 08019 Barcelona,
31 Catalonia, Spain

32 8. Universitat Autònoma de Barcelona, E08193 Bellaterra (Cerdanyola del Vallès), Catalonia,
33 Spain

34

35 *: corresponding author José V. Roces-Díaz (jyroces@gmail.com) CREAM, +34935814850 Autonomous
36 University of Barcelona, E-08193 Bellaterra (Cerdanyola del Vallès), Catalonia, Spain

37 **Abstract**

38 The implementation of the Ecosystem Services (ES) framework (including supply and demand)
39 should be based on accurate spatial assessments to make it useful for land planning or
40 environmental management. Despite the inherent dependence of ES assessments on the spatial
41 resolution at which they are conducted, the studies analysing these effects on ES supply and
42 their relationships are still scarce. To study the influence of the spatial level of analysis on ES
43 patterns and on the relationships among different ES, we selected seven indicators representing
44 ES supply and three variables that describe forest cover and biodiversity for Catalonia (NE
45 Iberian Peninsula). These indicators were estimated at three different scales: local, municipality
46 and county. Our results showed differences in the ES patterns among the levels of analysis. The
47 higher levels (municipality/county) removed part of the local heterogeneity of the patterns
48 observed at the local scale, particularly for ES indicators characterized by a finely grained,
49 scattered distribution. The relationships between ES indicators were generally similar at the
50 three levels. However, some negative relationships (potential trade-offs) that were detected at
51 the local level changed to positive (and significant) relationships at municipality and county.
52 Spatial autocorrelation showed similarities between patterns at local and municipality levels, but
53 differences with county level. We conclude that the use of high-resolution spatial data is
54 preferable whenever available, in particular when identifying hotspots or trade-offs/synergies is
55 of primary interest. When the main objective is describing broad patterns of ES, intermediate
56 levels (e.g., municipality) are also adequate, as they conserve many of the properties of
57 assessments conducted at finer scales, allowing the integration of data sources and, usually,
58 being more directly relevant for policy-making. In conclusion, our results warn against the
59 uncritical use of coarse (aggregated) spatial ES data and indicators in strategies for land use
60 planning and forest conservation.

61 **Keywords**

62 Indicators; forest biodiversity; administrative boundaries; scale effects; trade-off and synergy;
63 upscaling

64 **1. Introduction**

65 Ecosystem services (ES) can be defined as those benefits provided directly and indirectly by the
66 ecological functioning of nature, and they are key for the wellbeing of human societies (MEA,
67 2005). This concept bridges science-based and societal considerations and has been growing in
68 relevance since the 1990s. Thus, different international initiatives appeared in the last 20 years
69 focused on their assessment (i.e. MEA, 2005; TEEB, 2010; IPBES, 2012), together with a
70 growing scientific interest (Seppelt et al., 2011; Boerema et al., 2016). Different authors have
71 highlighted the potential applications of the ES concept for sustainable land use planning (Daily
72 et al., 2009; Baró et al., 2016), natural resources management (Tallis and Polasky, 2009) or
73 biodiversity conservation (Chan et al., 2011). At the same time, there is a need to develop
74 integrative frameworks for ES assessment (Kremen, 2005; de Groot et al., 2010) including
75 biodiversity, bio-physical and social aspects of the environment, and also covering as much as
76 possible the different components of ES (cascade approach including supply, demand and flow)
77 (Potschin and Haines-Young, 2010; Yahdjian et al., 2015).

78
79 The implementation of environmental management based on ES needs to be based on spatial
80 approaches (Egoh et al., 2008; Andrew et al., 2015) that involve mapping and characterizing
81 both ES supply and demand (Burkhard et al., 2012). Consistent with this, most ES assessments
82 (and ES-based studies) performed in recent years have included a spatially explicit perspective
83 (Seppelt et al., 2011). However, different authors have pointed out the need to account for
84 spatial patterns in more rigorous ways (Boerema et al., 2016) and to reduce the uncertainty
85 associated with ES mapping methods (Hou et al., 2013). The effect of scale on ES distribution
86 patterns and their spatial relationships has been highlighted in different works (e.g., Martín-
87 López et al., 2009; Geijzendorffer et al., 2015). As ES are generated by different ecosystem
88 types and ecological processes with different spatial patterns, their supply may differ between
89 scales (Hein et al., 2006; Roces-Díaz et al., 2014). Although the analysis of spatial patterns at
90 landscape and regional scales is extensively developed through spatial statistics, landscape
91 metrics and spatially explicit models (e.g. Wagner and Fortin, 2005; Uuemaa et al., 2009; Fortin
92 et al., 2012; Uuemaa et al., 2013), there is a limited knowledge on what are the most appropriate
93 scales of analysis to assess ES and their spatial relationships for different applications in the
94 context of land management, policy and decision making (Andrew et al., 2015; Schröter et al.,
95 2015).

96
97 Importantly, scale effects cannot only affect the absolute values of ES indicators but also the
98 relationships among them (Xu et al., 2017). When the provision of a given ES is increased at the
99 expense of another ES a trade-off occurs, while a mutual positive relationship, in which both ES
100 increase at the same time, can be defined as a synergy (Rodríguez et al., 2006; Bennett et al.,
101 2009). Previous work did not find large differences on the relationships between ES patterns
102 and biodiversity comparing different pixel sizes (Anderson et al., 2009), and similar ES patterns
103 across different administrative levels and spatial scales has been reported (Raudssep-Hearne and
104 Peterson, 2016). It is unknown, however, whether these results can be generalized.

105
106 For ES assessments to be useful for planning and management objectives they need to be
107 conducted at relevant spatial scales, which frequently correspond to administrative levels, as
108 those facilitate policy implementation (Tolvanen et al., 2014). The UN Strategic Plan for

109 Biodiversity 2011-2020 urges subnational administrations to consider the development of
110 biodiversity strategies to achieve the targets on biodiversity conservation, including the
111 provision of ES (Aichi goal D, CBD, 2011-2020). In this way the role of regional (Schulp et al.,
112 2014), county (Chen et al., 2009) and municipality (Rodriguez-Loinaz et al., 2015; Renard et
113 al., 2015) administrations is becoming more relevant to assess ES and the corresponding policy-
114 making at these subnational levels. At the same time, increasing the spatial level of analysis is at
115 the cost of homogenization of landscape patterns and loss of local information (Díaz-Varela et
116 al., 2009; 2016).

117

118 In this work we explored the effects of using different spatial levels of analysis on ES patterns
119 and their spatial relationships, in order to improve the integration of the ES framework on
120 national and sub-national strategies for planning and conservation of natural resources. The
121 specific objectives of this work are to: i) analyse the effects of spatial resolution on the spatial
122 patterns of forest ES-, including the location of the areas of highest supply (hotspots); and ii)
123 assess the impact of the level of analysis on the relationships (potential trade-offs and synergies)
124 among different ES, and between ES and forest biodiversity. We compare three levels of spatial
125 resolution: local (~1 km²), municipality and county, using 10 indicators, including seven ES
126 (food and water provision, climate regulation, soil fertility, flood regulation, erosion control and
127 recreation), forest cover and two biodiversity measures (woody plants and bird species
128 richness). Our study area is a highly populated and environmentally diverse Mediterranean
129 region in Catalonia (NE of Iberian Peninsula). In comparison with other regions in the
130 Mediterranean context, the study area shows high forest cover and population density, and a
131 wide variety of forest types due to the marked altitudinal and climatic gradient in this region.

132

133

134 **2. Material and methods**

135 **2.1. Study area**

136 Our study area is Catalonia (NE of Spain; Figure 1), an administrative region covering 32,114
137 km² and mainly located in the Mediterranean biogeographic region. Catalonia and its
138 subregional administrations have shared political responsibilities in planning and managing
139 biodiversity and ES. Catalonia has a population of 7.5 million people, most of them living in or
140 around the capital city (Barcelona). It is a mountainous area with an altitudinal range from the
141 sea level to more than 3,000 meters. It is a highly forested region (43% of its area is covered by
142 forests), with the main tree species belonging to the genera *Pinus* and *Quercus*. The forests from
143 coastal and low altitude areas are dominated by *Pinus halepensis* and *Quercus ilex*. At mid-
144 altitudinal ranges –from 800 to 1,500 m- the main species are *P. sylvestris*, *P. nigra*, *Q. humilis*
145 and *Q. faginea* and also *Fagus sylvatica* in the wettest zones. Finally, at altitudes higher than
146 1,500 m the main species are *P. uncinata* and *Abies alba*. The study area is divided in 41
147 counties (average extension = 783.1 km², range = 114.7 - 1784.1 km²) and 947 municipalities
148 (average extension = 33.9 km², range = 0.6 - 302.8 km²).

149

150 **2.2. Data sources**

151 In this work, we analysed the spatial patterns of a series of seven ES indicators (food and water
152 provision, climate regulation, soil fertility, flood regulation, erosion control and recreation) and
153 additional descriptors of forest cover and biodiversity (woody and bird species richness) at three

154 different spatial scales: local (1-km² cells or forest inventory plots), municipality and county.
155 Two sources of information were particularly important for estimating these indicators. On the
156 one hand, the Third Spanish National Forest Inventory (SNFI; MAGRAMA 1997-2007), which
157 provides detailed, plot-level information of forest characteristics, with a sampling density of one
158 plot every ~1 km² of forest area. The SNFI records species identity, height and diameter at
159 breast height (DBH) of all living and standing dead trees on circular plots of variable radius
160 depending on tree size (5 m radius for trees with DBH \geq 7.5 cm, 10 m radius for trees with DBH
161 \geq 12.5 cm, 15 m radius for trees with DBH \geq 22.5 cm and 25 m radius for trees with DBH \geq
162 42.5). On the other hand, the Land Cover Map of Catalonia (LCMC, 2009) was used to obtain
163 high resolution thematic cartography of the land cover of Catalonia. It is a vector map generated
164 by photo-interpreting on 1:5000 colour ortho-photo images, with a minimum patch area of
165 500 m², a working spatial scale of 1:1000, a pixel resolution of 0.25 meters and 241 different
166 legend categories.

167

168 **2.3. Description and calculation of forest ES indicators**

169 We worked with seven indicators of three main groups of ES (i.e., provisioning (P), cultural (C)
170 and regulating (R)), following widely used ES classifications (e.g. CICES (Haines-Young and
171 Potschin, 2012)). We also assessed two biodiversity indicators and a measure of forest cover. All
172 these indicators are described briefly in the following paragraphs; more detailed information can
173 be found elsewhere (Roces-Díaz et al., 2017b).

174

175 **Mushroom production** (P1, food provision, Provisioning ES) was estimated for each SNFI
176 plot using the models developed by de-Miguel et al. (2014) for the study area. These models
177 were developed based on the monitoring, from 1995 to 2012, of weekly mushroom production
178 from permanent sample plots representing most pine forest ecosystems throughout the study
179 region, i.e., pure and mixed stands of *P. sylvestris*, *P. nigra*, *P. halepensis* and *P. pinaster*.
180 Mixed-effects models based on site and stand characteristics were used to estimate the potential
181 production of edible mushrooms of commercial interest for a typical year, in kg·ha⁻¹·year⁻¹. In
182 the study area edible mushrooms occur primarily in pine forests.

183

184 **Exported water** (P2, water provision, Provisioning ES) was calculated also for each SNFI plot
185 using the water balance model developed by de Cáceres et al. (2015) for the study area. This
186 model provides information on the amount of water exported (blue water, in L·m⁻²·year⁻¹) from
187 each forest plot based on the physical properties of soil, climate and forest composition and
188 structure.

189

190 **Number of animal observations** (C1, recreation, Cultural ES) per 1 km² forest cell was
191 calculated using the data from the web portal of the Catalan Ornithological Institute (Sardà-
192 Palomera et al., 2012; Herrando-Moraira et al., 2016), www.ornitho.cat storing more than
193 3,000,000 observations. Bird watching is an important recreational activity in the study area.
194 Data represents observations of animal species (including birds, mammals, reptiles, amphibians
195 and some groups of invertebrates), and consists of observations uploaded by users of the
196 mentioned webpage all around Catalonia between 2010 and 2015. Only those cells where the
197 forest cover was dominant (>50% of the total area) were included in further analyses.

198

199 **Carbon sequestration** (R1, climate regulation, Regulating ES) was calculated following
200 Vayreda et al. (2012) as the carbon stock change in each SNFI plot (in t·ha⁻¹·year⁻¹) between
201 two consecutive surveys (2nd and 3rd SNFI survey)). In Catalonia, the 2nd SNFI was conducted in
202 1990 and the 3rd SNFI in 2000, so the period between surveys of a given plot was ~10 years.

203

204 **Soil organic Carbon** (SOC; R2, soil fertility, Regulating ES) was estimated as the average
205 value of Carbon stored in soil (t·ha⁻¹) obtained from the map of soil organic carbon of Spain
206 (Doblas-Miranda et al., 2013). This map provides the SOC content on a grid with a cell size 200
207 x 200 m. It is based on more than 900 field samples that were used to build a multiple
208 regression model with environmental data (climate, vegetation cover) as independent variables.
209 Only those cells where the forest cover was dominant (>50% of the total area) were included in
210 further analyses.

211

212 **Riparian forest cover** (R3, flood regulation, Regulating ES) was calculated as the percentage
213 of area along rivers (50 m width buffer) that is covered by forest in each 1x1 km cell (according
214 to the LCMC). Only those cells where river occurrence was relevant (the buffers around rivers
215 covered >5% of the pixel area) were used in further analyses.

216

217 **Forest cover in slopes** (R4, erosion control, Regulating ES) shows the percentage of landscape
218 with a slope $\geq 30\%$ that is covered by forest, in each 1x1 km cell, based on the original 0.25 m
219 cells of the LCMC. Only those cells where the forest cover was dominant (>50% of the total
220 area) were included in further analyses.

221

222 *2.4. Calculation of additional forest descriptors*

223 **Woody plant richness** (B1, Biodiversity) represents the number of different woody species
224 detected in each SNFI plot, municipality or county, based on the information collected in the
225 third SNFI survey.

226

227 **Forest bird richness** (B2, Biodiversity) correspond the number of forest bird species estimated
228 on each 1 x 1 km cell, municipality or county. The data was based on the second Catalan
229 Breeding Bird Atlas (Estrada et al. 2004). Fieldwork was conducted between 1999 and 2002
230 and represents a total of 3,077 cells of 1 km² resolution. Presence/absence data for breeding bird
231 species and environmental predictors were used to develop species distribution models and a
232 cross-validation procedure was applied for the assessment of model performance.

233

234 **Forest cover** (F1) represents the percentage of area of each level (1x1 km cell, municipality or
235 county) covered by forest ecosystems.

236

237 *2.5. Data analyses*

238 Most indicators (mushrooms production, exported water, number of animal observations,
239 carbon sequestration and soil organic carbon) were scaled up to the municipality and county
240 levels by averaging the point or grid data values at 1 km² resolution. To reduce methodological
241 artefacts, the variables that were computed as a percentage of forest cover (riparian forest cover,
242 forest cover in slopes and forest cover) were calculated directly from raw data at the
243 municipality and county levels. Finally, biodiversity variables always corresponded to the total

244 number of different species (woody species for B1 or forest birds for B2) detected at each
 245 spatial scale (plot/grid cell, municipality or county).

246

247 Some original variables were transformed by using square root or logarithmic functions to
 248 normalize their distribution prior to analysis. However, some variables could not be normalized
 249 (see table 1 for specific information on the transformation used for each variable).

250

251 As our main objectives was to assess the spatial pattern of ES, hotspots and spatial
 252 autocorrelation analyses were performed only for the seven ES indicators. The analyses were
 253 performed using ArcGIS 10.2 (ESRI, 2011). The hotspots analysis was conducted separately at
 254 the three spatial levels of analysis. The hotspot areas of each indicator were calculated using two
 255 different methods depending on the data type (point data from SNFI plots or grid data). For the
 256 point data, hotspots were defined following a clustering method based on the Getis-Ord G_i^*
 257 statistic (Getis and Ord, 1992; Ord and Getis, 1995), which is appropriate to feature data types
 258 (Schröter and Remme, 2015). This statistic is calculated as (eq.1):

$$259 \quad G_i^* = \frac{\sum_{j=1}^n w_{i,j} x_j - \left[\frac{\sum_{j=1}^n x_j}{n} \right] \sum_{j=1}^n w_{i,j}}{S \sqrt{\frac{n \sum_{j=1}^n w_{i,j}^2 - (\sum_{j=1}^n w_{i,j})^2}{n-1}}} \quad \text{Eq.1}$$

260 Where n is the number of spatial features; $w_{i,j}$ is the distance between features i and j ; x_j is the
 261 value of each ES indicator, and S is calculated as follows (eq.2):

$$262 \quad S = \sqrt{\frac{\sum_{j=1}^n x_j^2}{n} - \left[\frac{\sum_{j=1}^n x_j}{n} \right]^2} \quad \text{Eq.2}$$

263 For a given dataset, the Getis-Ord G_i^* statistic identifies those clusters of spatial features with
 264 values of ES supply higher than those expected by random chance we represented those features
 265 identified as hotspots with 95% of probability. For the grid data, the calculation of the Getis-Ord
 266 G_i^* statistic is not possible. Therefore, we defined hotspots as those areas that represented the
 267 highest values of supply (the cells where the supply was >80% percentile; Schröter and Remme,
 268 2015). In order to quantify the differences produced by the delimitation of hotspot areas at the
 269 three spatial levels, we compared the percentage overlap between these areas at the different
 270 spatial scales.

271

272 Differences in spatial patterns of ES indicators among the three levels of analysis were further
 273 inspected using Spatial Autocorrelation. Firstly, Moran's I coefficient (Moran, 1948) (eq.3) was
 274 calculated for the patterns of each ES indicator at the three different levels:

$$275 \quad I = \frac{n}{\sum_{i=1}^n \sum_{j=1}^n w_{i,j}} \frac{\sum_{i=1}^n \sum_{j=1}^n w_{i,j} z_i z_j}{\sum_{i=1}^n z_i^2} \quad \text{Eq.3}$$

276 Where n is the number of spatial features; $w_{i,j}$ is the distance between features i and j ; z_i and z_j
 277 are the deviation of the attribute (here the value of each indicator) for feature i and j ,
 278 respectively, from the mean value. This index allowed determining the spatial clustering of the
 279 ES indicators at the three levels of analysis. Secondly, Incremental Spatial Autocorrelation
 280 (ISA) was used for estimating spatial autocorrelation at increasing distances. This method is
 281 also based on the calculation of Moran's I coefficient using a moving window size. For our
 282 analyses window size ranged from 1,000 (based on minimum distance among plots) to 126,000
 283 meters, with an increment of 5,000 m. This range of values used was similar to a previous

284 application of this methodology on the analysis of ES patterns (Roces-Díaz et al., 2017a).

285

286 To analyse possible trade-offs and synergies between ES and forest variables (table 1) we
287 calculated Pearson and Spearman correlations between pairs of variables at the three different
288 levels of analysis (Mouchet et al., 2014). We calculated these correlations (Pearson for those
289 involving variables with normal distribution and Spearman for those involving at least one
290 variable we could not normalize; cf. Table 1) by using the R statistical software (v.3.2.5; R
291 Development Core Team, 2016).

292

293 Finally, to explore in more detail the relationships between the ES indicators and biodiversity
294 and forest cover we fitted linear models at the different levels: local, municipality and county.
295 We used woody species richness (B1), bird richness (B2) and forest cover (F1) as independent
296 variables and the ES indicators as dependent variables. We assessed multi-collinearity among
297 explanatory variables by calculating Variance Inflation Factors (VIF), which were always < 1.5 ,
298 confirming that multi-collinearity was not an issue in our models. For local and municipal level
299 we fitted linear mixed-effects models with municipality nested in county (local level) or county
300 (municipality level) as random factors. For the county level we fitted standard linear models.
301 The residuals of these models did not show any obvious pattern and their distribution was
302 approximately normal (Supplementary material Appendix A). In all statistical analyses
303 significance was accepted whenever the p-value < 0.05 .

304

305

306 **3. Results**

307 ***3.1. Spatial distribution of biodiversity variables, forest cover and ES indicators***

308 Forest cover and forest biodiversity variables showed differences in their spatial patterns at the
309 three levels analysed (Figure 2). In all three cases, relatively high values of forest cover (F1),
310 woody species richness (B1) and bird richness (B2) found at the local level in southern areas of
311 the study region progressively disappeared at coarser spatial scales (municipality and county),
312 for which highest values clustered more and more towards the northern part of the study area.

313

314 The distribution of ES indicators was more conservative and showed broadly similar patterns at
315 the three levels of analysis (Figures 3 and 4). Thus, those areas that showed higher supply
316 values at local levels also tended to present high supply at municipality and county levels.
317 However, some differences among these levels can also be highlighted. For example, exported
318 water (P2) and erosion control (R4) showed some of the highest values at local level in the
319 north-east of the study area, but this pattern disappeared at coarser scales (Figures 3 and 4).
320 Overall, most ES indicators showed highest values towards the north of the study area at all
321 levels, with the exception of animal observations (C1), which were clustered along the vertical
322 corridor linking Barcelona to the French border. Some indicators (e.g., carbon sequestration, R1
323 and riparian forest, R3) showed a scattered, fine-grained pattern with high supply zones across
324 all the study area, which was highly eroded at the county level.

325

326 ***3.2. Characterization of ES spatial patterns***

327 Spatial patterns of hotspots areas showed marked differences among the different ES indicators
328 (Figure 5). Overall, most ES indicators hotspots were located on the northern half of Catalonia.

329 However, two indicators (animal observations and riparian forest) presented their hotspots more
330 uniformly distributed across all the forests in the study area. Some remarkable differences were
331 detected among levels of analysis. While some indicators showed highly overlapping hotspots at
332 the three levels (e.g., mushroom production, exported water or soil organic carbon), the hotspots
333 of other indicators were clearly disjoint across scales. This was particularly the case for ES
334 indicators showing scattered hotspot patterns distributed over most of the study area at local
335 level (animal observations and riparian forest). In those cases, the fine grained spatial
336 distribution of hotspots was generalized at coarser scales and hotspots tended to concentrate
337 towards the north of the study area. The percentage of spatial agreement of hotspots at the three
338 levels is shown in Table 2. Mushroom production, exported water, soil carbon and erosion
339 control presented a high level of correspondence across scales (~90% agreement or higher
340 between local and municipality levels, >70% between local and county levels). At the other
341 extreme, animal observations and riparian forest presented much lower overlaps, in the order of
342 50% or lower, across spatial scales. Overall, the level of agreement was higher between local
343 and municipality levels than between local and county levels, with the only exception of
344 wildlife observations.

345
346 All ES indicators showed substantial spatial autocorrelation, characteristic of clustered patterns
347 (Supplementary material Appendix B). Moran's I coefficient showed higher values for most of
348 the ES indicators when the level of analysis increased (0.07-0.61 for the local level; 0.26-0.63
349 for municipality; and 0.17-0.69 for county). In a similar way, incremental spatial autocorrelation
350 analysis showed differences in the spatial aggregation patterns among levels of analysis for the
351 seven ES indicators (Figure 6). Local and municipality levels showed similar aggregation
352 patterns, with maximum spatial autocorrelation at distances between 1,000 and 11,000 meters.
353 However, the situation was very different at the county level, which showed maximum spatial
354 autocorrelation at distances >31,000 m, reflecting the larger characteristic size of these spatial
355 units.

356
357 The pair-wise relationships between ES indicators, and among ES indicators and forest cover
358 and biodiversity variables, were generally positive and significant, indicating a preponderance
359 of synergies over trade-offs (Table 3). The strongest and most consistent relationships across
360 scales were those between mushroom production, exported water, soil carbon and erosion
361 control. Interestingly, the value of most correlation coefficients increased from local to county
362 levels. In particular, negative (and significant) relationships were only observed at the local
363 level, most of them between woody species richness and some ES indicators, such as mushroom
364 production, exported water, soil carbon and erosion control. These negative relationships
365 disappeared at municipality and county levels, where they shifted to positive (but not always
366 significant) correlations (Table 3). Thus, some potential trade-offs detected at the finest spatial
367 resolution were not detected, or even became positive, potentially synergistic effects, at coarser
368 spatial scales.

369
370 Linear models of ES indicators as a function of forest cover and biodiversity variables also
371 showed substantial differences among levels of analysis (Table 4), which were generally
372 consistent with the results obtained from the correlation analysis. Thus, some explanatory
373 variables showed significant negative effects at local level but not at municipality or county

374 levels (e.g., woody species richness on mushroom production and exported water).
375 Interestingly, the negative relationship of woody species richness with soil carbon remained
376 significant at the three spatial scales according to the corresponding linear models. Bird richness
377 showed generally positive (and significant) relationships with most of the indicators. The local
378 effect of forest cover could be either positive (mushroom production, soil organic carbon,
379 erosion control) or negative (animal observations, riparian forest), and tended to decline at
380 larger spatial scales. Finally, the explained variance (model R^2) varied among ES indicators,
381 with R^2 (standard or conditional depending on model type) being ≥ 0.32 in all cases except for
382 animal observations and carbon sequestration, for which the explained variance was much lower
383 (Table 4). Scatter-plots among forest descriptors and ES indicators at three levels are showed in
384 supplementary material.

385

386 **4. Discussion**

387 ***4.1. Influence of the spatial level of analysis on ES patterns and hotspots areas***

388 At the international level, different policies are integrating ES approaches in biodiversity
389 conservation and environmental management strategies (e.g., EU Biodiversity Strategy to
390 2020). Although this integration requires the use of specific methodologies for the accurate
391 mapping and characterization of ES (Maes et al., 2012), spatial ES assessments frequently
392 show inconsistencies and mismatches, which often result in relatively high uncertainty
393 (Crossman et al., 2013; Hou et al., 2013; Geijzendorffer et al., 2015). Among these sources of
394 uncertainty, those derived from scale effects are particularly important because they underlie the
395 identification of ES spatial patterns, which is critical when ES assessments are translated into
396 land-use or management decisions (Xu et al., 2017).

397

398 Our results confirmed the influence of scale of analysis on the distribution patterns of different
399 ES indicators. These results appear logical considering the scale-dependency of the ecological
400 processes underlying ES provision (Hein et al., 2006) and are consistent with previous findings
401 showing the influence of spatial data characteristics (e.g., spatial, temporal or thematic
402 resolution) on ES patterns (Konarska et al., 2002; Kandziora et al., 2013; Gret-Regamey et al.,
403 2014), including indicators of the supply, demand and flow of ES (e.g., Bagstad et al., 2014;
404 Wolff et al., 2015). Although it is known that the spatial level of analysis may influence ES
405 assessments (Geijzendorffer et al., 2015), the number of works assessing the importance of this
406 effect is still limited (e.g., Raudssep-Hearne and Peterson, 2016) and most ES assessments are
407 based on the use of administrative (or similar) boundaries including municipalities, counties or
408 larger levels (referred generically as NUTS (Nomenclature of territorial units for statistics)
409 Rodriguez-Loinaz et al., 2015; Roces-Díaz et al., 2017b; Schulp et al., 2014).

410

411 In our study, the level with finest resolution (local) was based on 1-km² cells, and provides a
412 similar level of detail to previous works that assessed similar sets of ES for comparable study
413 areas (e.g., Anderson et al., 2009; Eigenbrod et al., 2010; Locatelli et al., 2013). In general, ES
414 indicators showed relatively heterogeneous spatial patterns, which turned into more
415 homogeneous patterns at coarser spatial scales. This is a recognized effect of spatial aggregation
416 (Levin, 1992; Constanza and Maxwell, 1994), with consequences for data interpretation and the
417 corresponding management decisions: while in some instances generalization may be a
418 necessary (and desirable) means of dealing with high spatial variability, the resultant averaging

419 effect can dismiss important fine-grained information.

420

421 The distribution of ES indicators and the corresponding hotspots showed higher agreement
422 between the local and municipality levels of analysis than between the local and the county
423 levels. Although this agreement was less marked for those ES with scattered spatial patterns
424 (i.e., animal observations and riparian forest), it shows that the analysis at the municipality level
425 reflected better the variability provided by fine-grain information and is more accurate than the
426 analysis at coarser administrative levels. The incremental spatial autocorrelation (ISA) analysis,
427 which showed similar spatial patterns at the local and municipality levels, but highly distinct
428 patterns at the county level, also supported this notion. Our results are generally consistent with
429 previous work exploring how spatial patterns of ES hotspots are affected by spatial resolution
430 (Eigenbrod et al., 2010; Homolova et al., 2014) and how spatial autocorrelation of ES patterns
431 depends on spatial resolution (Gret-Regamey et al., 2014). The municipality level can be
432 highlighted as a convenient scale for ES analysis that allows to integrate indicators from a
433 multiplicity of data sources (e.g., Raudsepp-Hearne et al., 2010; Rodriguez-Loinaz et al., 2015;
434 Rocés-Díaz et al., 2017b) and, at the same time, provides relatively accurate spatial patterns.

435

436 ***4.2. Spatial relationships and influence of analysis level***

437 The spatial level of analysis also influenced the relationships between pairs of ES indicators,
438 and between them and forest cover and biodiversity variables. Our results confirm how these
439 processes, including trade-offs and synergies between ES, are dependent on spatial scales
440 (Rodriguez et al., 2006). It is well recognized that the implementation of the ES approach on
441 land planning and natural resources management needs an accurate assessment of these types of
442 effects at different scales, as previous research has shown that administrative boundaries can
443 affect their identification (Deng et al., 2016). In this regard, we found that an increase in the
444 spatial level of analysis can mask potential trade-offs among ES, particularly when local
445 data are compared with aggregated indicators at broader scales (i.e. municipality or county).
446 These findings are in agreement with previous works (e.g., Yang et al., 2015; Liu et al., 2017)
447 where fewer trade-offs (and more synergies) among ES were detected at larger compared to
448 finer scales.

449

450 In general, our results showed stronger correlations at lower spatial resolution
451 (municipality/county levels), in agreement with other works where the influence of spatial
452 levels of analysis on trade-offs/synergies was explored (Anderson et al., 2009; Xu et al., 2017).
453 Although most indicators/variables showed consistent relationships across scales (cf. Raudsepp-
454 Hearne and Peterson, 2016), some of them presented contrasted relationships at local vs. coarser
455 scales. This was particularly the case of woody species richness, which showed negative and
456 significant correlations (i.e. potential trade-offs) with several ES indicators, including regulating
457 and provision ES, at the local level that turned into positive correlations (i.e. potential synergies)
458 at coarser spatial scales. This mismatch indicates the importance of landscape heterogeneity
459 (gamma vs. local, alpha diversity) in supporting high levels of ES supply, and it is consistent
460 with the notion that coarser scales may describe a spatial mosaic arrangement that allows
461 several ES to concur at the landscape level synergically (e.g. in a multifunctional rural
462 landscape).

463

464 The negative relationships obtained between woody species richness and ES contrast with many
465 studies showing consistently positive relationships between biodiversity and ES (Egoh et al.,
466 2009; Harrison et al., 2014; Strassburg et al., 2010), but agree with other works focused on
467 forest ES (i.e., Locatelli et al., 2013; Lautenbach et al., 2017). Most of the ES indicators for
468 which negative relationships with biodiversity were found are highly dependent on climatic
469 productivity (mushroom production in de-Miguel et al., 2014; exported water in de Caceres et
470 al., 2015; soil organic carbon in Doblas-Miranda et al., 2013). In the study area, these areas with
471 high productivity often involve historically managed forests that are characterized by low tree
472 richness (frequently monospecific stands focused largely on timber production-(Onaindia et al.,
473 2013; Rodriguez-Loinaz et al., 2013).

474

475 Although linear models generally confirmed the presence of some negative relationships
476 between ES indicators, particularly at the local level, they were generally more consistent across
477 spatial scales than simpler correlation analyses. This result suggests that some of the
478 inconsistencies across scales might be explained by covariance with third variables. Overall, our
479 results highlight the importance of assessing relationships among ES at different spatial and
480 temporal scales (Tomscha and Gergel, 2016) to obtain a robust (and interpretable)
481 characterization of potential trade-offs and synergies between them.

482

483 ***4.3. Main strengths and limitations***

484 The two coarser spatial scales that we used correspond to administrative boundaries often used
485 for land use planning and management, and are thus directly relevant from an applied
486 perspective. In addition, the set of ES indicators (and biodiversity variables) used in this work is
487 based on the combination of field and environmental data (as recommended by Martinez-Harms
488 et al., 2016) that account, whenever possible, for the underlying ecological processes. Finally,
489 the indicators we have chosen provide information on different types of forest ES and include
490 those believed to be more relevant in the study area.

491

492 On the other hand, this study has a series of potential limitations that should be highlighted.
493 Firstly, the local analysis level is based (for some indicators such as for example soil organic
494 carbon or erosion control) on regular grids, while municipality and county levels are derived
495 from administrative boundaries that involved a wide range of sizes and shapes. Thus,
496 differences on spatial patterns among these levels could be influenced by these inherent
497 differences in shape and distribution. In addition, the combination of several data sources
498 allowed analysing a wide range of ES. However, some of these indicators derived from primary
499 data, while others were based on ecological deterministic models or land use maps, and
500 differences in data sources and estimation approaches may affect spatial patterns (Eigenbrod et
501 al., 2010; Martínez-Harms et al., 2016). Finally, to provide a consistent set of ES indicators at
502 the three spatial levels of analysis, some of them had to be obtained using relatively simplified
503 approaches (compare for example with the indicators developed in Guerra et al. (2016) for
504 erosion control).

505

506

507 **5. Conclusions**

508 We explored the effect of using different spatial scales and administrative boundaries on ES

509 assessment and mapping. We report substantial information loss when coarser spatial scales
510 (county level) were used, whereas spatial patterns at the local and municipality level remained
511 similar. Some trade-offs among ES and between ES and biodiversity were only detected at the
512 local scale, implying that caution is needed when interpreting relationships between ES at
513 relatively coarse spatial scales. Following this, the use of high-resolution-data (when available)
514 is recommended, in particular when identifying hotspots areas or trade-offs/synergies are of
515 primary interest. In more descriptive assessments in which the main objective is describing
516 broad spatial patterns of ES distribution, intermediate levels (municipality) are also adequate, as
517 they conserve many of the spatial properties of assessments conducted at finer spatial scales and
518 have the advantage of being more directly relevant for policy-making.

519

520 **Supplementary material**

521 Appendix A shows the residuals of the linear models fitted at different spatial scales.

522 Appendix B shows the results of Spatial Autocorrelation tests.

523 Appendix C shows scatter-plots among forest descriptors (biodiversity and forest cover
524 variables) and ES indicators at the three spatial levels of analysis.

525

526 **Acknowledgments**

527 We thank to the volunteers from the Catalan Ornithological Institute (ICO) and Dr. Miquel de Cáceres
528 Ainsa for providing data for the analyses presented in this study. Funding was obtained from the Catalan
529 Office for Climate Change (OCCC) through project ForESMap, from EU FORESTERRA program
530 (INFORMED project) and from the Spanish government (CGL2013-46808-R and AGL2015-66001-C3-
531 1-R). JVRD was supported by the Government of Asturias and the FP7-Marie Curie-COFUND program
532 of the European Commission (Grant 'Clarín' ACA17-02). We also thank the ECOMETAS (CGL2014-
533 53840-REDT) network for support. This study also received funding from the European Union's Horizon
534 2020 research and innovation programme within the framework of the MultiFUNGtionality Marie
535 Skłodowska-Curie Individual Fellowship (IF-EF) under grant agreement No655815 and from the
536 Generalitat de Catalunya (Serra-Hunter Fellow). We thank Gabriel Borrás and Gemma Cantos (OCCC)
537 for useful discussion during the elaboration of this work. We are very grateful to all persons who made
538 the two Spanish Forest Inventories possible and, especially, to their main coordinators, Ramon
539 Villaescusa (IFN2) and Jose Antonio Villanueva (IFN3). We also thank two anonymous reviewers who
540 helped us improve the quality of the manuscript.

541

542

543

543 **6. References**

- 544 Anderson, B.J., Armsworth, P.R., Eigenbrod, F., Thomas, C.D., Gillings, S., Heinemeyer, A., Roy, D.B.,
545 Gaston, K.J., 2009. Spatial covariance between biodiversity and other ecosystem service priorities. *J.*
546 *Appl. Ecol.* 46, 888–896. doi:10.1111/j.1365-2664.2009.01666.x
- 547 Andrew, M.E., Wulder, M.A., Nelson, T.A., Coops, N.C., 2015. Spatial data, analysis approaches, and
548 information needs for spatial ecosystem service assessments: a review. *GIScience Remote Sens.* 52,
549 344–373. doi:10.1080/15481603.2015.1033809
- 550 Bagstad, K.J., Villa, F., Batker, D., Harrison-Cox, J., Voigt, B., Johnson, G.W., 2014. From theoretical to
551 actual ecosystem services: mapping beneficiaries and spatial flows in ecosystem service assessments.
552 *Ecology and Society* 19: 64. doi.org/10.5751/ES-06523-19026
- 553 Baró, F., Palomo, I., Zulian, G., Vizcaino, P., Haase, D., Gómez-Baggethun, E., 2016. Mapping
554 ecosystem service capacity, flow and demand for landscape and urban planning: A case study in the
555 Barcelona metropolitan region. *Land use policy* 57, 405–417. doi:10.1016/j.landusepol.2016.06.006
- 556 Bennett, E.M., Peterson, G.D., Gordon, L.J., 2009. Understanding relationships among multiple
557 ecosystem services. *Ecology Letters*, 12, 1394–1404. doi:10.1111/j.1461-0248.2009.01387.x
- 558 Boerema, A., Rebelo, A.J., Bodi, M.B., Esler, K.J., Meire, P., 2016. Are ecosystem services adequately
559 quantified? *J. Appl. Ecol.* doi:10.1111/1365-2664.12696
- 560 Burkhard, B., Kroll, F., Nedkov, S., Müller, F., 2012. Mapping ecosystem service supply, demand and
561 budgets. *Ecol. Indic.* 21, 12–29. doi:10.1016/j.ecolind.2011.06.019

562 CBD, Convention on Biological Biodiversity, 2011-2020. Aichi Biodiversity Targets.
563 <https://www.cbd.int/sp/targets/default.shtml>

564 Chan, K.M. a, Hoshizaki, L., Klinkenberg, B., 2011. Ecosystem services in conservation planning:
565 Targeted benefits vs. co-benefits or costs? *PLoS One* 6. doi:10.1371/journal.pone.0024378

566 Chen, N., Li, H., Wang, L., 2009. A GIS-based approach for mapping direct use value of ecosystem
567 services at a county scale: Management implications. *Ecol. Econ.* 68, 2768–2776.
568 doi:10.1016/j.ecolecon.2008.12.001

569 Costanza, R., Maxwell, T., 1994. Resolution and predictability: an approach to the scaling problem.
570 *Landsc. Ecol.* 9, 47–57.

571 Crossman, N.D., Burkhard, B., Nedkov, S., Willemen, L., Petz, K., Palomo, I., Drakou, E.G., Martín-
572 Lopez, B., McPhearson, T., Boyanova, K., Alkemade, R., Egoh, B., Dunbar, M.B., Maes, J., 2013. A
573 blueprint for mapping and modelling ecosystem services. *Ecosyst. Serv.* 4, 4–14.
574 doi:10.1016/j.ecoser.2013.02.001

575 Daily, G.C., Polasky, S., Goldstein, J., Kareiva, P.M., Mooney, H. a, Pejchar, L., Ricketts, T.H., Salzman,
576 J., Shallenberger, R., 2009. Ecosystem services in decision making: time to deliver. *Front. Ecol.*
577 *Environ.* 7, 21–28. doi:10.1890/080025

578 de Cáceres, M., Martínez-Vilalta, J., Coll, L., Llorens, P., Casals, P., Poyatos, R., Brotons, L. 2015.
579 Coupling a water balance model with forest inventory data to predict drought stress: the role of forest
580 structural changes vs. climate changes. *Agricultural and Forest Meteorology.* 213: 77-90.

581 de Groot, R.S., Alkemade, R., Braat, L., Hein, L., Willemen, L., 2010. Challenges in integrating the
582 concept of ecosystem services and values in landscape planning, management and decision making.
583 *Ecol. Complex.* 7, 260–272. doi:10.1016/j.ecocom.2009.10.006

584 de-Miguel, S., Bonet, J. A., Pukkala, T., Martínez de Aragón, J. 2014. Impact of forest management
585 intensity on landscape-level mushroom productivity: A regional model-based scenario analysis. *For.*
586 *Ecol. Manage.* 330, 218–227. <http://doi.org/10.1016/j.foreco.2014.07.014>

587 Deng, X., Li, Z., Gibson, J., 2016. A review on trade-off analysis of ecosystem services for sustainable
588 land-use management. *J. Geogr. Sci.* 26, 953–968. doi:10.1007/s11442-016-1309-9

589 Díaz-Varela, E., Rocas-Díaz, J.V., Álvarez-Álvarez, P., 2016. Detection of landscape heterogeneity at
590 multiple scales: Use of the Quadratic Entropy Index. *Landsc. Urban Plan.* 153, 149–159.
591 doi:10.1016/j.landurbplan.2016.05.004

592 Diaz-Varela, E.R., Marey-Pérez, M.F., Álvarez-Álvarez, P. 2009. Use of simulated and real data to
593 identify heterogeneity domains in scale-divergent forest landscapes. *For. Ecol. Manage.* 258, 2490–
594 2500. doi:10.1016/j.foreco.2009.09.005

595 Doblas-Miranda, E., Rovira, P., Brotons, L., Martínez-Vilalta, J., Retana, J., Pla, M., & Vayreda, J. 2013.
596 Soil carbon stocks and their variability across the forests, shrublands and grasslands of peninsular
597 Spain. *Biogeosciences*, 10(12), 8353–8361. <http://doi.org/10.5194/bg-10-8353-2013>

598 Egoh, B., Reyers, B., Rouget, M., Bode, M., Richardson, D., 2009. Spatial congruence between
599 biodiversity and ecosystem services in South Africa. *Biol. Conserv.* 142, 553–562.
600 doi:10.1016/j.biocon.2008.11.009

601 Egoh, B., Reyers, B., Rouget, M., Richardson, D., Lemaitre, D., Vanjaarsveld, a, 2008. Mapping
602 ecosystem services for planning and management. *Agric. Ecosyst. Environ.* 127, 135–140.
603 doi:10.1016/j.agee.2008.03.013

604 Eigenbrod, F., Armsworth, P.R., Anderson, B.J., Heinemeyer, A., Gillings, S., Roy, D.B., Thomas, C.D.,
605 Gaston, K.J., 2010. The impact of proxy-based methods on mapping the distribution of ecosystem
606 services. *J. Appl. Ecol.* 47, 377–385. doi:10.1111/j.1365-2664.2010.01777.x

607 ESRI 2011. ArcGIS Desktop: Release 10.2. Redlands, CA: Environmental Systems Research Institute.

608 Fortin, M.J., James, P.M. a, MacKenzie, A., Melles, S.J., Rayfield, B., 2012. Spatial statistics, spatial
609 regression, and graph theory in ecology. *Spat. Stat.* 1, 100–109. doi:10.1016/j.spasta.2012.02.004

610 Geijzendorffer, I.R., Martín-López, B., Roche, P.K., 2015. Improving the identification of mismatches in
611 ecosystem services assessments. *Ecol. Indic.* 52, 320–331. doi:10.1016/j.ecolind.2014.12.016

612 Getis, A., Ord, J.K., 1992. The Analysis of Spatial Association. *Geogr Anal* 24:189–206. doi:
613 10.1111/j.1538-4632.1992.tb00261.x

614 Grêt-Regamey, A., Weibel, B., Bagstad, K., Ferrari, M., Geneletti, D., Klug, H., Schirpke, U., Tappeiner,
615 U., 2014. On the Effects of Scale for Ecosystem Services Mapping. *PLoS One* 1–26.
616 doi:10.1371/journal.pone.0112601

617 Guerra, C.A., Maes, J., Geijzendorffer, I., Metzger, M.J., 2016. An assessment of soil erosion prevention
618 by vegetation in Mediterranean Europe: Current trends of ecosystem service provision. *Ecol. Indic.*
619 60, 213–222. doi:10.1016/j.ecolind.2015.06.043

620 Harrison, P. a., Berry, P.M., Simpson, G., Haslett, J.R., Blicharska, M., Bucur, M., Dunford, R., Egoh, B.,

621 Garcia-Llorente, M., Geamănă, N., Geertsema, W., Lommelen, E., Meiresonne, L., Turkelboom, F.,
622 2014. Linkages between biodiversity attributes and ecosystem services: A systematic review.
623 *Ecosyst. Serv.* 9, 191–203. doi:10.1016/j.ecoser.2014.05.006

624 Haines-Young, R., Potschin, M., 2013. Common International Classification of Ecosystem Services
625 (CICES): Consultation on Version 4, August-December 2012. EEA Framework Contract No
626 EEA/IEA/09/003 (Download at www.cices.eu or www.nottingham.ac.uk/cem)

627 Hein, L., van Koppen, K., de Groot, R.S., van Ierland, E.C., 2006. Spatial scales, stakeholders and the
628 valuation of ecosystem services. *Ecol. Econ.* 57, 209–228. doi:10.1016/j.ecolecon.2005.04.005

629 Herrando-Moraira, S. Franch, M., Anton, M., Garcia, D., Villero, D., Brotons, Ll., Herrando, S., 2016. Is
630 there any relation between observers' preference to visit a given site and its conservation value? An
631 analysis with casual data from Ornitho.cat for birds. In Busch, M. & Gedeon, K.
632 (Eds.) *BirdNumbers 2016: Birds in a changing world. Programme and Abstracts of the 20th*
633 *conference of the European Bird Census Council. Dachverband Deutscher Avifaunisten, Münster.*

634 Homolová, L., Schaepman, M.E., Bello, F. De, Thuiller, W., Lavorel, S., 2014. Comparison of remote
635 sensing and plant trait-based modelling to predict ecosystem services in subalpine grasslands.
636 *Ecosphere* 5 (8): 100.

637 Hou, Y., Burkhard, B., Müller, F., 2013. Uncertainties in landscape analysis and ecosystem service
638 assessment. *J. Environ. Manage.* 127 Suppl, S117–31. doi:10.1016/j.jenvman.2012.12.002

639 IPBES, Intergovernmental Platform on Biodiversity and Ecosystem Services. 2012. <http://www.ipbes.net/>

640 Kandziora, M., Burkhard, B., Müller, F., 2013. Mapping provisioning ecosystem services at the local
641 scale using data of varying spatial and temporal resolution. *Ecosyst. Serv.* 4, 47–59.
642 doi:10.1016/j.ecoser.2013.04.001

643 Konarska, K.M., Sutton, P.C., Castellon, M., 2002. Evaluating scale dependence of ecosystem service
644 valuation: A comparison of NOAA-AVHRR and Landsat TM datasets. *Ecol. Econ.* 41, 491–507.
645 doi:10.1016/S0921-8009(02)00096-4

646 Kremen, C., 2005. Managing ecosystem services: what do we need to know about their ecology? *Ecol.*
647 *Lett.* 8, 468–79. doi:10.1111/j.1461-0248.2005.00751.x

648 Lautenbach, S., Jungandreas, A., Blanke, J., Lehsten, V., Mühlner, S., Kühn, I., Volk, M., 2017. Trade-
649 offs between plant species richness and carbon storage in the context of afforestation? Examples
650 from afforestation scenarios in the Mulde Basin, Germany. *Ecol. Indic.* 73, 139–155.
651 doi:10.1016/j.ecolind.2016.09.035

652 LCMC, Land Cover Map of Catalonia, 2009. Generalitat de Catalunya. CREAL, Universidad Autónoma
653 de Barcelona. <http://www.creaf.uab.es/mcsc/esp/index.htm>

654 Levin, S.A., 1992. The problem of pattern and scale in ecology. *Ecology* 73, 1943–1967.

655 Liu, Y., Bi, J., Lv, J., Ma, Z., Wang, C., 2017. Spatial multi-scale relationships of ecosystem services: A
656 case study using a geostatistical methodology. *Sci. Rep.* 7, 1–12. doi:10.1038/s41598-017-09863-1

657 Locatelli, B., Imbach, P., Wunder, S., 2013. Synergies and trade-offs between ecosystem services in
658 Costa Rica. *Environ. Conserv.* 41, 27–36. doi:10.1017/S0376892913000234

659 Maes, J., Egoh, B., Willemen, L., Liqueste, C., Vihervaara, P., Schägner, J.P., Grizzetti, B., Drakou, E.G.,
660 Notte, A. La, Zulian, G., Bouraoui, F., Luisa Paracchini, M., Braat, L., Bidoglio, G., 2012. Mapping
661 ecosystem services for policy support and decision making in the European Union. *Ecosyst. Serv.* 1,
662 31–39. doi:10.1016/j.ecoser.2012.06.004

663 MAGRAMA, Ministerio de Agricultura, Alimentación y Medio Ambiente. 1997-2007. Segundo y Tercer
664 Inventario Forestal Nacional. Gobierno de España. [online 15 July 2015] URL:
665 [http://www.magrama.gob.es/es/biodiversidad/servicios/banco-datos-naturaleza/informacion-](http://www.magrama.gob.es/es/biodiversidad/servicios/banco-datos-naturaleza/informacion-disponible/index_inventario_forestal.aspx)
666 [disponible/index_inventario_forestal.aspx](http://www.magrama.gob.es/es/biodiversidad/servicios/banco-datos-naturaleza/informacion-disponible/index_inventario_forestal.aspx)

667 Martínez-Harms, M.J., Quijas, S., Merenlender, A.M., Balvanera, P., 2016. Enhancing ecosystem
668 services maps combining field and environmental data. *Ecosyst. Serv.* 22, 32–40.
669 doi:10.1016/j.ecoser.2016.09.007

670 Martín-López, B., Gómez-Baggethun, E., Lomas, P.L., Montes, C., 2009. Effects of spatial and temporal
671 scales on cultural services valuation. *J. Environ. Manage.* 90, 150–159.
672 doi:10.1016/j.jenvman.2008.03.013

673 MEA, Millenium Ecosystem Assessment. 2005. *Ecosystems and Human Well-being: Current State and*
674 *Trends.* Island Press, Washington, DC Millennium Ecosystem Assessment.

675 Moran, P.A.P., 1948. The interpretation of statistical maps. *J. R. Stat. Soc.* 10, 243–251.

676 Mouchet, M. a., Lamarque, P., Martín-López, B., Crouzat, E., Gos, P., Byczek, C., Lavorel, S., 2014. An
677 interdisciplinary methodological guide for quantifying associations between ecosystem services.
678 *Glob. Environ. Chang.* 28, 298–308. doi:10.1016/j.gloenvcha.2014.07.012

679 Onaindia, M., Fernández de Manuel, B., Madariaga, I., Rodríguez-Loinaz, G., 2013. Co-benefits and

680 trade-offs between biodiversity, carbon storage and water flow regulation. *For. Ecol. Manage.* 289,
681 1–9. doi:10.1016/j.foreco.2012.10.010

682 Ord, J.K., Getis, A. 1995. Local Spatial Autocorrelation Statistics: Distributional Issues and an
683 Application. *Geogr Anal* 27:286–306. doi: 10.1111/j.1538-4632.1995.tb00912.x

684 Potschin, M.B. Haynes-Young R.H. 2011. Ecosystem services: Exploring a geographical perspective.
685 *Prog. Phys. Geogr*, 35: 575-594. doi.org/10.1177/0309133311423172

686 Raudsepp-Hearne, C., Peterson, G.D., 2016. Scale and ecosystem services: how do observation,
687 management, and analysis shift with scale — lessons from Québec. *Ecology and Society*, 21: 16.
688 doi:10.5751/ES-08605-210316

689 Raudsepp-Hearne, C., Peterson, G.D., Bennett, E.M., 2010. Ecosystem service bundles for analyzing
690 tradeoffs in diverse landscapes. *Proc. Natl. Acad. Sci. U. S. A.* 107, 5242–7.
691 doi:10.1073/pnas.0907284107

692 Renard, D., Rhemtulla, J.M., Bennett, E.M., 2015. Historical dynamics in ecosystem service bundles.
693 *Proc. Natl. Acad. Sci. U. S. A.* 112, 13411–13416. doi:10.1073/pnas.1502565112

694 Rocés-Díaz, J. V., Díaz-Varela, E.R., Álvarez-Álvarez, P., 2014. Analysis of spatial scales for ecosystem
695 services: Application of the lacunarity concept at landscape level in Galicia (NW Spain). *Ecol. Indic.*
696 36, 495–507. doi:10.1016/j.ecolind.2013.09.010

697 Rocés-Díaz, J.V., Burkhard, B, Kruse, M, Muller, F, Diaz-Varela, E, Alvarez, P. 2017a. Use of ecosystem
698 information derived from forest thematic maps for spatial analysis of ecosystem services in
699 northwestern Spain. *Landscape Ecol. Eng.* 13:45–57. doi.org/10.1007/s11355-016-0298-2

700 Rocés-Díaz, J.V., Vayreda, J., Banque-Casanovas, M., Cuso, M., Anton, M., Bonet, J.A., Brotons, Ll., De
701 Caceres, M., Herrando, S., Martinez de Aragon, J., de-Miguel, S., Martinez-Vilalta, J. 2017b.
702 Assessing the distribution of forest ecosystem services in a highly populated Mediterranean region
703 (Submitted to *Ecological Indicators*)

704 Rodríguez, J. P., T. D. Beard, Jr., E. M. Bennett, G. S. Cumming, S. Cork, J. Agard, A. P. Dobson, and G.
705 D. Peterson. 2006. Trade-offs across space, time, and ecosystem services. *Ecology and Society* 11(1):
706 28. URL: <http://www.ecologyandsociety.org/vol11/iss1/art28/>

707 Rodríguez-Loinaz, G., Amezcaga, I., Onaindia, M., 2013. Use of native species to improve carbon
708 sequestration and contribute towards solving the environmental problems of the timberlands in
709 Biscay, northern Spain. *J. Environ. Manage.* 120, 18–26. doi:10.1016/j.jenvman.2013.01.032

710 Rodríguez-Loinaz, G., Alday, J.G., Onaindia, M., 2015. Multiple ecosystem services landscape index: A
711 tool for multifunctional landscapes conservation. *J. Environ. Manage.* 147, 152–163.
712 doi:10.1016/j.jenvman.2014.09.001

713 Sardà-Palomera, F., Brotons, L., Villero, D., Sierdsema, H., Newson, S.E., Jiguet, F. 2012. Mapping from
714 heterogeneous biodiversity monitoring data sources. *Biodivers Conserv* 21, 2927-2948.
715 <https://doi.org/10.1007/s10531-012-0347-6>

716 Schröter, M., Remme, R.P., Sumarga, E., Barton, D.N., Hein, L., 2015. Lessons learned for spatial
717 modelling of ecosystem services in support of ecosystem accounting. *Ecosyst. Serv.* 13, 64-69.
718 doi.org/10.1016/j.ecoser.2014.07.003

719 Seppelt, R., Dormann, C.F., Eppink, F. V., Lautenbach, S., Schmidt, S., 2011. A quantitative review of
720 ecosystem service studies: approaches, shortcomings and the road ahead. *J. Appl. Ecol.* 48, 630–636.
721 doi:10.1111/j.1365-2664.2010.01952.x

722 Strassburg, B.B.N., Kelly, A., Balmford, A., Davies, R.G., Gibbs, H.K., Lovett, A., Miles, L., Orme,
723 C.D.L., Price, J., Turner, R.K., Rodrigues, A.S.L., 2010. Global congruence of carbon storage and
724 biodiversity in terrestrial ecosystems. *Conserv. Lett.* 3, 98–105. doi:10.1111/j.1755-
725 263X.2009.00092.x

726 Tallis, H., Polasky, S., 2009. Mapping and valuing ecosystem services as an approach for conservation
727 and natural-resource management. *Ann. N. Y. Acad. Sci.* 1162, 265–83. doi:10.1111/j.1749-
728 6632.2009.04152.x

729 TEEB, The Economics of Ecosystems and Biodiversity. 2010. <http://www.teebweb.org/>

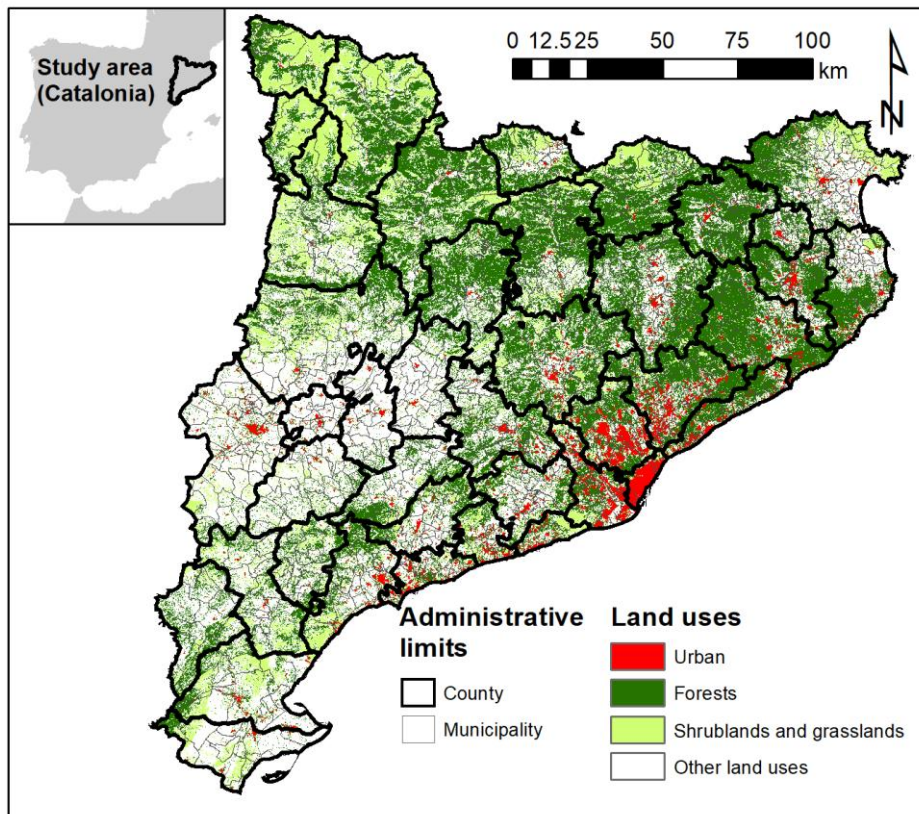
730 Tolvanen, H., Rönkä, M., Vihervaara, P., Kampainen, M., Arzel, C., Aarras, N., Thessler, S., 2014.
731 Spatial information in ecosystem service assessment: data applicability in the cascade model context.
732 *J. Land Use Sci.* 1–18. doi:10.1080/1747423X.2014.947642

733 Tomscha, S.A., Gergel, S.A., 2016. Ecosystem service trade-offs and synergies misunderstood without
734 landscape history. *Ecology and Society* 21, 43. doi.org/10.5751/ES-08345-210143

735 Uuemaa, E., Antrop, M., Marja, R., 2009. Landscape Metrics and Indices : An Overview of Their Use in
736 Landscape Research Imprint / Terms of Use. Landscape 1–28.

737 Uuemaa, E., Mander, Ü., Marja, R., 2013. Trends in the use of landscape spatial metrics as landscape
738 indicators: A review. *Ecol. Indic.* 28, 100–106. doi:10.1016/j.ecolind.2012.07.018

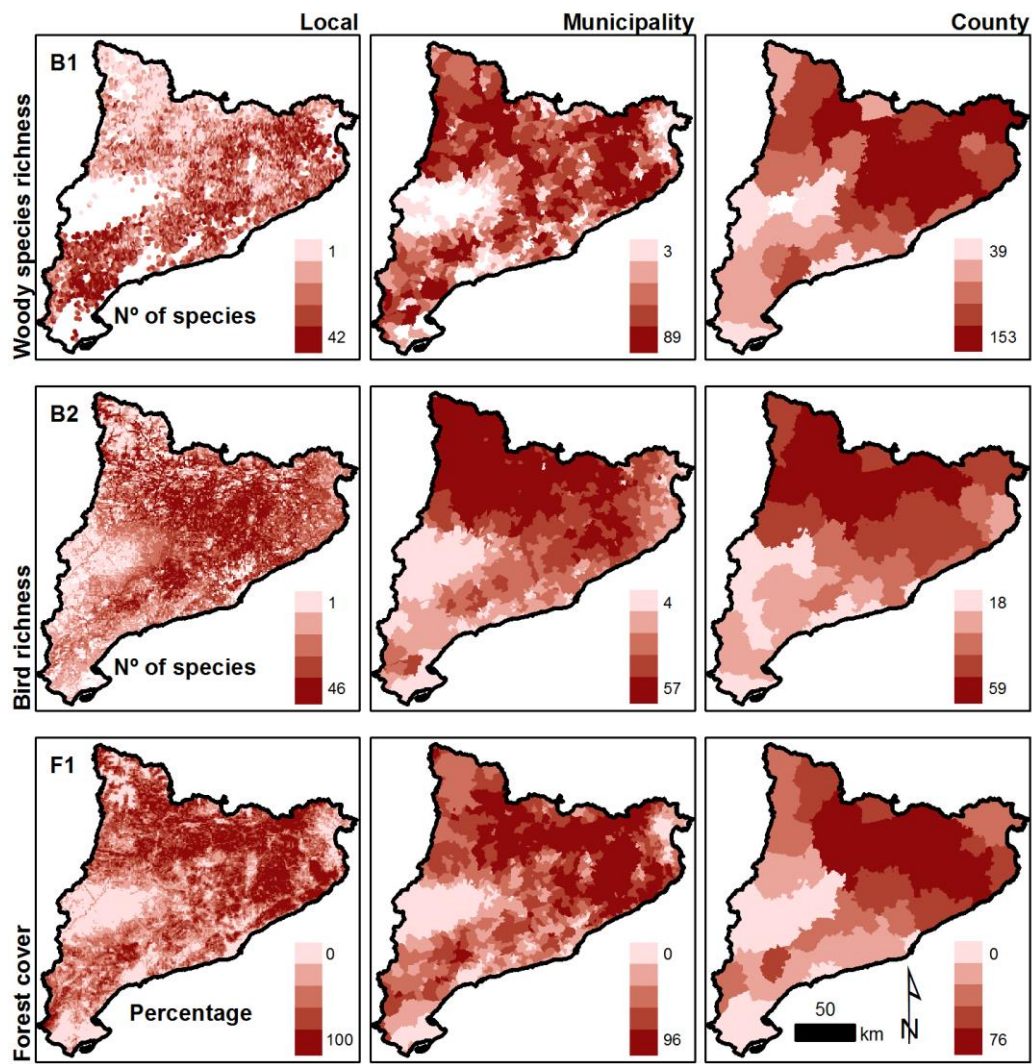
739 Vayreda, J., Gracia, M., Canadell, J.G., Retana, J., 2012. Spatial Patterns and Predictors of Forest Carbon
740 Stocks in Western Mediterranean. *Ecosystems* 15, 1258–1270. doi:10.1007/s10021-012-9582-7
741 Wagner, H.H., Fortin, M.J., 2005. Spatial analysis of landscapes: Concepts and statistics. *Ecology* 86,
742 1975–1987. doi:10.1890/04-0914
743 Wolff, S., Schulp, C.J.E., Verburg, P.H., 2015. Mapping ecosystem services demand: A review of current
744 research and future perspectives. *Ecol. Indic.* 55, 159-171. doi.org/10.1016/j.ecolind.2015.03.016
745 Xu, S., Liu, Y., Wang, X., Zhang, G., 2017. Scale effect on spatial patterns of ecosystem services and
746 associations among them in semi-arid area: A case study in Ningxia Hui Autonomous Region, China.
747 *Sci. Total Environ.* 598, 297–306. doi:10.1016/j.scitotenv.2017.04.009
748 Yahdjian, L., Sala, O.E., Havstad, K.M., 2015. Rangeland ecosystem services: Shifting focus from supply
749 to reconciling supply and demand. *Front. Ecol. Environ.* 13, 44–51. doi:10.1890/140156
750 Yang, G., Ge, Y., Xue, H., Yang, W., Shi, Y., Peng, C., Du, Y., Fan, X., Ren, Y., Chang, J., 2015. Using
751 ecosystem service bundles to detect trade-offs and synergies across urban–rural
752 complexes. *Landsc. Urban Plan.* 136, 110-121. doi:10.1016/j.landurbplan.2014.12.006
753



754

755 Figure 1. Location of the study area, including the different administrative limits (municipality and
 756 county) and main types of land cover (LCMC, 2009).

757

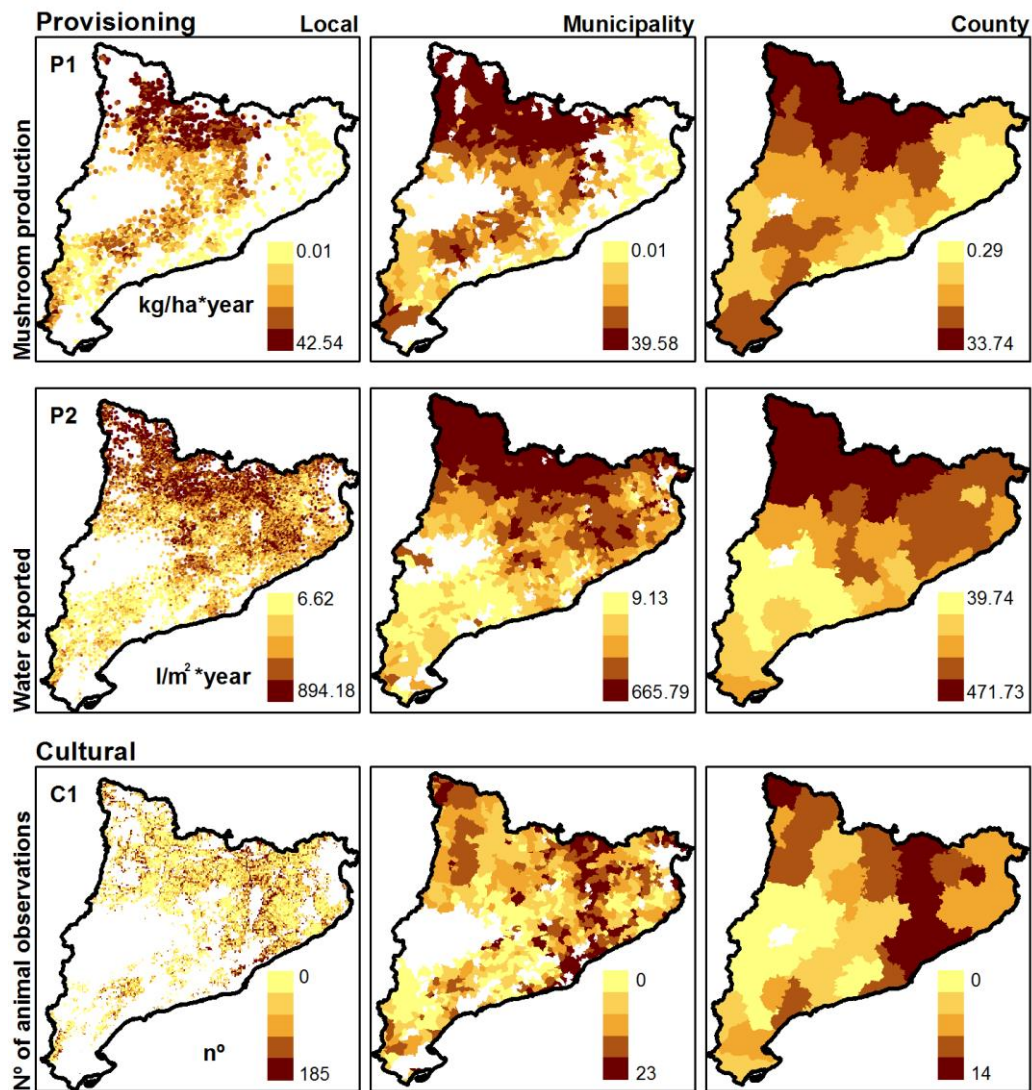


758

759 Figure 2. Spatial patterns of biodiversity (woody species richness (B1) and bird richness (B2)) and forest
 760 cover (F1) in Catalonia at the three levels of analysis (local, municipality and county) used in this work.
 761 Colour intensity indicates increasing, 20% percentile classes.

762

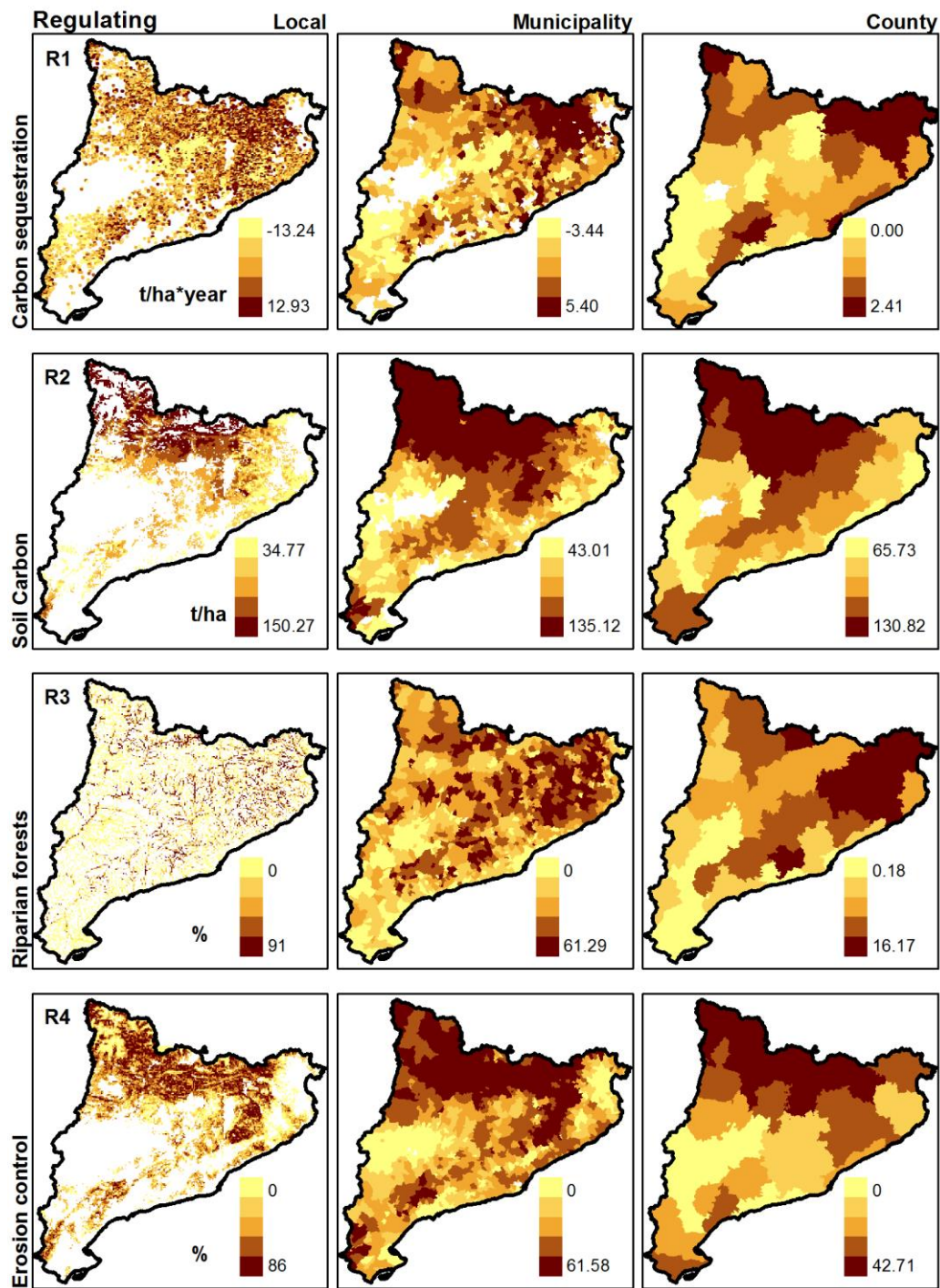
763



764

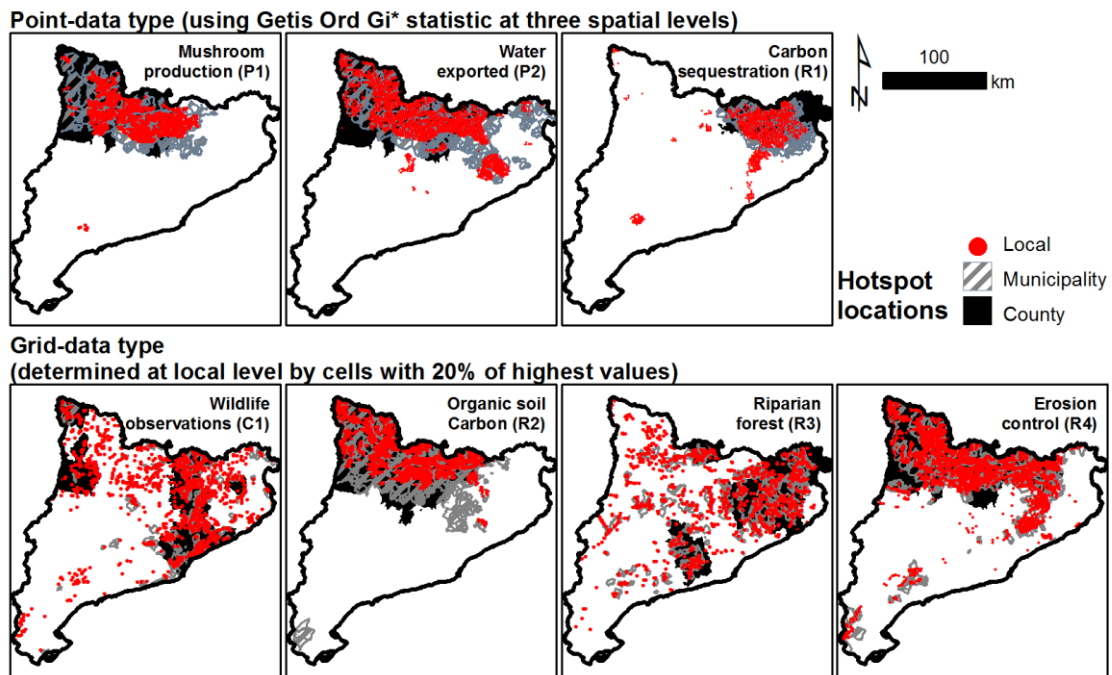
765 Figure 3. Spatial patterns of provisioning ES (mushroom production (P1) and water exported (P2)) and
 766 cultural ES (animal observations (C1)) indicators in Catalonia at the three levels of analysis (local,
 767 municipality and county) used in this work. Colour intensity indicates increasing, 20% percentile classes.

768



769

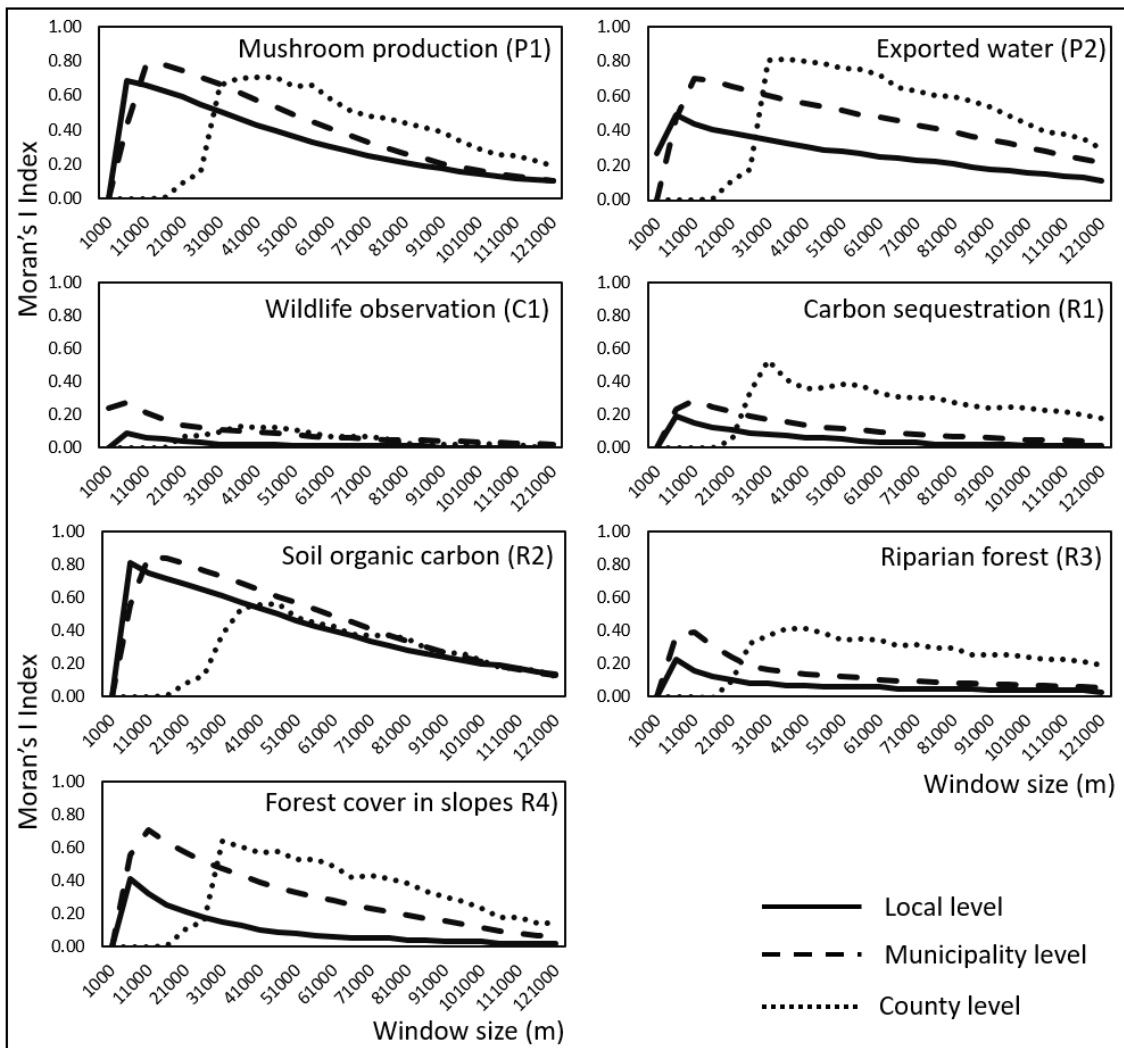
770 Figure 4. Spatial patterns of regulating ES (carbon sequestration (R1), soil organic carbon (R2), riparian
 771 forest (R3) and erosion control (R4)) indicators in Catalonia at the three levels of analysis (local,
 772 municipality and county) used in this work. Colour intensity indicates increasing, 20% percentile classes.



773

774 Figure 5. Spatial patterns of hotspots of ES indicators at the three levels of analysis (local, municipality
 775 and county) used in this work.

776



777

778 Figure 6. Results of Incremental Spatial Autocorrelation analysis that shows values of Moran's I index
 779 (vertical axis) from increasing sizes of windows (horizontal axis; meters).

780

781

782 Tables

783

784 Table 1. Description of the main variables used in this work. The transformations used to normalize the
 785 distribution of the variables were: root square (sqr) or logarithmic (ln). Normalization showed if each
 786 variable was (or not) normalized. Two different types of correlation coefficients were used: Pearson and
 787 Spearman (when the variable could not be normalized). Further details are provided in the text.

Category	ES/ variable	Indicator	Code	Units	Data	N/original resolution	Transform.	Normal.	Sources
Forest Biodiversity	Tree diversity	Woody species richness	B1	N°	Point (SNFI plot)	11,288 plots	sqr(x)	Yes	MAGRAMA (1997-2007)
	Bird diversity	Forest bird richness	B2	N°	Grid	1,000 m	-	No	ICO (2014)
Forest cover	Forest cover	Forest cover	F1	%	Grid	25 m	ln(x+0.1)	Yes	LCMC (2009)
Provisioning Ecosystem Services	Food provision	Mushrooms	P1	kg/ha/year	Point (SNFI plot)	3,272 plots	sqr(x)	Yes	de-Miguel et al. (2014)
	Water provision	Exported water	P2	L/m ² /year	Point (SNFI plot)	11,261 plots	sqr(x)	Yes	de Caceres et al. (2015)
Cultural Ecosystem Services	Recreation	Wildlife observation	C1	N°	Grid	1,000 m	ln(x+0.1)	No	ICO (2014)
Regulating Ecosystem Services	Climate regulation	Carbon sequestration	R1	t/ha/year	Point (SNFI plot)	8,726 plots	sqr(x)	Yes	Vayreda et al. (2012)
	Soil fertility	Soil organic Carbon	R2	t/ha	Grid	200 m	-	Yes	Doblas-Miranda et al. (2013)
	Flood regulation	Riparian forest cover	R3	%	Grid	25 m	ln(x+0.1)	No	LCMC (2009)
	Erosion control	Forest cover in slopes	R4	%	Grid	25 m	sqr(x)	No	LCMC (2009)

788

789

790 Table 2. Percentage of hotspots at local level included inside the hotspot areas at municipality and county
 791 levels.

		Ecosystem Services indicators ^a							
		Direction of change	P1	P2	C1	R1	R2	R3	R4
% of local hotspots included at higher levels	Local to municipality		98.1	92.9	38.8	74.2	98.9	53.7	88.0
	Local to county		72.8	75.3	52.2	61.3	97.2	48.2	71.5

792 ^a Ecosystem services indicators: Mushrooms (P1), Exported water (P2), Wildlife observation (C1), Carbon sequestration (R1), Soil
 793 organic Carbon (R2), Riparian forest cover (R3), Erosion control (R4).
 794

795 Table 3. Correlation coefficients between pairs of environmental variables and ES indicators (*: p-value <
 796 0.05). Red colours: significant negative relationships and green colours: significant positive relations

Local		B2	F1	P1	P2	C1	R1	R2	R3	R4
Woody sp. rich.	B1	0.02	-0.06*	-0.41*	-0.47*	-0.02	-0.01	-0.44*	0.07*	-0.32*
Bird richness	B2	1	0.79*	0.20*	0.08*	0.28*	0.18*	0.05*	0.13*	0.12*
Forest cover	F1		1	0.42*	0.12*	-0.18*	0.24*	0.06*	0.18*	0.47*
Mushroom prod.	P1			1	0.53*	0.03	0.13*	0.82*	-0.01	0.53*
Exported water	P2				1	0.07*	0.11*	0.48*	0.05*	0.34*
Animal obs.	C1					1	0.06*	0.06*	0.09*	0.02
Carbon seq.	R1						1	0.04	0.13*	0.12*
Soil Carbon	R2							1	-0.14*	0.49*
Riparian forest	R3								1	0.03
Erosion control	R4									1
Municipality		B2	F1	P1	P2	C1	R1	R2	R3	R4
Woody sp. rich.	B1	0.50*	0.69*	0.23*	0.10	-0.06	0.14*	0.28*	0.17*	0.55*
Bird richness	B2	1	0.73*	0.54*	0.65*	0.24*	0.29*	0.69*	0.47*	0.70*
Forest cover	F1		1	0.37*	0.34*	0.06	0.33*	0.47*	0.37*	0.79*
Mushroom prod.	P1			1	0.43*	0.13*	0.00	0.81*	0.04	0.61*
Exported water	P2				1	0.30*	0.14*	0.48*	0.27*	0.39*
Animal obs.	C1					1	0.23*	0.13*	0.00	0.19*
Carbon seq.	R1						1	0.23*	0.24*	0.24*
Soil Carbon	R2							1	0.23*	0.56*
Riparian forest	R3								1	0.20*
Erosion control	R4									1
County		B2	F1	P1	P2	C1	R1	R2	R3	R4
Woody sp. rich.	B1	0.71*	0.83*	0.16	0.58*	0.38	0.56*	0.35	0.69*	0.56*
Bird richness	B2	1	0.81*	0.57*	0.83*	0.47*	0.44	0.70*	0.61*	0.80*
Forest cover	F1		1	0.25	0.69*	0.54*	0.45	0.47*	0.72*	0.69*
Mushroom prod.	P1			1	0.53*	0.09	0.15	0.88*	0.18	0.65*
Exported water	P2				1	0.51*	0.49*	0.64*	0.50*	0.75*
Animal obs.	C1					1	0.58*	0.31	0.19	0.54*
Carbon seq.	R1						1	0.38	0.46*	0.51*
Soil Carbon	R2							1	0.36	0.75*
Riparian forest	R3								1	0.42
Erosion control	R4									1

797
 798

799
800
801
802
803
804

Table 4. Linear models for the seven ES indicators as dependent variables and the three biodiversity and forest cover indicators as independent variables. Local and municipality level analyses correspond to mixed-effects models and county level analyses were conducted using simple linear models. Local level models used municipality (nested in county) and county as random factors; municipality level models used county as random factor. Significance level *** <0.001; **<0.01, *<0.05; -<0.1. R2m: marginal R-squared, R2C: conditional R-squared, R2: R-squared; SE: standard error.

Indicators	Spatial level	Intercept		Woody sp. richness (B1)		Bird richness (B2)		Forest cover (F1)		R-squared		
		Coef.	SE	Coef.	SE	Coef.	SE	Coef.	SE	R2m	R2c	R2
Mushrooms Production (P1)	Local	0.112	0.952	-0.338***	0.064	0.001	0.004	0.914***	0.202	0.027	0.798	-
	Municip.	-1.889***	0.517	0.049	0.041	0.048***	0.011	0.655***	0.116	0.153	0.808	-
	County	-1.676	1.3	-0.356-	0.194	0.209***	0.038	-0.553	0.564	-	-	0.44
Water Exported (P2)	Local	18.062***	3.831	-1.843***	0.256	0.031*	0.014	-0.13	0.832	0.062	0.53	-
	Municip.	5.087***	1.273	-0.01	0.107	0.179***	0.027	-0.003	0.296	0.149	0.753	-
	County	-6.791*	2.741	-0.793	0.408	0.523***	0.081	0.628	1.189	-	-	0.64
Wildlife observations (C1)	Local	8.187***	1.437	-0.16	0.096	0.034***	0.005	-2.114***	0.313	0.132	0.356	-
	Municip.	-1.43***	0.416	0.021	0.042	0.042***	0.009	-0.12	0.112	0.114	0.341	-
	County	-1.404-	0.784	-0.16	0.117	0.019	0.023	0.646-	0.34	-	-	0.14
Carbon sequestration (R1)	Local	1.086**	0.416	-0.026	0.028	0.003*	0.002	0.03	0.09	0.01	0.184	-
	Municip.	0.369**	0.141	-0.024	0.014	0.009***	0.003	0.099	0.039	0.1	0.363	-
	County	0.12	0.288	0.008	0.043	0.014	0.008	0.024	0.125	-	-	0.14
Soil organic Carbon (R2)	Local	81.911***	7.506	-4.462***	0.493	-0.069*	0.027	5.136**	1.573	0.044	0.871	-
	Municip.	50.872***	4.998	-1.395***	0.404	0.775***	0.104	3.114**	1.113	0.161	0.817	-
	County	19.399	13.05	-6.946**	1.94	2.613***	0.385	2.276	5.661	-	-	0.61
Riparian forest (R3)	Local	3.146	2.321	0.296	0.154	0.052***	0.009	-1.394**	0.506	0.085	0.317	-
	Municip.	-2.389***	0.649	0.106	0.065	0.059***	0.014	0.124	0.181	0.149	0.392	-
	County	-3.804***	0.833	0.294	0.124	0.027	0.025	0.307	0.362	-	-	0.53
Erosion Control (R4)	Local	-10.93***	1.89	-0.233	0.128-	0.024***	0.007	3.551***	0.411	0.07	0.62	-
	Municip.	-5.162***	0.661	0.102	0.061-	0.055***	0.014	1.708***	0.170	0.381	0.677	-
	County	-2.185*	0.918	-0.078	0.137	0.119***	0.027	0.456	0.398	-	-	0.59

805
806
807

808 *The spatial level of analysis affects the patterns of forest*
809 *ecosystem services supply and their relationships*

810

811

812 **Supplementary material.**

813

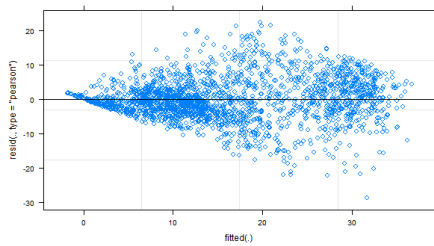
814 **Appendix A. Residuals of the linear models fitted at different spatial scales.**

815

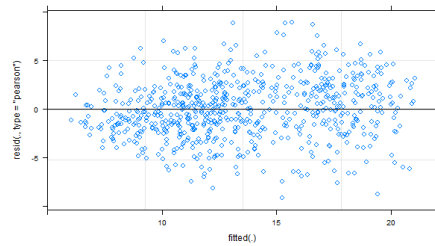
816

Local level (mixed-effect models)

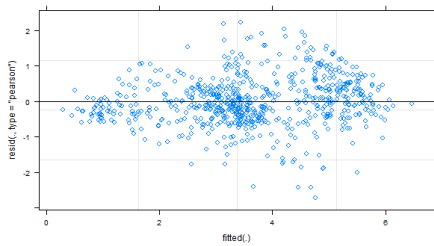
P1 (mushroom production)



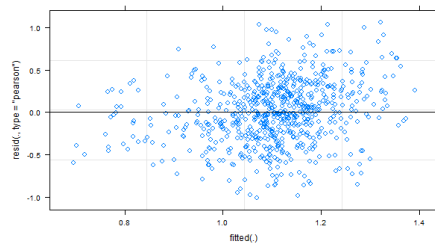
P2 (exported water)



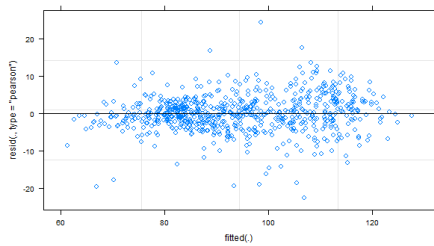
C1 (wildlife observation)



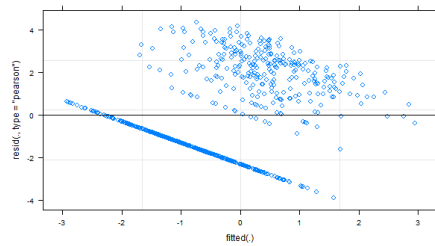
R1 (carbon sequestration)



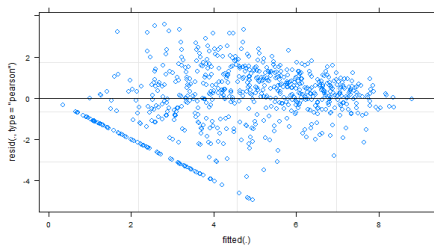
R2 (soil organic carbon)



R3 (riparian forest)



R4 (forest cover in slopes)



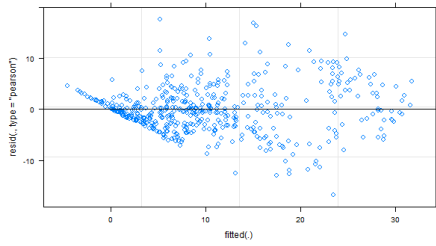
817 Figure S1. Residuals of mixed-effect models for ES indicators at local level.

818

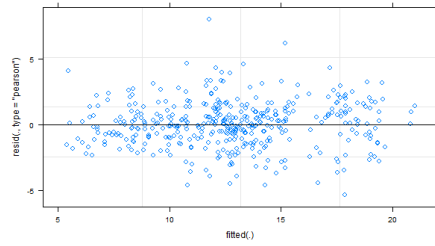
819

Municipality level (mixed-effect models)

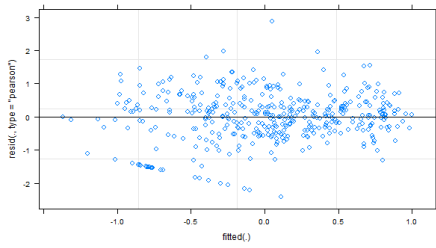
P1 (mushroom production)



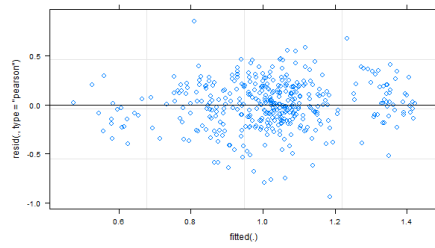
P2 (exported water)



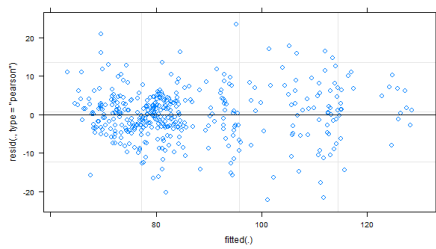
C1 (wildlife observation)



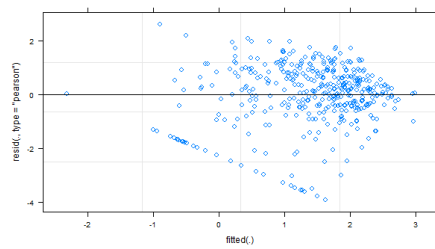
R1 (carbon sequestration)



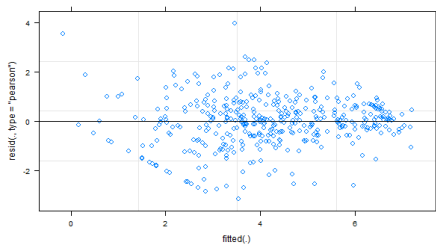
R2 (soil organic carbon)



R3 (riparian forest)



R4 (forest cover in slopes)



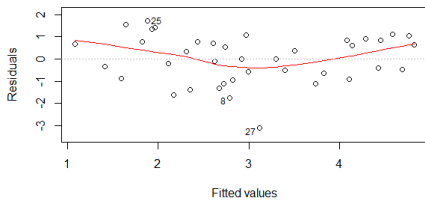
820 Figure S2. Residuals of mixed-effect models for ES indicators at municipality level.

821

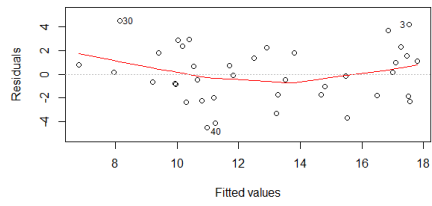
822

County level (general linear models)

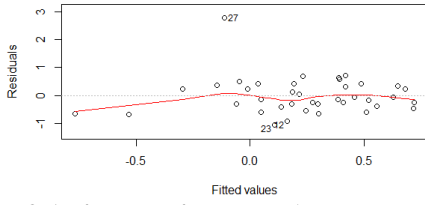
P1 (mushroom production)



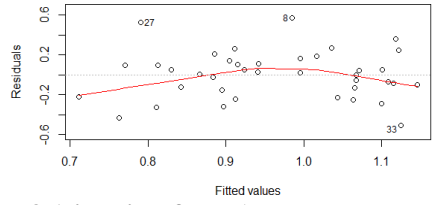
P2 (exported water)



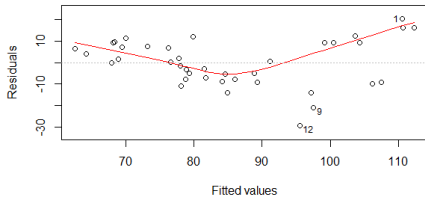
C1 (wildlife observation)



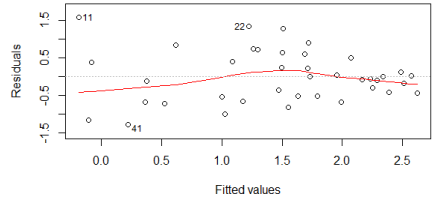
R1 (carbon sequestration)



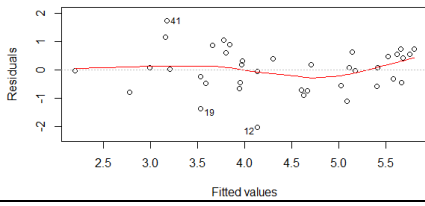
R2 (soil organic carbon)



R3 (riparian forest)



R4 (forest cover in slopes)



823 Figure S3. Residuals of general lineal models for ES indicators at county level.
824

825 **Appendix B. Results of spatial autocorrelation tests.**

826

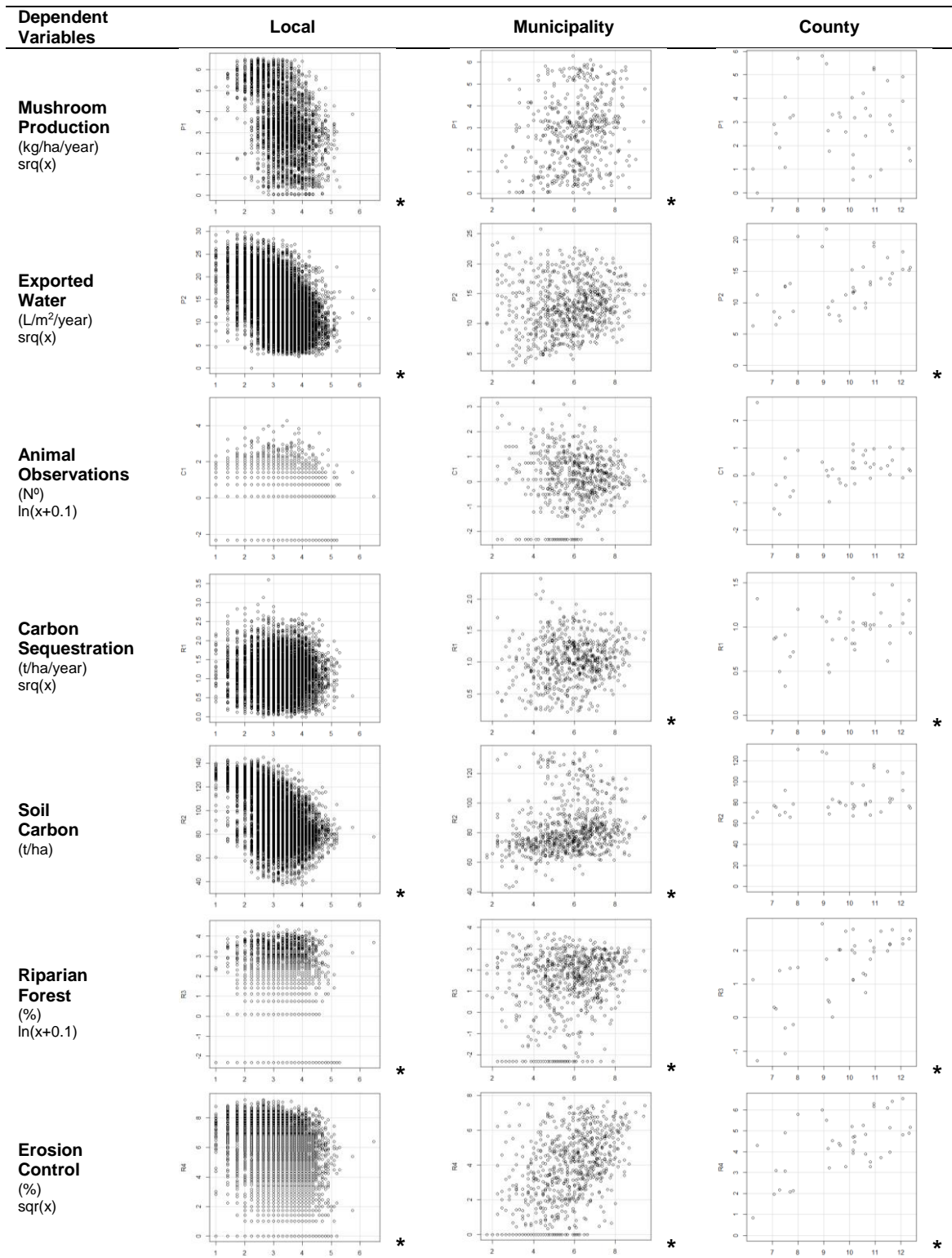
827 Table S1. Spatial autocorrelation results.

ES indicator	Level	Moran's I coefficient	Z score	p-value
Mushroom production	Local	0.61	261.24	< 0.01
	Municipality	0.63	45.24	< 0.01
	County	0.51	3.85	< 0.01
Exported water	Local	0.41	284.78	< 0.01
	Municipality	0.53	40.43	< 0.01
	County	0.69	5.31	< 0.01
Wildlife observation	Local	0.07	74.07	< 0.01
	Municipality	0.45	23.69	< 0.01
	County	0.17	2.82	< 0.01
Carbon sequestration	Local	0.19	127.58	< 0.01
	Municipality	0.26	19.13	< 0.01
	County	0.46	3.59	< 0.01
Soil organic carbon	Local	0.37	386.28	< 0.01
	Municipality	0.61	19.19	< 0.01
	County	0.52	4.03	< 0.01
Riparian forest	Local	0.36	100.11	< 0.01
	Municipality	0.49	35.97	< 0.01
	County	0.35	2.84	< 0.01
Forest cover in slopes	Local	0.41	423.21	< 0.01
	Municipality	0.56	41.41	< 0.01
	County	0.52	4.11	< 0.01

828

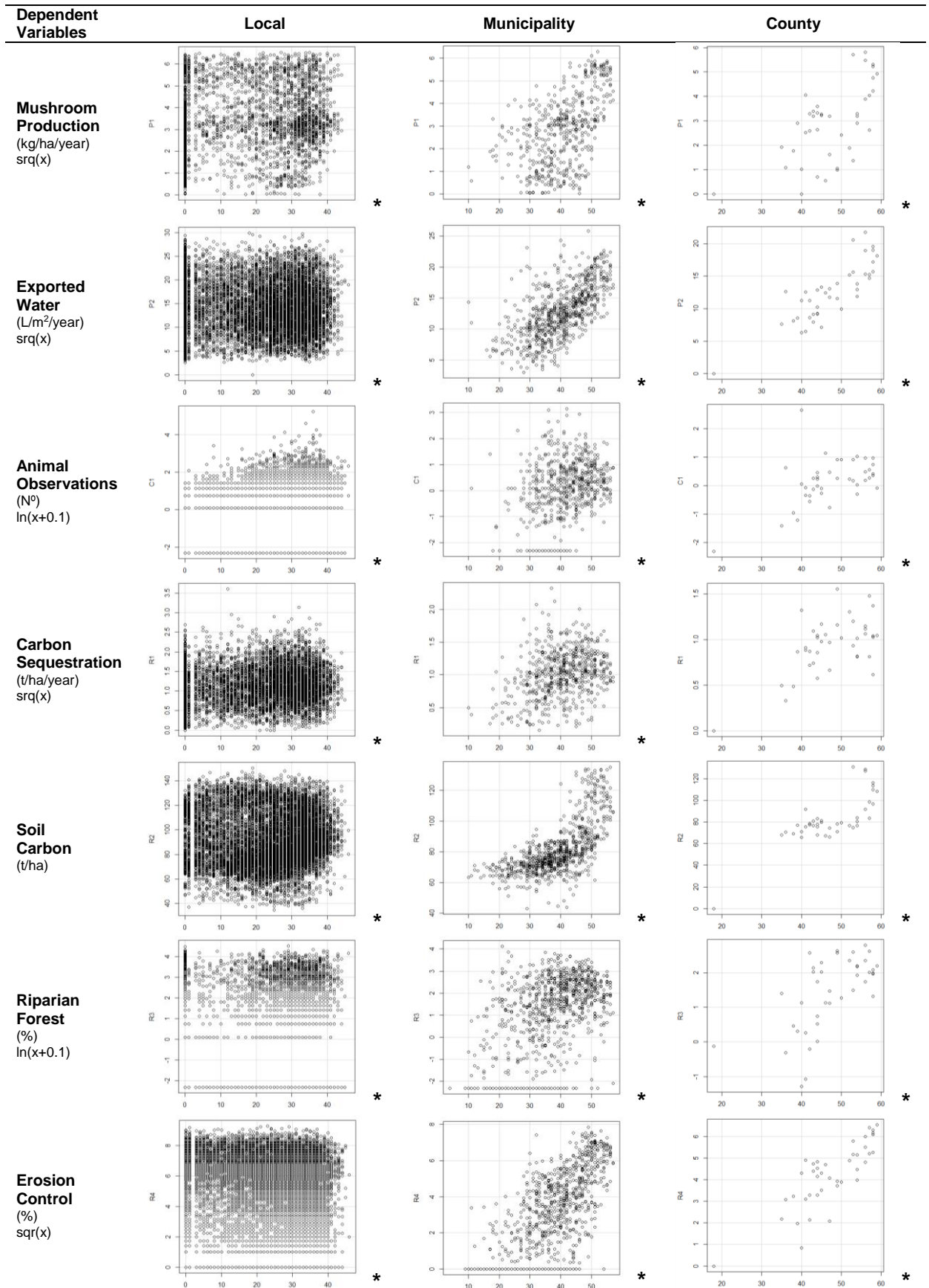
829
830
831

Appendix C. Scatter-plots among forest descriptors (biodiversity and forest cover variables) and ES indicators at the three spatial levels of analysis.

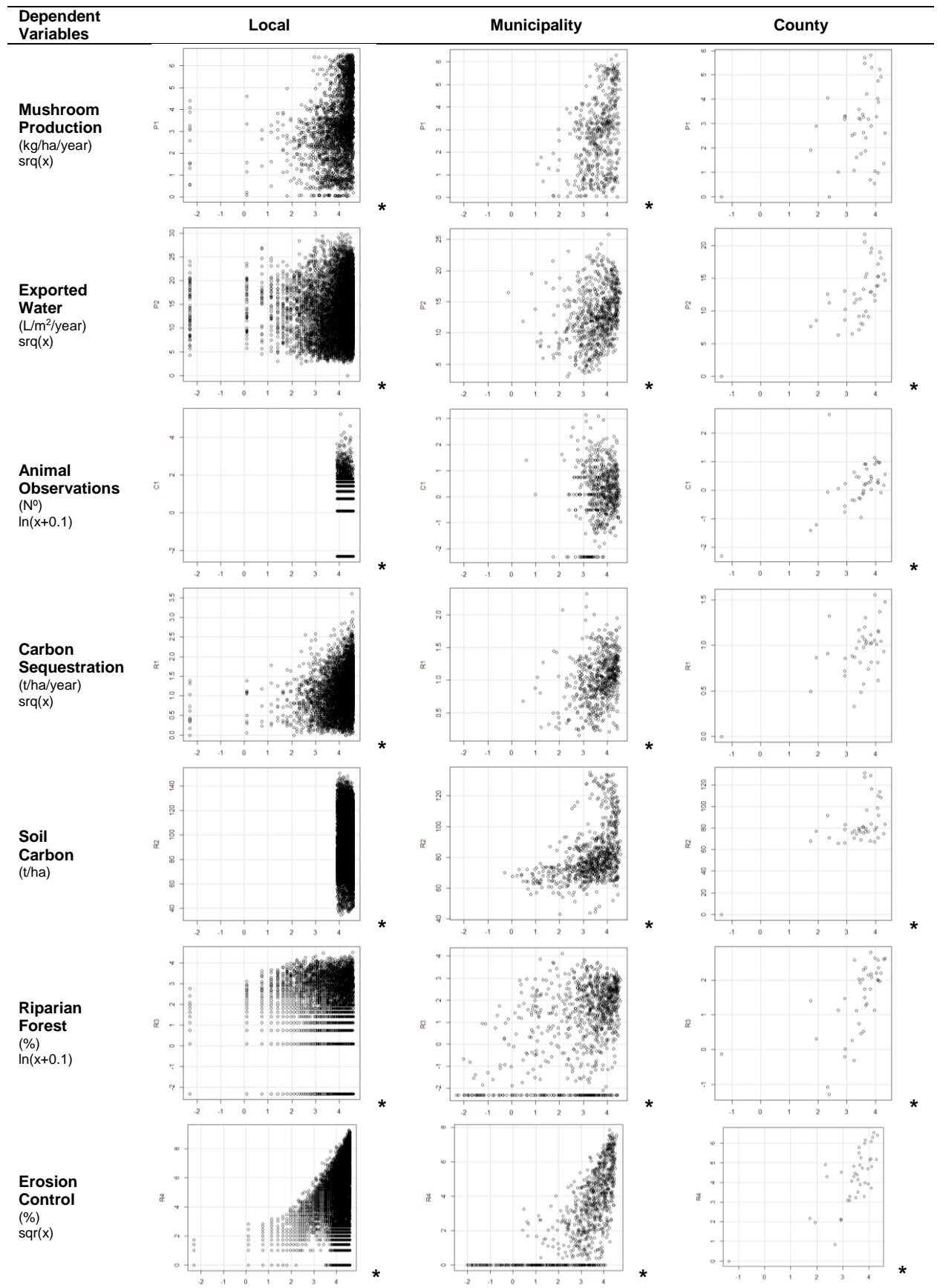


832
833

Figure S4. Scatter-plots between woody species richness (square root transformed, X-axis) and ES indicators (*: p-value < 0.05). Units of dependent variables and their transformations (if any) are showed in the first column.



834 Figure S5. Scatter-plots between bird species richness (X-axis) and ES indicators (*: p-value < 0.05). Units of
835 dependent variables and their transformations (if any) are showed in the first column.



836
837
838

Figure S6. Scatter-plots between forest cover (X-axis; transformation: ln (forest cover + 0.1) and ES indicators (*: p-value < 0.05). Units of dependent variables and their transformations (if any) are showed in the first column.