Rank Aggregation from Pairwise Comparisons in the Presence of Adversarial Corruptions

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Abstract

Rank aggregation from pairwise preferences has widespread applications in recommendation systems and information retrieval. Given the enormous economic and societal impact of these applications, and the consequent incentives for malicious players to manipulate ranking outcomes in their favor, making rank aggregation algorithms robust to adversarial manipulations in data is a crucial challenge. In this paper, we initiate the study of robustness in rank aggregation under the popular Bradley-Terry-Luce (BTL) model for pairwise comparisons. We consider a setting where pairwise comparisons are initially generated according to a BTL model, but a fraction of these comparisons are corrupted adversarially prior to being reported to us. We consider a strong contamination model, where an adversary having complete knowledge of the initial truthful data and the true BTL weights, can corrupt this data by inserting, deleting, or changing data points. The goal is to recover the true BTL weights given this corrupted data. We characterize the extent of corruption under which the true BTL weights are uniquely identifiable. We also provide a novel algorithm that provably filters out the adversarial corruption from data under reasonable conditions on data generation and corruption. We support our theory with experiments on both synthetic as well as real data, showing the resilience of our algorithm to a substantial degree of corruption and the vulnerability of existing approaches to even small amounts of corruption.

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1. Introduction

The problem of rank aggregation from pairwise comparisons, where the goal is to aggregate pairwise preferences between items into rankings/scores for each item, has a wide range of applications in the areas of recommendation systems and information retrieval (Dwork et al., 2001; Negahban et al., 2017; Maystre & Grossglauser, 2015; Agarwal et al., 2018; Hendrickx et al., 2019; Wauthier et al., 2013; Ailon et al., 2008; Gleich & Lim, 2011; Guiver & Snelson, 2009; Volkovs & Zemel, 2012). In these large scale web-applications for recommendation and retrieval, one obtains pairwise preferences from different users either explicitly through survey questions or implicitly through clicks, ratings, reviews etc. and aggregates these preferences to score/rank items/products for these users.

The massive economic and societal impact of these applications has also meant that some players are trying to boost the ranking/scores of their products by resorting to malicious practices such as creating fake user accounts, manufacturing fake reviews and ratings, click-fraud etc. Hence, it has become increasingly important to guard against these malicious players by designing ranking algorithms that are robust to adversarial corruption in data.

In order to address this challenge, we initiate the study of *robustness* in rank aggregation under the Bradley-Terry-Luce (BTL) model (Bradley & Terry, 1952; Luce, 1959), which is arguably the most popular parametric model for rank aggregation using pairwise comparisons. We describe the exact setting below.

1.1. Problem Formulation

Given a set of n items, the BTL model associates a positive weight/score w_i^* with each item $i \in [n]$, and postulates that item i wins in a pairwise comparison against item j with probability $p_{ij}^* = w_i^*/(w_i^* + w_j^*)$. Since this model is invariant under multiplicative scaling, for uniqueness, it is assumed that $\mathbf{w}^* \in \Delta_n$, the open n-simplex, where $\mathbf{w}^*$ is the vector of the aforementioned BTL weights. In our framework, nature first draws a comparison graph $G^* = (V, E^*)$ which is an undirected graph with vertex set $V = [n]$ and edge set $E^* = \{(\{u_i, v_i\}, \hat{p}_{u_i v_i})\}_{i=1}^{m^*}$

consisting of m^* edges, where the label $\hat{p}_{uv}$[1] on edge $(u, v) \in E^*$ corresponds to the fraction of times i beats j out of L pairwise comparisons drawn according to the underlying BTL model with (unknown) weights $\mathbf{w}^*$, for a parameter $L \in \mathbb{N}$. We consider a *powerful contamination model* where an adversary having complete knowledge of the truthful graph $G^* = (V, E^*)$, as well as true weights $\mathbf{w}^*$, can subsequently contaminate some fraction of E^* by adding spurious new edges with arbitrary labels, deleting and corrupting existing edges/labels. As a result, we receive as input a comparison graph $G = (V, E)$ with edge set $E = \{(\{u_i, v_i\}, p_{u_i v_i})\}_{i=1}^m$ consisting of a subset $E_u = E^* \cap E$ of uncorrupted edges from the initial truthful data, where for each $(\{u, v\}, p_{uv}) \in E_u$, the reported probability value p_{uv} is equal to the uncorrupted probability $\hat{p}_{uv}$. The remaining subset $E_a = E \setminus E_u$ consists of either newly introduced edges, or edges already existing in E^* whose labels were corrupted by the adversary. In either case, no assumptions can be made on the reported probability values p_{uv} for edges $(\{u, v\}, p_{uv}) \in E_a$. The set $E^* \setminus E$ is the set of edges deleted by the adversary.

In this adversarial contamination model, our work addresses the following fundamental questions:

- *For an arbitrary truthful comparison graph $G^* = (V, E^*)$, what is the extent of adversarial corruption that can be tolerated up to which the true BTL parameters are uniquely identifiable?*
- *Are there structural properties of $G^* = (V, E^*)$ that allow tolerance to high degrees of adversarial corruption?*
- *Do there exist efficient algorithms to estimate the true BTL parameters (with low error) given pairwise comparison data with a non-trivial fraction of adversarial corruption?*

Notation. Given any subset of edges E' and cut $(S, V \setminus S)$, we use $E'(S, V \setminus S)$ to refer to the set of edges in E' that cross the cut $(S, V \setminus S)$. In the event that S is a singleton vertex $u \in V$, we use $E'(u) := E'(\{u\}, V \setminus \{u\})$ to refer to the set of edges in E' incident on u. Given any subset of edges E' and a vertex $u \in V$, we use $\delta_{E'}(u)$ to refer to the set of neighbors of u in the graph $G' = (V, E')$.

1.2. Overview of Results

We naturally consider *structural identifiability* of the true BTL weights within our contamination model, i.e. unique identifiability of $\mathbf{w}^*$ when the uncorrupted labels $\hat{p}_{uv}$ for all $(u, v) \in E_u$ are exactly equal to the true pairwise probabilities p^*_{uv}, a setting corresponding to the limit $L \to \infty$.

We first present a candidate hard example of an adversarial corruption, which not just demonstrates the kind of carefully crafted corruptions that make this setting challenging, but also helps form a basis for our identifiability results later.

[1]Since the probability $\hat{p}_{uv}$ can be inferred from the probability $\hat{p}_{vu}$, we will assume that there is fixed ordering over items, if $u < v$ then the label corresponds to $\hat{p}_{uv}$ corresponding to pair $\{u, v\}$ otherwise it corresponds to $\hat{p}_{vu}$.

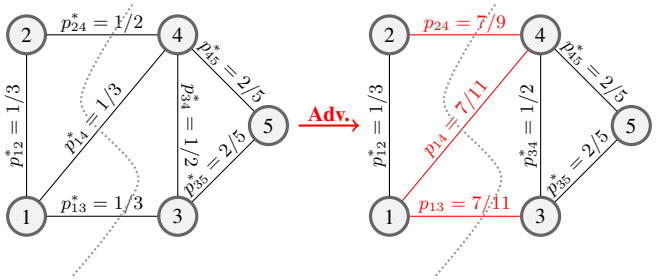

Figure 1. An instance of Example 1, with $S = \{1, 2\}$, $\alpha = 3/5$, and $\mathbf{w}^* = (7, 14, 14, 14, 21)/70$; By corrupting just the edges crossing the cut (dotted line), the resulting graph is entirely consistent with $\mathbf{w}^{(\alpha,S)} = (14, 28, 8, 8, 12)/70$. Note how the items with some of the lowest scores have the highest scores post corruption.

Example 1 (Single Cut Corruption)**.** *Given the truthful comparison graph $G^* = (V, E^*)$, and true weights $\mathbf{w}^*$ with $\hat{p}_{uv} = p^*_{uv}, \forall (u, v) \in E^*$, fix an arbitrary cut $(S, V \setminus S)$. Let $w^*_S := \sum_{u \in S} w^*_u$ be the total weight of all vertices in S. We create new weights $\mathbf{w}^{(\alpha,S)}$, where for every vertex $u \in S$, we scale up its weight as $w^{(\alpha,S)}_u = \alpha w^*_u / w^*_S$, and for every vertex $v \in V \setminus S$, we scale down its weight as $w^{(\alpha,S)}_v = (1 - \alpha) w^*_v / (1 - w^*_S)$, where $w^*_S < \alpha < 1$ is any arbitrary scaling factor. Note that the relative weights within S and $V \setminus S$ are unaffected by this change. If the adversary corrupts only the edges $E^*(S, V \setminus S)$ crossing the cut $(S, V \setminus S)$ to be consistent with the new weights $\mathbf{w}^{(\alpha,S)}$, leaving all other edges untouched, then the resulting graph is entirely consistent with $\mathbf{w}^{(\alpha,S)}$. (See Figure 1)*

This example shows a coordinated corruption where the adversary only needs to corrupt the edges in a single cut in the graph to make the entire comparison graph consistent with completely different BTL weights, leaving behind no evidence of corruption. This also has an intuitive interpretation: The set S consists of items of interest to the adversary, and $V \setminus S$ consists of the rest of the items. By only corrupting the comparison data between the items of interest and the rest of the items, the adversary manipulates the relative ranking between items of interest and the rest of the items, leaving the internal ranking within both these sets unchanged.

This example provides the key intuition behind the condition which we prove is both necessary and sufficient for unique identifiability of the true weights $\mathbf{w}^*$.

Theorem 1 (Informal)**.** *Given an arbitrary, connected, corrupted input comparison graph $G = (V, E)$, the true weights are uniquely identifiable in the limit $L \to \infty$ if and only if every cut in G has strictly more uncorrupted edges than corrupted edges crossing the cut.*

The above theorem is essentially a majority condition for

unique identifiability. In some sense, this demonstrates that Example 1 is a canonical type of corruption that must be guarded against. This also shows that in comparison graphs with sparse cuts, even small amounts of carefully crafted corruptions can make the true weights unidentifiable.

This motivates the study of Erdős-Rényi comparison graphs, a well-studied family of graphs in ranking, as they are known to have dense cuts. In the following Theorem, we show that if the initial truthful comparison graph is drawn according to the Erdős-Rényi model, then the global cut-based condition for identifiability reduces to a local bound on the fraction of corrupted edges incident on any vertex.

Theorem 2 (Informal)**.** *When the initial truthful comparison graph G^* is an Erdős-Rényi graph, with high probability, the true weights are uniquely identifiable in the limit $L \to \infty$ if the fraction of corrupted edges per vertex is at most $\frac{1}{4} - \epsilon$, and conversely are not uniquely identifiable if the fraction of corrupted edges per vertex exceeds $\frac{1}{4} + \epsilon$, where ϵ is any arbitrarily small positive constant.*

The above theorem shows that a corruption rate of $1/4$-th per vertex is a sharp threshold for unique identifiability. The proof follows by exploiting the structural regularities imposed by the Erdős-Rényi model, which imply that a corruption rate of at most $1/4 - \epsilon$ per vertex is sufficient to guarantee the majority condition described in Theorem 1 for every cut in the graph, and contrarily, if the corruption rate per vertex exceeds $1/4 + \epsilon$, then there exists a cut which violates the majority condition. Due to the randomness in the graph model, these claims hold with high probability.

Although these theorems characterize conditions for unique identifiability, they do not imply an efficient algorithmic procedure for recovering the true weights from a corrupted comparison graph. Our final contribution is an efficient algorithm with provable recovery guarantees when the initial truthful comparison graph is an Erdős-Rényi graph.

Theorem 3 (Informal)**.** *When the initial truthful comparison graph is an Erdős-Rényi graph and the fraction of corrupted edges per vertex is at most $O(\frac{\log d}{\log n})$ where d is the average degree in the graph, there exists an algorithm that recovers the true weights exactly in the limit $L \to \infty$, and approximately (with low error) in case of finite L.*

Our efficient recovery algorithm can provably tolerate an inverse logarithmic corruption rate $O(\log \log n / \log n)$ in sparse graphs with $O(n \log n)$ edges, and a constant corruption rate in slightly denser graphs with $O(n^{1+\epsilon})$ edges for any constant $0 < \epsilon \leq 1$. Under this corruption rate of $O(\log d / \log n)$, our recovery error in terms of L in this adversarial setting matches the best known error rate for Erdős-Rényi in the non-robust setting (Agarwal et al., 2018).

At the heart of this result lies a filtering algorithm that removes *every edge with significant deviation* from the true pairwise probability. This algorithm is based on the key idea that the ratios of pairwise probabilities p_{uv}/p_{vu} correspond to ratios of weights w_u/w_v, and if the product of these ratios over some cycle significantly deviates from 1, then there must have been at least one significantly corrupted edge on that cycle. Hence, this inconsistency in a cycle can be used as a *certificate of corruption* for corrupted edges in the cycle. The algorithm solves a linear program (LP) with a *hitting set* constraint for all inconsistent cycles and rounds the solution identify the significantly corrupted edges in these cycles.

A key structural property of Erdős-Rényi graph that makes this approach feasible is the existence of *short certificates* of corruption for every significantly corrupted edge in the input graph. The main challenge here is proving that every significantly corrupted edge would be pruned, and simultaneously, sufficiently many uncorrupted edges would survive to allow weight recovery after rounding the fractional solution, which is non-trivial to prove. To this end, we prove an *adversarially robust* structural property of Erdős-Rényi graphs, that guarantees that if some significantly corrupted edge survived the filtering, then some short certificate of corruption for that edge must have also survived the filtering, which would imply a violated constraint. Due to this coupling between corruptions and corresponding certificates of corruption, the corruption rate that can be provably handled by the linear program is inherently tied to the lengths of these certificates. Due to this, the corruption rates that our algorithm can provably recover from increases as the density of the underlying comparison graph increases, as denser graphs admit shorter certificates.

1.3. Related Work

The general problem of rank aggregation using pairwise comparisons under the BTL model has been well-studied, and there are several consistent algorithms for recovering the BTL parameters (Hunter, 2004; Negahban et al., 2017; Hendrickx et al., 2019). Moreover, there are also consistent algorithms for rank aggregation using multiway comparisons under the MNL model (Maystre & Grossglauser, 2015; Agarwal et al., 2018), which is a generalization of the BTL model. However, these algorithms were not designed with robustness in mind, and as a consequence, have recovery guarantees only when the comparison data is drawn stochastically from the underlying model; unbiased noise due to sampling is benign compared to the arbitrary adversarial corruption we allow.

Another related line of work is parameter recovery under a mixture of BTL models using pairwise comparisons, where the goal is to recover parameters of all the components along with the mixture weights (Oh & Shah, 2014; Chierichetti et al., 2018; Suh et al., 2017; Zhao & Xia, 2019). However, these mixture models crucially differ from our adversarial contamination model as the pairwise probability on any edge

in these models is a convex combination of the pairwise probabilities defined by the individual BTL components, whereas in our model the pairwise probability on an edge is either consistent with the underlying true BTL model, or is arbitrary. Hence, the identifiability and recoverability results for these models do not apply to our setting. There has been some effort in addressing adversarial mixtures (Suh et al., 2017), but their model is in fact a mixture of 2 specific BTL models: the true BTL model and its inverse. As a consequence, their mixture model is incomparable to our adversarial contamination model.

An adversarial corruption model similar to ours has been studied in the computer vision literature (Goldstein et al., 2016; Hand et al., 2018) for a problem of recovering locations of objects given direction (unit) vectors between pairs of locations. Although our problem is very different than theirs, it is worth noting that their recovery results assume an extremely dense Erdős-Rényi comparison graph over locations, whereas our recovery results hold for even very sparse Erdős-Rényi comparison graphs. Moreover, our corruption model is somewhat stronger than theirs as the adversary in their model can only corrupt existing data points, while the adversary in our model can even add or delete data points.

Concurrent to our work, (Lerman & Shi, 2019) studied the problem of robust group synchronization, which asks to recover ground truth group elements given measurements consisting of noisy pairwise group ratios, some of which may be adversarially corrupted. The results of (Lerman & Shi, 2019) specifically aim to estimate the corruption level on every pairwise measurement, which is a variant of the recovery objective. Although we present our work in the context of ranking from pairwise comparisons, it is straightforward to see that our results are applicable to this more general problem as well. Our algorithmic ideas, and consequently, our recovery guarantees differ significantly. Their proposed algorithm is applicable to a specific family of comparison graphs and has provable guarantees only under a weak sufficiency condition on the amount of corruption in the input graph: for every edge in the graph, there is at least one cycle of length at most k (a fixed constant) involving that edge where the rest of the cycle edges are uncorrupted and the number of such good cycles is at least a 3/4 fraction of the total number of cycles of length at most k involving that edge. On the other hand, our algorithm is designed for the family of Erdős-Rényi comparison graphs, and has much stronger guarantees in this regime: it can tolerate a $O(\log\log n/\log n)$ fraction of corruption per vertex in the sparse regime and can tolerate a constant fraction of corruption per vertex in the dense regime. The former regime cannot be handled at all by their approach, and in the latter regime, there is no clear bound on the fraction of corruption per vertex required to satisfy the weak sufficiency condition under which their approach has provable guarantees.

In addition to our algorithmic results, we also provide an exact necessary and sufficient condition under which recovery is fundamentally possible for arbitrary contaminated comparison graphs, which is not addressed in their work.

Our proposed framework is very closely related to robust estimation theory in classical statistics, in particular, the $\epsilon-$contamination model of Huber (Huber, 1965; 1992) and its generalizations (Diakonikolas et al., 2017). A canonical problem in this literature is robust estimation of parameters of a Gaussian distribution under a corruption model where an ϵ fraction of the truthful Gaussian samples are arbitrarily corrupted by an omniscient adversary. Until recently, all known algorithms for this problem had an inherent tradeoff between computational tractability and the quality of the recovered estimates, and it was a long standing open problem of whether it was possible to have a computationally efficient estimator that also had information theoretically optimal error guarantees. This was resolved in Diakonikolas et al. (2017). Also, see Chen et al. (2016); Diakonikolas et al. (2019; 2018) for other interesting results.

2. A Cut-Based Characterization for Identifiability in General Graphs

In this section, we study unique identifiability in the limit that the number of samples per pair L goes to infinity, i.e. the setting where uncorrupted edge labels are exactly equal to the true pairwise probabilities under BTL. We show that the true weights are uniquely identifiable if and only if the comparison graph induced by the input data satisfies a cut-based majority condition.

Theorem 1. *Given any arbitrary comparison graph $G = (V, E)$ as input, it is possible to uniquely identify the true weights $\mathbf{w}^*$ in the limit $L \to \infty$, if and only if for every cut $(S, V \setminus S)$*

$$|E_u(S, V \setminus S)| > |E_a(S, V \setminus S)|\,,$$

where $E_u \subseteq E$ is the set of uncorrupted edges, and $E_a = E \setminus E_u$ is the set of adversarially corrupted edges.

The above theorem exactly characterizes the extent of adversarial corruption that one can recover from in *any* corrupted comparison graph G based on a *cut-majority* condition. The above theorem also provides a verification algorithm which, given a comparison graph G and a candidate solution $\mathbf{w}$, can identify whether $\mathbf{w}$ is the true weight vector. Before discussing this verification algorithm, we will first give a basic notion of an edge being consistent with a solution $\mathbf{w}$.

Definition 1 (Consistent-Edge)**.** *Given an input comparison graph $G = (V, E)$, we say that an edge $(u, v) \in E$ is consistent with a solution $\mathbf{w}$, and vice-versa, if and only if $p_{uv} = w_u/(w_u + w_v)$.*

Note that any uncorrupted edge is always consistent with the true weights $\mathbf{w}^*$. Given that the cut-majority condition is

satisfied for G, the following simple corollary to Theorem 1 gives a way to verify whether a solution $\mathbf{w}$ is correct or not.

Corollary 1. *If a input comparison graph $G = (V, E)$ satisfies the recoverability condition in Theorem 1 then in the limit $L \to \infty$, $\mathbf{w}^*$ is the unique solution that, for every cut $(S, V \setminus S)$, is consistent with a strict majority of the edges crossing the cut.*

Although, this characterization works for all graphs, it might be computationally infeasible to check all possible cuts in order to verify if a solution is correct.

A key implication of this theorem is that the structure of the comparison graph induced by pairwise comparison data plays a crucial role in determining tolerance to corruption. While it is clear to see that true weights are unidentifiable if a majority of the edges incident on any vertex get corrupted, *even restricting the fraction of corrupted edges incident on any vertex is not enough to guarantee unique identifiability*.

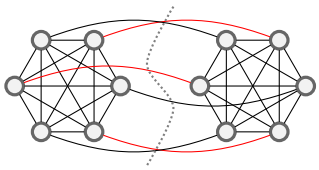

Figure 2. Sparse cuts across dense subgraphs can easily be exploited, even by a limited budget adversary.

Figure 2 demonstrates why: An adversary merely needs to corrupt a majority of the edges crossing a *sparse cut* to obfuscate the true weights. In an extreme case, where the true comparison graph is a regular graph consisting of dense subgraphs with $\Omega(n^2)$ edges separated by a sparse cut with $O(n)$ edges, even restricting the fraction of corrupted edges incident on any vertex to as low as $O(1/n)$ cannot guarantee identifiability, which is a trivial bound as $\Omega(1/n)$ is needed to allow even one corrupted edge in the comparison graph.

3. Results for Erdős-Rényi Comparison Graphs

From the discussion in the previous section, we can conclude that for a comparison graph to be resilient to corruption, the number of edges crossing any cut should be comparable to the number of edges on one side of the cut (the smaller side), also known as *edge expansion*. A natural candidate for graphs having this property are Erdős-Rényi graphs, which are random graphs that have constant edge expansion with high probability. Given a parameter $p \in [0, 1]$, an Erdős-Rényi graph $G_{n,p}$ is a random graph over n vertices where each edge (u, v) is sampled independently with probability p. These graphs have been widely studied in various domains, including ranking from pairwise comparisons (Chen & Suh, 2015; Jang et al., 2016; Chen et al., 2017).

These graphs have another very interesting property: the global cut-majority condition for unique identifiability of the true weights effectively reduces to a much simpler local vertex-majority condition. This is attractive for several reasons, the foremost being that verifying this condition is extremely efficient, making it usable in practice. Before elaborating on these observations, we will first formalize the contamination model for Erdős-Rényi graphs.

3.1. Adversarial Contamination Model

Given a parameter $p \geq (k \log n)/n$ for any k larger than some sufficiently large constant, the comparison graph $G^* = (V, E^*)$ generated by nature is a random $G_{n,p}$ graph.

Given a *corruption rate* parameter $\gamma > 0$, the adversary can introduce arbitrary contaminations into the realized comparison graph G^*, resulting in a corrupted comparison graph $G = (V, E)$, albeit subject to the constraint

$$|E_r(u) \cup E_a(u)| \leq \gamma |E^*(u)|, \qquad \forall u \in V \tag{1}$$

where for any vertex $u \in V$, $E^*(u)$ is the initial set of uncorrupted edges incident on u in G^*, $E_a(u) := \{(u, v) \in E : p_{uv} \neq \widehat{p}_{uv}\}$ is the set of corrupted edges incident on u in G, and $E_r(u) := E^*(u) \setminus E(u)$ is the set of edges that were incident on u in G^* but were later deleted in G.

This condition effectively limits the adversary to contaminating at most a γ fraction of the incident edges on any vertex in the graph. Observe that this condition further implies that at most a γ fraction of the edges incident on any vertex in the corrupted graph can have spurious labels, i.e. $|E_a(u)| \leq \gamma |E(u)|$, which we will crucially use later.

3.2. A Sharp Threshold Condition for Identifiability

Given the contamination model from above, we show that there is a sharp threshold on the per-vertex corruption rate for unique identifiability of the true weights; if the corruption rate γ is smaller than this threshold, then $\mathbf{w}^*$ is uniquely identifiable with high probability for any choice of adversarial corruption. Contrarily, if the corruption rate γ is larger than this threshold, then with high probability, there exists a choice of corruption such that the $\mathbf{w}^*$ is unidentifiable. The proof of this claim crucially exploits the following strong edge expansion property of Erdős-Rényi graphs.

Fact 1. *Given any arbitrarily small constant $\epsilon > 0$, there exists a sufficiently large constant k, such that given a graph $G = (V, E) \sim G_{n,p}$ with parameter $p \geq (k \log n)/n$, we have for every cut $(S, V \setminus S)$*

$$(1 - \epsilon)\,|S||V \setminus S|p < |E(S, V \setminus S)| < (1 + \epsilon)\,|S||V \setminus S|p$$

This claim holds with probability at least $1 - 1/\text{poly}(n)$.

This fact roughly guarantees that with high probability, the number of edges crossing any cut in an Erdős-Rényi graph will not deviate from its expected value by a large amount.

Theorem 2. *Given any arbitrarily small constant $\epsilon > 0$, there exists a sufficiently large constant k, such that given a input comparison graph $G = (V, E)$ conforming to the contamination model in Section 3.1 with Erdős-Rényi graph parameter $p \geq (k \log n)/n$, if the corruption rate $\gamma \leq \frac{1}{4} - \epsilon$, then with probability at least $1 - 1/\text{poly}(n)$, the cut-majority condition described in Theorem 1 is satisfied for every cut in G, and as a consequence, the true weights $\mathbf{w}^*$ are uniquely identifiable as the number of samples per pair $L \to \infty$. Conversely, if the corruption rate $\gamma \geq \frac{1}{4} + \epsilon$, then with probability at least $1 - 1/\text{poly}(n)$, there exists a choice of adversarial corruption such that the cut-majority condition described in Theorem 1 is violated for at least one cut in G, rendering the true weights unidentifiable, even as $L \to \infty$*

Given that the vertex-majority condition holds, the following simple corollary to the above lemma shows that there is a linear time algorithm to verify whether a candidate solution $\mathbf{w}$ is in fact correct solution $\mathbf{w}^*$.

Corollary 2. *In the setting of Theorem 2, if $\gamma \leq 1/4 - \epsilon$, then $\mathbf{w}^*$ is the unique solution such that for every vertex $v \in V$, at least $3/4 + \epsilon$ fraction of its incident edges in G are consistent with $\mathbf{w}^*$, where consistency of an edge with a solution is defined in Definition 1.*

Hence, a candidate solution $\mathbf{w}$ is the true solution $\mathbf{w}^*$ if and only if for every vertex $v \in V$, at least $3/4 + \epsilon$ fraction of its incident edges are consistent with $\mathbf{w}$. Note that unlike the cut-majority condition where a simple majority is enough, here we necessarily need a majority of a little over $3/4$, as even incorrect weights can achieve close to $3/4$ majority. Note that checking this condition just requires knowledge of a lower bound on ϵ, and not the exact value of γ.

Although this vertex-majority condition makes verification of a candidate set of weights easy, it does not directly imply a polynomial time algorithm for recovering the true weights. In the next section we will design such a recovery algorithm.

3.3. An Algorithm for Weight Recovery

The following theorem gives the main result of this section.

Theorem 3. *Given an input comparison graph $G = (V, E)$ conforming to the contamination model in Section 3.1 with Erdős-Rényi graph parameter $p \geq (k \log n)/n$ for any k larger than some sufficiently large constant, true BTL weights $\mathbf{w}^*$, and number of samples per pair L; if the corruption rate per vertex $\gamma \leq \log(np)/(125 \log n)$, then there is an efficient algorithm that, with probability at least $1 - 1/\text{poly}(n)$, recovers an estimate $\mathbf{w} \in \Delta_n$ such that*

$$||\mathbf{w}^* - \mathbf{w}||_1 \leq cb^2 \log b \sqrt{\log n / L}\,,$$

for an absolute constant c, where b is an upper bound on the skew in item quality $\max_{i,j \in [n]} w_i^*/w_j^*$.

The corruption rate that can be tolerated by our recovery algorithm varies depending on the density of the underlying comparison graph. When the initial graph is very sparse, i.e. when the average degree is $O(\log n)$, then our algorithm can tolerate a corruption rate of approximately $(\log \log n)/\log n$, which is lower than the theoretical limit of identifiability described in Theorem 2. However, for slightly denser graphs, i.e. when the average degree is $O(n^\epsilon)$ for any constant $0 < \epsilon \leq 1$, then our algorithm can handle a constant corruption rate.

To contrast the above guarantee with the results in the usual non-adversarial BTL setting, in (Negahban et al., 2017; Agarwal et al., 2018) the recovery error is $O(\sqrt{\log n/L})$, and hence, $L = \omega(\log(n))$ is enough to ensure consistency. It is surprising to see that our result matches this bound exactly (up to constants), implying there is *no additional statistical cost for achieving consistency* even under completely adversarial corruptions in the input pairwise comparison data when the corruption rate is $O(\log(np)/\log n)$.

In the case where we receive exact pairwise probabilities for every uncorrupted edge in the input, i.e. $L \to \infty$, we have

Corollary 3. *Let $G = (V, E)$ be any input comparison graph conforming to the contamination model in Section 3.1 with Erdős-Rényi graph parameter $p \geq (k \log n)/n$ for any k larger than some sufficiently large constant, and true BTL weights $\mathbf{w}^*$, where for every uncorrupted edge $(u, v) \in E_u$, we have $p_{uv} = p^*_{uv}$; if the corruption rate per vertex $\gamma \leq \log(np)/(125 \log n)$, there exists an efficient algorithm that with probability at least $1 - 1/\text{poly}(n)$ recovers $\mathbf{w}^*$ exactly.*

3.3.1. Algorithm

Our algorithm is based on solving a linear programming relaxation, and rounding the solution to remove all edges that *deviate significantly* from the true probability values. In the process, we might remove some uncorrupted edges as well, but the graph would still remain connected with high probability which will be enough to obtain a consistent estimate $\widehat{\mathbf{w}}$ of the true weights. When we are given the true pairwise probabilities for all uncorrupted edges, then our algorithm in fact removes *all* corrupted edges from the input, and subsequently returns the true weights $\mathbf{w}^*$.

Before describing our algorithm, we will formalize the notions of *significant deviation* from the true probability value, and *approximate consistency* within a cycle.

Definition 2 (Significant Deviation)**.** *Given an input comparison graph $G = (V, E)$ with pairwise probabilities $\{p_{uv}\}_{(u,v) \in E}$, conforming to the contamination model in Section 3.1 with Erdős-Rényi graph parameter p, number of comparisons per edge L, and true BTL weights $\mathbf{w}^*$; we use $E_A \subset E$ to refer to the set of edges that deviate significantly from their true probability value, where*

$$E_A := \left\{ (u, v) \in E : |p_{uv} - p^*_{uv}| > 4 \left(4 + \frac{\log n}{\log(np)} \right) \epsilon_L \right\}$$

where $\epsilon_L = (1+b)\sqrt{\log n / L}$.

From Hoeffding's inequality, we have with probability at least $1-1/\text{poly}(n)$, that $|p_{uv} - p^*_{uv}| \le \epsilon_L/(1+b)$ for every uncorrupted edge $(u,v) \in E_u$, due to which no uncorrupted edge would be included in this set E_A. Thus, $E_A \subseteq E_a$ with $E_A = E_a$ when $\epsilon_L = 0$, i.e. $p_{uv} = p^*_{uv}$ for all $(u,v) \in E_u$.

Definition 3 (Approximate Consistency)**.** *Given an input comparison graph $G = (V, E)$ with pairwise probabilities $\{p_{uv}\}_{(u,v)\in E}$, conforming to the contamination model in Section 3.1, with number of comparisons per edge L; given a simple cycle $C = (v_1, \cdots, v_l, v_1)$ of length l in G, we call C approximately consistent if*

$$\frac{1-(2l-1)\epsilon_L}{1+\epsilon_L} \le \prod_{i=1}^{l} \frac{p_{v_{c_i} v_{c_{i+1}}}}{p_{v_{c_{i+1}} v_{c_i}}} \le \frac{1+\epsilon_L}{1-(2l-1)\epsilon_L},$$

and inconsistent otherwise.

The underlying intuition becomes clear when $\epsilon_L = 0$, i.e. we receive exact probabilities for every uncorrupted edge in the input. For any pair of vertices (u,v), we have that $p_{uv}/p_{vu} = w_u/w_v$ if the pairwise probabilities were defined according to the BTL model with weights $\mathbf{w}$. In this case, a simple cycle being consistent intuitively means that there exists *some* set of BTL weights consistent with the pairwise probabilities on *all* the edges in the simple cycle. While, a consistent cycle does not guarantee every edge in the cycle is uncorrupted as the adversary can introduce self-consistent corruptions, every inconsistent cycle must necessarily contain at least one corrupted edge. The constraints in our LP essentially capture this condition. For finite L, we need to allow some slack due to noise from sampling, due to which we have the slightly weaker guarantee that every inconsistent cycle must necessarily contain some edge that *deviates significantly* from its true probability value (i.e. from E_A).

Linear Program. We formulate a LP (Figure 3) with decision variables $x(e)$ for each edge $e \in E$, which indicate whether an edge is corrupted; a higher mass on $x(e)$ intuitively corresponds to a higher confidence of the LP solution in e being a corrupted edge. The LP has two types of constraints: Firstly, for each inconsistent cycle of length at most $4 + \log n/\log(np)$, we have a constraint requiring the total mass of all edges in the cycle be at least 1, reflecting the fact that each inconsistent cycle contains at least one corrupted edge. Secondly, for each vertex $u \in V$, we have a constraint requiring the total mass of all edges incident on u be at most[2] a $\gamma_{\text{LP}} = \log(np)/125\log n$ fraction of the degree of u, reflecting the fact that the number of corrupted edges incident on any vertex are bounded.

[2]The LP is oblivious to the exact corruption rate γ, and will work for any $\gamma \le \gamma_{\text{LP}}$, which is an upper bound on the corruption rate that the LP can *provably recover from.*

$$\begin{aligned}
\text{Minimize} \quad & \sum_{e\in E} x(e) & \\
\text{Subject to} \quad & \sum_{e\in C} x(e) \ge 1 & \forall C \in \mathbb{C} \\
& \sum_{e\in E(u)} x(e) \le \gamma_{\text{LP}} |E(u)| & \forall u \in V \\
& 0 \le x(e) \le 1 & \forall e \in E
\end{aligned}$$

Figure 3. The LP for identifying corrupted edges; $\mathbb{C}$ is the set of inconsistent cycles in G of length at most $4 + \log n/\log(np)$; $\gamma_{\text{LP}} = \log(np)/(125\log n)$ is the maximum tolerable corruption rate.

Lemma 1. *The LP in Fig 3 is solvable in $O(n^{2+o(1)}d^6)$ time where d is the average degree in the input graph.*

The proof leverages the Multiplicative Weight Update method (Plotkin et al., 1995) for approximately solving Linear Programs. We defer a detailed proof to the appendix.

Observation 1. *A solution that assigns $x(e) = 1$ to every edge $e \in E_a$, the set of adversarially corrupted edges, and $x(e) = 0$ to every edge $e \in E_u$, the set of uncorrupted edges is a feasible solution to the above LP.*

The proof follows by showing that no cycle consisting of only uncorrupted edges can be inconsistent. Thus, every inconsistent cycle must contain at least one corrupted edge, due to which every inconsistent cycle constraint is satisfied. Furthermore, the constraint for each vertex is satisfied due to the corruption condition in Equation 1. This shows that the feasible set of the above linear program is not empty.

Observation 2. *For any edge $(u,v) \in E_A$, any path from u to v consisting of edges only from E_u of length at most $4 + \log n/\log(np)$ will induce an inconsistent cycle.*

Threshold Pruning. Given any feasible solution $\mathbf{x}$ to the above LP, let $E_{lpr} := \{e \in E : x(e) \ge \log(np)/(5\log n)\}$ be the set of edges with large $x(e)$ values. We subsequently delete all edges from E_{lpr} from the input, producing a cleaned comparison graph $\widetilde{G} = (V, \widetilde{E} = E \setminus E_{lpr})$.

The key idea is to show that for every edge with significant corruption $(u,v) \in E_A$, there exists a *short* path from u to v consisting of only uncorrupted edges, which along with (u,v) would induce an inconsistent cycle (Obs 2), and hence, would be captured by our LP as a constraint. The harder challenge is in showing that *every* such edge in E_A would be removed by our threshold pruning scheme. The proof of this essentially involves showing that the residual comparison graph $\widetilde{G}$ still contains short paths consisting of only uncorrupted edges between every pair of vertices (which automatically implies connectedness), and thus, if some edge with significant corruption $(u,v) \in E_A$ survived, this would induce an inconsistent cycle. Furthermore, since $\widetilde{G}$ consists of only edges with small $x(e)$ values, this inconsistent cycle must have cumulative mass less than 1,

Algorithm 1 Adversarially Robust Recovery

1: **Input**: items $[n]$, graph $G = (V, E)$, parameters p and ϵ_L.
2: $\mathbf{x} \leftarrow$ Solution of LP in Figure 3.
3: $\forall (u, v) \in E, \hat{x}(u, v) \leftarrow \mathbf{1}[x(u, v) \geq \log(np)/(5 \log n)]$.
4: If $\hat{x}(u, v) = 1$ then delete data point corresponding to (u, v)
5: Return the output of Accelerated Spectral Ranking (Agarwal et al., 2018) algorithm on this pruned dataset.

implying a violated constraint, contradicting the assumption that we were given a feasible solution to the LP.

Lemma 2. *In the setting of Theorem 3, with probability at least* $1 - 1/\text{poly}(n)$*, we have that the residual graph* $\widetilde{G}$ *is connected, and furthermore, contains no edges from* E_A*.*

Since the residual graph $\widetilde{G} = (V, \widetilde{E})$ is free from edges that deviate significantly from their true probability value, the next step is to use an algorithm for recovery in the usual non-adversarial setting on $\widetilde{G}$. We use the Accelerated Spectral Ranking (ASR) algorithm (Agarwal et al., 2018), which defines a lazy random walk over $\widetilde{G}$ with probability of transition $\widetilde{P}_{uv}$ from vertex u to vertex v given by

$$\widetilde{P}_{uv} = \begin{cases} \frac{1}{\widetilde{d}_u} p_{vu} & \text{if } u \neq v,\ (u, v) \in \widetilde{E}, \\ \frac{1}{\widetilde{d}_u} \sum_{v \in \delta_{\widetilde{E}}(u)} p_{uv} & \text{if } u = v, \\ 0 & \text{otherwise.} \end{cases}$$

where $\widetilde{d}_u$ is the degree $|\delta_{\widetilde{E}}(u)|$ of vertex u in the graph $\widetilde{G} = (V, \widetilde{E})$. Let $\widetilde{\mathbf{P}} := [\widetilde{P}_{uv}]$ be the corresponding transition probability matrix, with transition probabilities as defined above. The solution $\widetilde{\mathbf{w}}$ returned by the ASR algorithm is a linear transformation $\widetilde{\mathbf{w}} = \widetilde{\mathbf{D}}^{-1}\boldsymbol{\pi}$, where $\boldsymbol{\pi} = \widetilde{\mathbf{P}}^\top \boldsymbol{\pi}$ is the stationary distribution of this Markov chain, and $\widetilde{\mathbf{D}}$ is the diagonal matrix of degrees $\widetilde{D}_{uu} = \widetilde{d}_u$.

The recovery guarantees for the ASR algorithm (and other existing algorithms) are only known in a setting where the estimates of pairwise probabilities are unbiased, which is not the case here as the residual graph may contain edges with biased probabilities. Nevertheless we show that the analysis of this algorithm can be extended to allow for biased pairwise probabilities satisfying a uniform deviation bound.

Lemma 3. *In the setting of Theorem 3, let* $\mathbf{w}^*$ *be the set of true BTL weights, and let* $\mathbf{w}$ *be the estimate returned by the ASR algorithm with input* $\widetilde{G} = (V, \widetilde{E})$*. Then we have that*

$$||\mathbf{w} - \mathbf{w}^*||_1 \leq (Cb \log b)\epsilon_L$$

where C *is an absolute constant.*

This, along with Lemma 1 gives us the claim of Theorem 3.

4. Experiments

In this section, we validate our theoretical guarantees with experiments on both synthetic and real data. In the interest of space, we show just one type of experiment here, where we compare the performance of our algorithm against existing non-robust algorithms when the input data has been contaminated according to the single cut corruption method as described in Example 1. We encourage the interested reader to refer to the Appendix for an additional type of experiment, where the contamination in the input data is semi-random in nature. The results obtained in that case are fairly similar to the ones reported in this section.

4.1. Synthetic Data

We fix $n = 50$, and generate a set of uniformly at random weights $\mathbf{w}^*$ normalized to sum to 1. We generate an Erdős-Rényi random comparison graph $G^* \sim G_{n,p}$ with parameter $p = (2 \log n)/n$. We choose a uniformly at random partition $(S, V \setminus S)$ of $n/2$ vertices each, and construct the adversarial vector $\mathbf{w}^{(\alpha,S)}$ as described in Example 1, for a fixed value of the scaling factor α set to 0.02. For every vertex $u \in S$, we pick a uniformly at random 2γ fraction[3] of its incident edges crossing the cut $(u, V \setminus S)$ to corrupt. We generate two datasets: (1) For every uncorrupted edge (u, v), we report the exact pairwise probability p^*_{uv} according to $\mathbf{w}^*$, and for every corrupted edge (u, v), we report the exact pairwise probability $p^{(\alpha,S)}_{uv}$ according to $\mathbf{w}^{(\alpha,S)}$, and (2) For every uncorrupted edge (u, v), we generate a random sample $X_{uv} \sim \text{Binomial}(L, p^*_{uv})$ and report $p_{uv} = X_{uv}/L$, $p_{vu} = 1 - p_{uv}$, and for every corrupted edge (u, v), we generate a random sample $Y_{uv} \sim \text{Binomial}(L, p^{(\alpha,S)}_{uv})$ and report $p_{uv} = Y_{uv}/L$, $p_{vu} = 1 - p_{uv}$. In our experiments, we set $L = \log n/\epsilon^2$, where $\epsilon = 5\%$ is the chosen accuracy parameter. We test all algorithms on both datasets.

4.2. Real Data

Experimentation with real datasets is challenging, primarily due to scarcity of datasets that are structurally robust to contamination. The datasets (`GIF`, `Youtube`) studied in (Agarwal et al., 2018; Maystre & Grossglauser, 2015) are found to be particularly vulnerable to manipulation; they contain cuts where corrupting just one edge is sufficient to completely fail the cut-majority condition (Thm 1) required for identifiability of the true weights. We circumvent this topology-dependent limitation by identifying datasets (`Sushi`, `Irish`) that come with full rankings. For these datasets, we extract pairwise comparisons from the complete orderings, giving us empirically observed pairwise probabilities p_{uv} for every pair of items (u, v) in the dataset (effectively inducing a complete comparison graph). Another difficulty with real data is that the true weights $\mathbf{w}^*$ are undefined. We resolve this issue by passing the datasets to a standard algorithm for parameter estimation in the BTL model (we choose the algorithm of (Agarwal et al., 2018)),

[3] since we corrupt only the cut edges, this is an effective corruption rate of γ

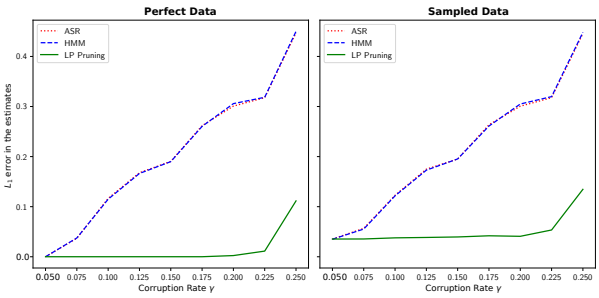

Figure 4. (**Synthetic data**) L_1 error in the recovered weights vs corruption rate γ

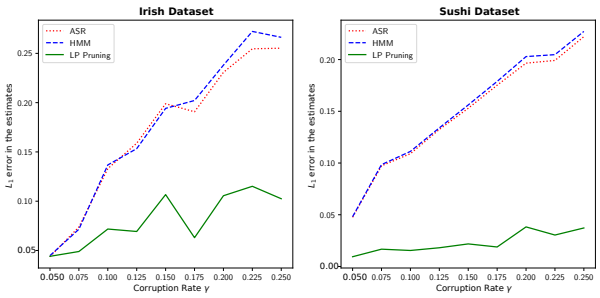

Figure 5. (**Real data**) L_1 error in the recovered weights vs corruption rate γ

and treating the returned estimates for each dataset as their corresponding ground truth weights.

For each real dataset, we create an artificially contaminated dataset as follows: given the complete comparison graph (`Sushi` $n = 16$, `Irish` $n = 12$) and the assumed ground truth weights $\mathbf{w}^*$, we first generate a Erdős-Rényi random comparison graph $G^* \sim G_{n,p}$ with parameter $p = 0.3$ by subsampling edges from the complete comparison graph. We choose a uniformly at random partition $(S, V \setminus S)$ of $n/2$ vertices each, and construct the adversarial vector $\mathbf{w}^{(\alpha,S)}$ as described in Example 1, for a fixed value of the scaling factor α set to 0.02. For every vertex $u \in S$, we pick $\gamma|E^*(u)|$ vertices in $V \setminus S$ uniformly at random, and insert corrupted edges between u and each of these vertices. For every uncorrupted edge (u, v), we report the empirically observed pairwise probability p_{uv}, and for every corrupted edge (u, v), we report the pairwise probability $p_{uv}^{(\alpha,S)}$ according to the adversarial vector $\mathbf{w}^{(\alpha,S)}$. We use this resulting contaminated comparison graph as input to all algorithms.

4.3. Algorithm Details

We implement our algorithm 1 in Python, and use the default LP solver in the `cvxpy` package to solve the LP described in Figure 3. We compare the performance of our algorithm against two standard algorithms for parameter estimation in the BTL model: Hunters minorization-maximization algorithm (Hunter, 2004) (abbr. HMM), and Accelerated Spectral Ranking (Agarwal et al., 2018) (abbr. ASR).

4.4. Experimental Results

In our experiments, we vary γ in the range 5%-25% in increments of 2.5%, and plot the average L_1 error in the returned weight vectors across 50 random trials. The results observed for synthetic data essentially verify our theoretical guarantees: the error in the estimates returned by our algorithm does not depend on the corruption rate up until the corruption rate becomes too large, after which a few corrupted edges pass through our filtering subroutine, whereas for existing algorithms, the error monotonically increases with increasing corruption rate. The results obtained for real data provide compelling evidence for the practical applicability of our approach. Despite the possibility that the observed pairwise preference probabilities in practice might not adhere to the BTL model, our filtering subroutine is still able to identify and eliminate corrupted comparisons, while retaining enough of the uncorrupted comparisons to return weight estimates that are very close to the estimates that would have been received given purely uncontaminated data. This strongly contrasts the performance of the existing non-robust algorithms, which return significantly erroneous estimates even for small corruption rates. The results are also promising as they seem to suggest the applicability of our linear programming based pruning approach for corruption rates well beyond what we were able to prove theoretical guarantees for.

5. Conclusion and Discussion

We initiate the study of robustness in rank aggregation under the BTL model by introducing a powerful adversarial contamination model. Within this model, we characterize the exact necessary and sufficient condition for structural identifiability of the true BTL weights in arbitrary comparison graphs. For the family of Erdős-Rényi comparison graphs, we prove a simpler necessary and sufficient condition for identifiability. We also design a linear-programming based recovery algorithm for Erdős-Rényi graphs, which for sparse graphs, has nearly a quadratic runtime, and can tolerate a corruption rate of $O(\log \log n / \log n)$. For denser graphs, it can tolerate a constant corruption rate albeit with a worse runtime. Our work motivates several open problems. Firstly, can we have an efficient recovery algorithm for sparse Erdős-Rényi comparison graphs that improves upon the corruption rate tolerable by our algorithm. Even more generally, can we have a polynomial time recovery algorithm for arbitrary comparison graphs that satisfy the sufficient condition for identifiability, or are there intractability barriers precluding either of these possibilities. Aside from these algorithmic questions, this paper also opens the possibility of considering more restricted contamination models such as ones with oblivious or semi-random adversaries, which could potentially allow us to handle even higher corruption rates.

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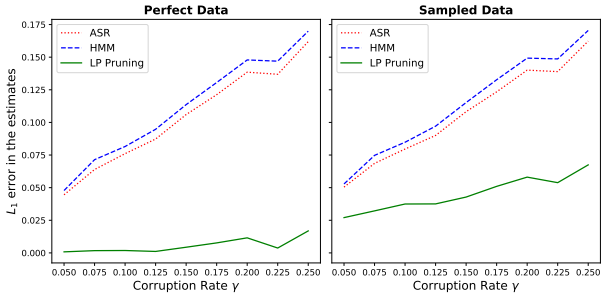

Figure 6. (**Synthetic Data**) L_1 error in the recovered weights vs corruption rate γ in the semi-random contamination model described in Section A.1

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A. Appendix

A.1. Additional Experiments

In addition to the experiments mentioned in the main paper, we experiment with another type adversarial corruption, where given a corruption rate γ, the adversary generates a set of BTL weights $\mathbf{w}^a$, and an Erdős-Rényi random graph $G^a = (V, E^a) \sim G_{n,\gamma p}$, where p is the Erdős-Rényi graph parameter of the initial truthful comparison graph G^* drawn by nature. For every edge $(u, v) \in E^a$, the adversary reports true label consistent with $\mathbf{w}^a$. The input received by our algorithm is $G = G^* \cup G^a$ (if an edge (u, v) is present in both E, E^a, we assume that we receive the label from the uncorrupted graph instead). This is an example of a *semi-random* adversary. We test all algorithms in this more restricted contamination model.

A.1.1. Synthetic Data

We fix $n = 50$, and generate two sets of uniformly at random weights $\mathbf{w}^*, \mathbf{w}^a$, both normalized to sum to 1. Given a corruption rate γ, we generate two Erdős-Rényi random comparison graphs $G^* \sim G_{n,p}$, and $G^a \sim G_{n,\gamma p}$ with parameter $p = (2 \log n)/n$. We generate two datasets: (1) For every uncorrupted edge $(u, v) \in G^*$, we report the exact pairwise probability p^*_{uv} according to $\mathbf{w}^*$, and for every corrupted edge $(u, v) \in G^a$, we report the exact pairwise probability p^a_{uv} according to $\mathbf{w}^a$, and (2) For every uncorrupted edge (u, v), we generate a random sample $X_{uv} \sim \text{Binomial}(L, p^*_{uv})$ and report $p_{uv} = X_{uv}/L$, $p_{vu} = 1 - p_{uv}$, and for every corrupted edge (u, v), we generate a random sample $Y_{uv} \sim \text{Binomial}(L, p^a_{uv})$ and report $p_{uv} = Y_{uv}/L$, $p_{vu} = 1 - p_{uv}$. In our experiments, we set $L = \log n/\epsilon^2$, where $\epsilon = 5\%$ is a chosen accuracy parameter. We test all algorithms on both datasets.

A.1.2. Real Data

For each real dataset (`Sushi` $n = 16$, `Irish` $n = 12$), we create an artificially contaminated dataset as follows: we first generate an adversarial weight vector $\mathbf{w}^a$ by drawing weights uniformly at random, normalizing it to sum to 1. Given a corruption rate γ, we generate two Erdős-Rényi random comparison graphs $G^* \sim G_{n,p}$, and $G^a \sim G_{n,\gamma p}$ with parameter $p = 0.3$. For each uncorrupted edge $(u, v) \in G^*$, we report the empirically observed pairwise probability p_{uv}, and for every corrupted edge $(u, v) \in G^a$, we report the pairwise probability p^a_{uv} according to the adversarial vector $\mathbf{w}^a$. We use this resulting contaminated comparison graph as input to all algorithms.

A.1.3. Experimental Results

In our experiments, we vary γ in the range 5%-25% in increments of 2.5%, and plot the average L_1 error in the returned weight vectors across 50 random trials.

The results again demonstrate the vast difference in performance between our algorithm and the existing non-robust algorithms for parameter recovery in the BTL model, even in this weaker model where the adversary adds random edges into the truthful comparison graph and labels them consistently with his chosen set of BTL weights without corrupting any of the existing edges. For the synthetic data case, given exact pairwise probabilities, our algorithm almost exactly recovers the true BTL parameters, even at substantial corruption rates. Given sampled probabilities, our algorithm significantly outperforms the existing non-robust algorithms. This vast difference in performance is also evident in real datasets, further supporting the practical viability of our proposed algorithm.

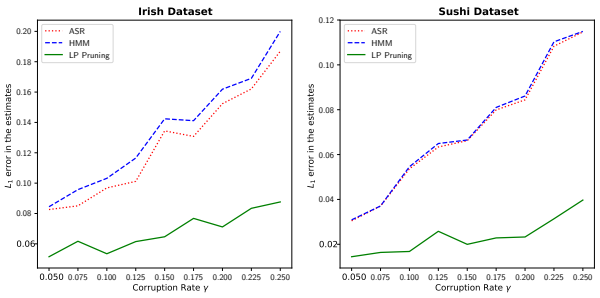

Figure 7. (**Real Data**) L_1 error in the recovered weights vs corruption rate γ in the semi-random contamination model described in Section A.1

A.2. Proof of Theorem 1

Theorem 1. *Given any arbitrary comparison graph $G = (V, E)$ as input, it is possible to uniquely identify the true weights $\mathbf{w}^*$ in the limit $L \to \infty$, if and only if for every cut $(S, V \setminus S)$*

$$|E_u(S, V \setminus S)| > |E_a(S, V \setminus S)|,$$

where $E_u \subseteq E$ is the set of uncorrupted edges, and $E_a = E \setminus E_u$ is the set of adversarially corrupted edges.

Proof. In the limit $L \to \infty$, we have that $\hat{p}_{uv} = p^*_{uv}$ for every edge in the initial truthful graph $G^* = (V, E^*)$. To prove that it is necessary for this condition to hold in order to be able to identify the true weights $\mathbf{w}^*$, consider the following construction: given an uncorrupted graph $G^* = (V, E^*)$ together with pairwise probabilities $\{p_{uv} = p^*_{uv}\}_{(u,v)\in E^*}$, and true set of weights $\mathbf{w}^*$, fix any cut $(S, V \setminus S)$. Let $E^*(S, V \setminus S)$ be the set of edges crossing this cut in G^*. Consider an arbitrary disjoint partition E_1, E_2 of the edges $E^*(S, V \setminus S)$ such that $E^*(S, V \setminus S) = E_1 \cup E_2$. The adversary constructs an obfuscating set of weights $\mathbf{w}^{(\alpha,S)}$ such that for any item $u \in S$, $w^{(\alpha,S)}_u = \alpha w^*_u / w^*_S$, and for any $v \in V \setminus S$, $w^{(\alpha,S)}_v = (1-\alpha) w^*_v / (1 - w^*_S)$, where $w^*_S := \sum_{v\in S} w^*_v$, and $\alpha \neq w^*_S$ is any scaling factor in $(0, 1)$. The adversary then picks $E_a = E_1, E_u = E_2$ or $E_a = E_2, E_u = E_1$ uniformly at random, and for every edge $(u, v) \in E_a$, corrupts the pairwise probability $p_{uv} = w^{(\alpha,S)}_u / \left(w^{(\alpha,S)}_u + w^{(\alpha,S)}_v\right)$ to be consistent with the obfuscating weights $\mathbf{w}^{(\alpha,S)}$. By construction of $\mathbf{w}^{(\alpha,S)}$, one can verify that for every edge $(u, v) \in E^*$ where $u, v \in S$ or $u, v \in V \setminus S$, the pairwise probability p_{uv} is consistent with both $\mathbf{w}^*$ and $\mathbf{w}^{(\alpha,S)}$. Thus, every edge $(u, v) \in E^* \setminus E_a$ is consistent with $\mathbf{w}^*$ and every edge $(u, v) \in E^* \setminus E_u$ is consistent with $\mathbf{w}^{(\alpha,S)}$. Clearly, in the absence of the this condition it is impossible for any algorithm to distinguish E_u from E_a.

To prove that this condition is also sufficient in order to be able to identify the set of uncorrupted edges, suppose for the sake of contradiction, there is some graph $G = (V, E)$ for which there are two sets of weights $\mathbf{w}, \mathbf{w}^* \in \Delta_n, \mathbf{w} \neq \mathbf{w}^*$ such that for every cut $(S, V \setminus S)$, a majority of the edges crossing this cut are consistent with both $\mathbf{w}, \mathbf{w}^*$. We first claim that the two sets of weights $\mathbf{w}, \mathbf{w}^*$ induce equivalence classes on vertices. Formally, we say that vertices $u, v \in V$ belong to the same equivalence class with respect to $\mathbf{w}, \mathbf{w}^*$ if $w_u/w_v = w^*_u/w^*_v$. One can verify that this satisfies symmetry, and transitivity. Intuitively, this says that the two sets of weights are in agreement with each other with respect to the relative qualities of vertices within the same equivalence class, and are in disagreement with each other with respect to the relative qualities of vertices across equivalence classes. The idea is to show that while edges that connect vertices within an equivalence class can be consistent with both $\mathbf{w}, \mathbf{w}^*$, the edges connecting vertices from different equivalence classes can only be consistent with one of $\mathbf{w}, \mathbf{w}^*$. Thus, any cut separating different equivalence classes will act as a certificate of difference between $\mathbf{w}, \mathbf{w}^*$, and hence, a majority of the edges crossing such a cut can be consistent with only one of these sets. Since $\mathbf{w} \neq \mathbf{w}^*$, this partitions V into at least two equivalence classes V_1, V_2. Consider any equivalence class, say V_1 and consider the cut $(V_1, V \setminus V_1)$. For any edge $(u, v) \in E(V_1, V \setminus V_1)$, we claim that p_{uv} is consistent with either $\mathbf{w}$ or $\mathbf{w}^*$, but not both. This is easy to see because if $p_{uv} = w^*_u/(w^*_u + w^*_v) = w_u/(w_u + w_v)$ then $w^*_u/w^*_v = w_u/w_v$, and u, v would belong to the same equivalence class. Thus, the edges $E(V_1, V \setminus V_1)$ are disjointly partitioned into those consistent with $\mathbf{w}$ and those consistent with $\mathbf{w}^*$, and clearly, only one of them can have majority, which is a contradiction. □

A.3. Proof of Fact 1

Fact 1. *Given an arbitrarily small constant $\epsilon > 0$, there exists a sufficiently large constant k, such that given a graph $G = (V, E) \sim G_{n,p}$ with parameter $p \geq (k \log n)/n$, we have for every cut $(S, V \setminus S)$*

$$(1-\epsilon)\,|S||V \setminus S|p < |E(S, V \setminus S)| < (1+\epsilon)\,|S||V \setminus S|p$$

This claim holds with probability at least $1 - 1/\text{poly}(n)$.

Proof. Given any cut $(S, V \setminus S)$ in G, where $s = |S|$, the number of edges crossing the cut $|E(S, V \setminus S)|$ is the sum of $s(n-s)$ independent bernoulli random variables, each with parameter p. The expected number of such edges is $\mathbf{E}_{s,p} = s(n-s)p$. Thus, by the union bound followed by

the Chernoff bound, we have that

$$\begin{aligned}
&\Pr\left(\exists S : |E(S, V \setminus S)| \leq (1-\epsilon)\,\mathbf{E}_{s,p}\right) \\
&\leq \sum_{s=1}^{n/2} \binom{n}{s} \exp\left(-\epsilon^2 \frac{\mathbf{E}_{s,p}}{2}\right) \\
&\leq \sum_{s=1}^{n/2} \exp\left(s \log \frac{en}{s} - \frac{\epsilon^2 sk \log n}{4}\right) \\
&\leq \frac{1}{n^{\frac{\epsilon^2 k}{4}-2}}
\end{aligned}$$

where the second inequality follows from the fact that $p = (k \log n)/n$ and $s \leq n/2$.

Similarly,

$$\begin{aligned}
&\Pr\left(\exists S : |E(S, V \setminus S)| \geq (1+\epsilon)\,\mathbf{E}_{s,p}\right) \\
&\leq \sum_{s=1}^{n/2} \binom{n}{s} \exp\left(-\epsilon^2 \frac{\mathbf{E}_{s,p}}{3}\right) \\
&\leq \sum_{s=1}^{n/2} \exp\left(s \log \frac{en}{s} - \frac{\epsilon^2 sk \log n}{6}\right) \\
&\leq \frac{1}{n^{\frac{\epsilon^2 k}{6}-2}}
\end{aligned}$$

where the second inequality follows from the fact that $p = (k \log n)/n$ and $s \leq n/2$.

By a union bound over the above events, the claim of Fact 1 holds with probability at least $1 - 2/n^{\frac{\epsilon^2 k}{6}-2} = 1 - 1/\text{poly}(n)$ since ϵ is a constant. □

A.4. Proof of Theorem 2

Theorem 2. *Given any arbitrarily small constant $\epsilon > 0$, there exists a sufficiently large constant k, such that given a input comparison graph $G = (V, E)$ conforming to the contamination model in Section 3.1 with Erdős-Rényi graph parameter $p \geq (k \log n)/n$, if the corruption rate $\gamma \leq \frac{1}{4} - \epsilon$, then with probability at least $1 - 1/\text{poly}(n)$, the cut-majority condition described in Theorem 1 is satisfied for every cut in G, and as a consequence, the true weights $\mathbf{w}^*$ are uniquely identifiable as the number of samples per pair $L \to \infty$. Conversely, if the corruption rate $\gamma \geq \frac{1}{4} + \epsilon$, then with probability at least $1 - 1/\text{poly}(n)$, there exists a choice of adversarial corruption such that the cut-majority condition described in Theorem 1 is violated for at least one cut in G, rendering the true weights unidentifiable, even as $L \to \infty$*

Proof. We will start by proving the sufficient condition for recoverability using the cut-majority condition discussed in Lemma 1. Given the arbitrarily small constant $\epsilon > 0$, from Fact 1, we know that there exists a sufficiently large constant k such that given a graph $G^* = (V, E^*) \sim G_{n,p}$ with parameter $p \geq (k \log n)/n$, with probability at least $1 - 1/\text{poly}(n)$, the number of edges crossing any cut in G^* is within $(1 \pm \epsilon)$ of its expected value. Let G^* be the underlying uncorrupted graph satisfying this condition. This also implies that the degree of every vertex in G^* is upper bounded by $(1+\epsilon)np$. Thus, the total number of edges incident on any vertex that the adversary can add, delete or modify (in total) in G is upper bounded by $(1/4 - \epsilon)(1 + \epsilon)np < (1-3\epsilon)np/4$. Now fix any cut $(S, V \setminus S)$ in G^* such that $|S| := s \leq n/2$. We have that the total number of edges crossing the cut is at least $(1-\epsilon)s(n-s)p \geq (1-\epsilon)snp/2$. Furthermore, the total amount of corruption, i.e. addition of spurious edges, deletion, or modification of existing edges in total that the adversary can introduce into this cut in G is upper bounded by $(1-3\epsilon)snp/4$, which is strictly less than half of the total number of edges crossing this cut in G^*. Thus, at least half of the original uncorrupted edges crossing the cut in G^*, survive in G, while at most half this number gets either added, deleted or modified in total. Thus the cut majority condition in Lemma 1 holds for the cut $(S, V \setminus S)$, and this is true for every cut in G, which proves the former claim of Theorem 2.

We will now prove the latter part of the lemma which shows a necessary condition for recoverability. Given the arbitrarily small constant $\epsilon > 0$, let $\epsilon' = \epsilon/3$. From the proof of Fact 1, we know that given a graph $G^* = (V, E^*) \sim G_{n,p}$ with parameter $p \geq (k \log n)/n$, with probability at least $1 - 2/n^{\frac{\epsilon'^2 k}{6}-2}$, the number of edges crossing any cut in G^* is within $(1 \pm \epsilon')$ of its expected value. Let G^* be the underlying uncorrupted graph satisfying this condition. Fix a uniformly at random cut $(S, V \setminus S)$ in G^* such that $|S| := s = n/2$. Using Chernoff bounds, we can claim that with high probability, for every vertex $u \in S$, $|E^*(u, V \setminus S)| \leq (1+\epsilon')np/2$. This is true because for every vertex $u \in S$, the number of edges $|E^*(u, V \setminus S)|$ is the sum of $n/2$ independent Bernoulli random variables, each with parameter p. Thus,

$$\begin{aligned}
&\Pr(\exists u \in S : |E^*(u, V \setminus S)| \geq (1+\epsilon')\frac{np}{2}) \\
&\leq \sum_{u \in S} \exp(-\epsilon'^2 np/6) \\
&\leq \frac{n}{2} \exp(-\epsilon'^2 k \log n/6) \\
&\leq \frac{1}{2n^{\frac{\epsilon'^2 k}{6}-1}}
\end{aligned}$$

The adversary corrupts this cut in the following manner. Given the true set of weights $\mathbf{w}^*$, he constructs an obfuscating set of weights $\mathbf{w}^{(\alpha,S)}$ such that for any $u \in S$, $w_u^{(\alpha,S)} = \alpha w_u^*/w_S^*$ and for any $v \in V \setminus S$, $w_v^{(\alpha,S)} = (1-\alpha)w_v^*/(1-w_S^*)$, where $w_S^* := \sum_{v \in S} w_v^*$, and $\alpha \neq w_S^*$ is any scaling factor in $(0, 1)$. Now for every vertex $u \in S$, the adversary chooses a uniformly at random set of

$(1/2+\epsilon'/2)|E^*(u, V\setminus S)|$ neighbours of u from $V\setminus S$, and corrupts the probabilities on each of these edges to be consistent with $\mathbf{w}^{(\alpha,S)}$. From the above claim, we know that every vertex $u \in S$ has at most $(1+\epsilon')np/2$ neighbours in $V\setminus S$ in G^*, and has $|E^*(v)| \geq (1-\epsilon')np$. Thus, we have that the total fraction of corrupted edges per vertex $u \in S$ is at most $(1+\epsilon')^2/(4(1-\epsilon)) = 1/4+(3\epsilon'+\epsilon'^2)/(4(1-\epsilon')) < 1/4+\epsilon$ (by choice of ϵ'). Furthermore, in the corrupted graph G, for every vertex $u \in S$, a majority of its neighbours in $V\setminus S$ are connected through a corrupted edge, and hence, a majority of the edges crossing the cut $(S, V\setminus S)$ are corrupted, failing the cut-majority condition. The only thing left to show is there is no vertex $v \in V\setminus S$ that has more than $(1/4+\epsilon)|E^*(v)|$ corrupted edges incident upon it. Since each vertex $u \in S$ independently selects a set of uniformly at random $(1+\epsilon')/2$ fraction of its incident *cut edges* to corrupt, for every vertex $v \in V\setminus S$, each of its incident edges from a vertex $u \in S$ is corrupted independently with probability $(1+\epsilon')/2$. Thus, the number of corrupted edges $|E_a(v)|$ incident on any vertex $v \in V\setminus S$ is the sum of $n/2$ independent Bernoulli random variables, each with parameter $(1+\epsilon')p/2$. Thus, the expected number of corrupted edges incident on any vertex $v \in V\setminus S$ is $\mathbf{E}_a(v) = (1+\epsilon')np/4$. Thus, by Chernoff bounds,

$$\begin{aligned}
&\Pr(\exists v \in V\setminus S : |E_a(v)| \geq (1+\epsilon')\mathbf{E}_a(v)) \\
&\leq \sum_{v\in V\setminus S} \exp(-\epsilon'^2\frac{\mathbf{E}_a(v)}{3}) \\
&\leq \frac{n}{2}\exp\left(-\epsilon^2(1+\epsilon')\frac{k\log n}{12}\right) \\
&\leq \frac{1}{2n^{\frac{\epsilon'^2(1+\epsilon')k}{12}-1}}
\end{aligned}$$

Thus, in the corrupted graph G, every vertex $v \in V\setminus S$ has at most $(1+\epsilon')^2np/4$ corrupted edges incident on it. Furthermore, from Fact 1, we have that every vertex $v \in V\setminus S$ has at least $(1-\epsilon')np$ edges incident on it in G^*, giving a corruption rate of at most $(1+\epsilon')^2/(4(1-\epsilon')) = 1/4+(3\epsilon'+\epsilon'^2)/(4(1-\epsilon')) < 1/4-\epsilon$ by choice of ϵ'. By a union bound over the above events, the latter claim of Theorem 2 holds with probability at least $1-3/n^{\frac{\epsilon'^2k}{12}-2} = 1-1/\text{poly}(n)$ since $\epsilon' = \epsilon/3$, and ϵ is a constant.

□

A.5. Proof of Lemma 2

Lemma 2. *In the setting of Theorem 3, with probability at least $1-1/\text{poly}(n)$, we have that the residual graph $\widetilde{G}$ is connected, and furthermore, contains no edges from E_A.*

To prove this Lemma, we first introduce the following supporting Lemmas.

Lemma 4. *Given a graph $G = (V, E) \sim G_{n,p}$ with parameter $p = (k\log n)/n$ for any k larger than some sufficiently large constant, let $E' \subseteq E$ be an **arbitrarily** chosen set of edges such that for every vertex $u \in V$, $|E'(u)| \leq np/20$. Then with probability at least $1-1/\text{poly}(n)$, there exists a path in $G' = (V, E\setminus E')$ of length at most $3+\log n/\log(np)$ between every pair of vertices $u, v \in V$.*

The above lemma shows that the expansion properties of the Erdős-Rényi graphs hold in a *robust* sense. The astute reader will recognize that this claim seems stronger than what we need, considering that the adversarial corruption rate was bounded by $O(\log(np)/\log n)$. However, this is claim is necessary, because there are in fact two sources of edge deletion, the first being the adversary, and the second being the threshold pruning step of our algorithm, which can delete an arbitrary constant fraction of the incident edges on every vertex. The following Lemma shows this.

Lemma 5. *Let $G = (V, E)$ be any input comparison graph conforming to the contamination model in Section 3.1, and $\mathbf{x}$ be any feasible solution to the LP defined in Figure 3. Then for any vertex $u \in V$, we have $|\{E_{lpr} \cup E_a \cup E_r\}(u)| \leq np/20$ with probability at least $1-1/\text{poly}(n)$.*

These two Lemmas will be crucial in proving Lemma 2.

Proof. (of Lemma 2) Given a feasible solution $\mathbf{x}$ to the LP, let $\widetilde{G} = (V, \widetilde{E})$ be the residual subgraph that survives the threshold pruning step. By a direct application of Lemmas 5, 4, we can conclude that in the surviving graph $\widetilde{G}$, with probability at least $1-1/\text{poly}(n)$, there will exist a path of length at most $3+\log n/\log(np)$ between every pair of vertices $u, v \in V$ consisting of edges only from $E\setminus\{E_{lpr}\cup E_a\cup E_r\} = E_u\setminus E_{lpr}$. Suppose that the surviving set of edges $\widetilde{E}$ contains some significantly corrupted edge $e = (u, v) \in E_A$. From the above claim, we know that there is a path of length at most $3+\log n/\log(np)$ consisting of edges only from E_u, and from Observation 2, this would create an inconsistent cycle C of length at most $4+\log n/\log(np)$. Furthermore, since this cycle consists of edges that survived our rounding scheme, we know that for all edges on this cycle, $x(e) < \log(np)/(5\log n)$, which implies that $\sum_{e\in C} x(e) < 1$. However, this is a contradiction to the claim that $\mathbf{x}$ was a feasible solution to the LP as $C \in \mathbb{C}$. □

A.6. Proof of Lemma 4

Lemma 4. *Given a graph $G = (V, E) \sim G_{n,p}$ with parameter $p = (k\log n)/n$ for any k larger than some sufficiently large constant, let $E' \subseteq E$ be an **arbitrarily** chosen set of edges such that for every vertex $u \in V$, $|E'(u)| \leq np/20$. Then with probability at least $1-1/\text{poly}(n)$, there exists a path in $G' = (V, E\setminus E')$ of length at most $3+\log n/\log(np)$ between every pair of vertices $u, v \in V$.*

Proof. We shall prove this Lemma by first showing that the graph G has vertex expansion of at least $np/4$ for every vertex set S of size $s \leq 1/p$. Given any vertex set S of size $1 \leq s \leq 1/p$, and any vertex $v \in V \setminus S$, we have $\Pr(v \in \delta_E(S)) \geq 1 - e^{-sp}$. Thus, we have $\mathbb{E}(|\delta_E(S)|) \geq (n-s)(1-e^{-sp})$. Given any vertex set S of size $s \leq 1/p$, the size of the vertex neighbourhood $|\delta_E(S)|$ is the sum of $n-s$ independent Bernoulli random variables, each with parameter at least $(1-e^{-sp})$. Thus, by a union bound followed by Chernoff bounds

$$\begin{aligned}
&\Pr(\exists S, s \leq 1/p : |\delta_E(S)| \leq \frac{1}{2}(n-s)(1-e^{-sp})) \\
&\leq \sum_{s=1}^{1/p} \binom{n}{s} \exp\left(-\frac{(n-s)(1-e^{-sp})}{8}\right) \\
&\leq \sum_{s=1}^{1/p} \exp\left(s \log \frac{en}{s} - \frac{(n-s)(1-e^{-sp})}{8}\right) \\
&\leq \sum_{s=1}^{1/p} \exp\left(s \log \frac{en}{s} - \frac{(n-s)(sp - \frac{(sp)^2}{2})}{8}\right) \\
&\leq \sum_{s=1}^{1/p} \exp\left(s \log \frac{en}{s} - \frac{snp}{8}\left(1 - \frac{sp}{2} - \frac{s}{n}\right)\right) \\
&\leq \sum_{s=1}^{1/p} \exp\left(s \log \frac{en}{s} - \frac{sk \log n}{8}\left(\frac{1}{2} - \frac{1}{k \log n}\right)\right) \\
&\leq \frac{1}{np} n^{2 + \frac{1}{8 \log n} - \frac{k}{16}}
\end{aligned}$$

where the second to last inequality follows from the fact that $s \leq 1/p$. Thus, for every vertex set S such that $1 \leq s \leq 1/p$, with high probability, the size of the vertex neighbourhood $|\delta_E(S)| \geq \mathbb{E}(|\delta_E(S)|)/2$. We claim that given any vertex set S of size $1 \leq s \leq 1/p$, $\mathbb{E}(|\delta_E(S)|) \geq snp/2$. To see this, consider

$$\begin{aligned}
&\mathbb{E}(|\delta_E(S)|) \geq (n-s)(1-e^{-sp}) \\
&\mathbb{E}(|\delta_E(A)|) - \frac{snp}{2} \geq n\left(\left(1 - \frac{sp}{np}\right)(1-e^{sp}) - \frac{sp}{2}\right)
\end{aligned}$$

Letting $sp = x$, and $f(x) = (1 - x/np)(1-e^{-x}) - x/2$, we have $f''(x) \leq 0$, which implies that $f(x)$ is concave in x and achieves minimum value at one of the boundaries of its domain $[0,1]$. It can easily be verified that this value is positive at both of these points for sufficiently large k, n, proving the claim. Thus, from the above proof, we can infer that with high probability, the vertex neighbourhood $|\delta_E(S)|$ of any S where $1 \leq s \leq 1/p$ is at least $\mathbb{E}(|\delta_E(S)|)/2$ which is at least $snp/4$. However, we also have that $E'(v) \leq np/20$, due to which $E'(S, V \setminus S) \leq snp/20$. Thus, we can conclude that $|\delta_{E \setminus E'}(S)| \geq snp/4 - snp/20 \geq snp/5$.

Next, we shall prove that for any vertex set S of size $1/p \leq s \leq 8/p$, the size of the vertex neighbourhood $|\delta_E(S)|$ of S is at least $nsp/20 + n/20$. Consider any vertex set S of size $1/p \leq s \leq 8/p$. Recall from our previous argument that the size of the vertex neighbourhood $|\delta_E(S)|$ is the sum of $n-s$ independent Bernoulli random variables, each with parameter at least $(1-e^{-sp})$, due to which we have $\mathbb{E}(|\delta_E(S)|) \geq (n-s)(1-e^{-sp})$. Thus, by Hoeffding's Inequality, we have

$$\begin{aligned}
&\Pr\left(|\delta_E(S)| \leq \frac{n(sp+1)}{20}\right) \\
&\leq \exp\left(-2(n-s)(1-e^{-sp})\left(1 - e^{-sp} - \frac{2(sp+1)}{20(1-s/n)}\right)\right) \\
&\leq \exp\left(-\frac{39n}{20}(1-e^{-sp})\left(1 - e^{-sp} - \frac{4}{39}(sp+1)\right)\right)
\end{aligned}$$

where the final inequality follows from the fact that $(1 - s/n) \geq (1 - 8/np) \geq 39/40$ for $np \geq 320$. Letting $sp = x$, and $f(x) = (1-e^{-x})(1 - e^{-x} - 4(x+1)/39)$, we have that $f''(x) < 0$ in the range $[1, 8]$, and hence, achieves minimum value at one of the boundaries $[1, 8]$. Thus, we have $f(x) \geq \min\{f(1), f(8)\} = f(8)$. Thus, we have

$$\Pr\left(|\delta_E(S)| \leq \frac{n(sp+1)}{20}\right) \leq \exp\left(-\frac{39 f(8) n}{20}\right)$$

Thus, by a union bound,

$$\begin{aligned}
&\Pr\left(\exists S, \frac{1}{p} \leq s \leq \frac{8}{p} : |\delta_E(S)| \leq \frac{n(sp+1)}{20}\right) \\
&\leq \sum_{s=1/p}^{8/p} \binom{n}{s} \exp\left(-\frac{39 f(8) n}{20}\right) \\
&\leq \sum_{s=1/p}^{8/p} \exp\left(s \log \frac{en}{s} - \frac{39 f(8) n}{20}\right) \\
&\leq \sum_{s=1/p}^{8/p} \exp\left(-n\left(\frac{39 f(8)}{20} - \frac{s}{n} \log \frac{en}{s}\right)\right) \\
&\leq \frac{7}{p} \exp\left(-n\left(\frac{39 f(8)}{20} - \frac{8}{np} \log \frac{enp}{8}\right)\right) \\
&\leq \frac{7}{p} \exp\left(-n\left(\frac{39 f(8)}{20} - \frac{\log 40e}{40}\right)\right) \\
&\leq \frac{7n}{320} \exp\left(-\frac{n}{32}\right)
\end{aligned}$$

Where the second to last inequality follows from the fact that $(s/n)\log(en/s)$ is a concave increasing function in the argument s/n, and achieves maximum at the right boundary $s/n = 8/np$, and the final inequality follows from the fact that $np \geq 320$. Thus, we have that with high probability, $|\delta_E(S)|$ of S is at least $snp/20 + n/20$. However, we also have that $E'(v) \leq np/20$, due to which $E'(S, V \setminus S) \leq snp/20$. Thus, we can conclude that $|\delta_{E \setminus E'}(S)| \geq n/20$.

Finally, we shall prove that for any vertex set S of size $8/p < s \leq n/10$, the size of the vertex neighbourhood $\delta_{E\setminus E'}(S)$ of S is at least $n/10$. As before, due to the constraint $E'(v) \leq np/20$, we have that $E'(S, V \setminus S) \leq snp/20$ for any vertex set S of size s. We shall first prove that with high probability, given any vertex set S of size $8/p < s \leq n/10$, then for any subset of vertices $S_0 \subset V \setminus S$ of size $4n/10$, the total number of edges between S and S_0 is $|E(S, S_0)| \geq snp/20$. Given any vertex sets $S, S_0 \subset V \setminus S$, $|S_0| = 4n/10$, the total number of edges $|E(S, S_0)|$ is the sum of $4sn/10$ Bernoulli random variables, each with parameter p. Thus, its expected value is $4snp/10$. Thus, by a union bound, followed by Chernoff Bounds, we have

$$\begin{aligned}
&\Pr\left(\exists S, S_0 : |E(S, S_0)| \leq \frac{snp}{20}\right) \\
&\leq \sum_{s=8/p}^{n/10} \binom{n}{s}\binom{n-s}{4n/10} \exp\left(-\left(\frac{7}{8}\right)^2 \frac{4snp}{20}\right) \\
&\leq \sum_{s=8/p}^{n/10} \exp\left(-\frac{49snp}{320} + \frac{4n}{10}\log\frac{10e}{4}\left(1-\frac{s}{n}\right) + s\log\frac{en}{s}\right) \\
&\leq \sum_{s=8/p}^{n/10} \exp\left(-n\left(\frac{49sp}{320} - \frac{4}{10}\log\frac{10e}{4} - \frac{1}{10}\log 10e\right)\right)
\end{aligned}$$

Where the final inequality follows from the fact that $(s/n)\log(en/s)$ is a concave increasing function in the argument s/n, and achieves maximum at the right boundary $s/n = 1/10$. Substituting $sp \geq 8$, we get

$$\begin{aligned}
&\sum_{s=8/p}^{n/10} \exp\left(-\frac{n}{40}\left(49 - 16\log\frac{10e}{4} - 4\log 10e\right)\right) \\
&\leq \frac{24n}{320}\exp\left(-\frac{n}{8}\right)
\end{aligned}$$

where the final inequality follows from the fact that $np \geq 320$. This implies that with high probability, given any set S of size $8/p \leq s \leq n/10$, any set of $4n/10$ vertices from the remaining $n - s$ vertices must have atleast $snp/20$ edges going into it from S. This implies two things: (1) The number of vertices that are not in the vertex neighbourhood $\delta_E(S)$ of S is at most $4n/10$. The reason why this is true is because if the number of vertices that are not in the vertex neighbourhood of S exceeds $4n/10$, then this set would be an example of a set with fewer than $snp/20$ edges going into it from S (it would in fact have 0 edges going into it). (2) To disconnect any set of $4n/10$ vertices from the vertex neighbourhood of S, the entire edge deletion budget must be consumed. This implies that $|\delta_{E\setminus E'}(S)| \geq n - s - 8n/10$. However, $n - s \geq 9n/10$, and thus, $|\delta_{E\setminus E'}(S)| \geq n/10$.

Now consider the graph $G' = (V, E \setminus E')$. For every vertex $u \in V$, let $\delta^d_{E\setminus E'}(u)$ be the set of vertices within distance d of u in the graph G'. From the above proof, we have that with high probability, $\delta^D_{E\setminus E'}(u) \geq n/10$ where $D = 2 + \log n/\log(np)$. This follows by performing a BFS starting from u in G', and applying the vertex expansion condition to the BFS tree with depth $1 \leq d \leq D$. For a fixed pair of vertices (u, v), if v was not already in $\delta^D_{E\setminus E'}(u)$, then we have that the number of edges $|E(\delta^D_{E\setminus E'}(u), v)|$ is the sum of at least $n/10$ independent Bernoulli random variables, each with parameter p, due to which its expected value is at least $np/10$. In order to disconnect vertex v from $\delta^D_{E\setminus E'}(u)$, we would need all edges $E(\delta^D_{E\setminus E'}(u), v) \subseteq E'$, and due to constraint $|E'(v)| \leq np/20$, this would only be possible if $|E(\delta^D_{E\setminus E'}(u), v)| \leq np/20$. Thus, by a union bound followed by Chernoff Bounds, we have

$$\begin{aligned}
&\Pr(\exists u, v : |E(\delta^D_{E\setminus E'}(u), v)| \leq \frac{np}{20}) \\
&\leq n^2 \exp\left(-\left(\frac{1}{2}\right)^2 \frac{np}{20}\right) \\
&\leq n^2 \exp\left(\frac{-k\log n}{80}\right) \\
&\leq n^{2-\frac{k}{80}}
\end{aligned}$$

By a union bound over all the above events, we can conclude that for $np = k\log n \geq 320$ with probability

$$\begin{aligned}
&\geq 1 - \frac{1}{320}\left(n^{2+\frac{1}{8\log n}-\frac{k}{16}} + \frac{7n}{e^{\frac{n}{32}}} + \frac{24n}{e^{\frac{n}{8}}} + 320n^{2-\frac{k}{80}}\right) \\
&= 1 - \frac{1}{\text{poly}(n)},
\end{aligned}$$

every pair of vertices are connected within $G' = (V, E \setminus E')$ with a path of length at most $3 + \log n/\log(np)$. □

A.7. Proof of Lemma 5

Lemma 5. *Let $G = (V, E)$ be any input comparison graph conforming to the contamination model in Section 3.1, and $\mathbf{x}$ be any feasible solution to the LP defined in Figure 3. Then for any vertex $u \in V$, we have $|\{E_{lpr} \cup E_a \cup E_r\}(u)| \leq np/20$ with probability at least $1 - 1/\text{poly}(n)$.*

Proof. For any fixed arbitrarily small constant $0 < \epsilon \leq 1/30$, let the initial comparison graph $G^* = (V, E^*) \sim G_{n,p}$ satisfy the condition specified by Fact 1, and let $G = (V, E)$ the subsequent adversarially contaminated graph conforming to the contamination model specified in Section 3.1 that is received as input to our algorithm. The subsequent results follow deterministically, assuming G^* satisfies the condition specified by Fact 1, which itself holds with probability at least $1 - 1/\text{poly}(n)$.

Recall that E is the set of all edges returned by the adversary, out of which $E_u \subseteq E^*$ is the set of uncorrupted

edges from the original comparison graph, E_a is the set of edges that were either introduced to G by the adversary or already existed in E^*, but were subsequently corrupted by the adversary, and $E_r = E^* \setminus \{E_u \cup E_a\}$ is the set of edges from E^* deleted by the adversary. Recall that, given any solution x to the LP in Figure 3, we delete any edge e such that $x(e) \geq \log(np)/(5\log n)$. Observe that per vertex u, we have $|E_{lpr}(u)| \leq |E(u)|/25$. This is straightforward to prove, because if this were to be untrue for some vertex $v \in V$, then the constraint $\sum_{e \in E(v)} x(e) \leq \gamma_{\text{LP}}|E(v)|$ would be violated. By assumption about G^*, we have $|E^*(u)| \leq (1+\epsilon)np$ and subsequently, by Eq. (1), we have $|E(u)| \leq (1+\epsilon)(1+\gamma)np$. By the same constraint on the corruption rate, we have for any vertex u, $|\{E_a \cup E_r\}(u)| \leq \gamma(1+\epsilon)np$. Thus, we have $|\{E_{lpr} \cup E_a \cup E_r\}(u)| \leq (1+\epsilon)(1+\gamma)np/25 + \gamma(1+\epsilon)np \leq (1+\epsilon)(1+\gamma+25\gamma)np/25 < np/20$ for $\epsilon \leq 1/30$. Thus, for every vertex u, the algorithm and the adversary combined can discard or corrupt at most $np/20$ edges incident on u in E^*. □

A.8. Proof of Lemma 1

Lemma 1. *The LP in Fig 3 is solvable in $O(n^{2+o(1)}d^6)$ time where d is the average degree in the input graph.*

Proof. For any fixed arbitrarily small constant $0 < \epsilon \leq 1/30$, let the initial comparison graph $G^* = (V, E^*) \sim G_{n,p}$ satisfy the condition specified by Fact 1, and let $G = (V, E)$ the subsequent adversarially contaminated graph conforming to the contamination model specified in Section 3.1 that is received as input to our algorithm. The subsequent results follow deterministically, assuming G^* satisfies the condition specified by Fact 1, which itself holds with probability at least $1 - 1/\text{poly}(n)$.

We begin by showing that the total number of decision variables and constraints is small. Since the inconsistent cycle constraints dominate the rest of the constraints, we begin by proving that the total number of cycles in G of length at most $4 + \log n/\log(np)$ is just quadratic. From our assumption about G^*, the degree of every vertex in G^* is bounded by $(1+\epsilon)np$. Furthermore, the adversary can add at most a γ fraction of the realized edges as new adversarially corrupted edges, giving us an upper bound of $(1+\epsilon)(1+\gamma)np$ on the degree of every vertex in G. Thus, the number of cycles of length at most $4 + \log n/\log(np)$ that any vertex participates in is bounded by

$$\begin{aligned}
&\sum_{l=1}^{4+\log n/\log(np)} ((1+\epsilon)(1+\gamma)np)^l \\
&\leq \left(4 + \frac{\log n}{\log(np)}\right) e^{1+4\epsilon} n^{1+\frac{4+\epsilon}{\log(np)}} (np)^4 \\
&= O(n^{1+o(1)}(np)^4)
\end{aligned}$$

where the second inequality follows from the bound $\gamma \leq \log(np)/\log n$. Thus, the total number of cycles of length at most $4 + \log n/\log(np)$ is bounded by $O(n^{2+o(1)}(np)^4)$, and thus, the total number of constraints is bounded by $O(n^{2+o(1)}(np)^4)$, and the total number of decision variables is bounded by $O(n(np))$. The latter claim follows directly from our assumption about G^*.

Next, we utilize the fact that the Multiplicative Weight Update framework for approximately solving Linear Programs is extremely efficient. Our problem is further simplified by the fact that we do not need to solve the entire minimization problem. In fact, the latter steps in our proposed algorithm provably work given *any* feasible solution to the Linear Program described in Figure 3, which is one of the reasons why the MWU method becomes a compelling approach. Given a system of linear constraints $\mathbf{A}^\top \mathbf{x} \geq \mathbf{b}, \mathbf{x} \in C$ for some convex domain C, along with an oracle that takes as input a single constraint $\alpha^\top \mathbf{x} \geq \beta$ and either (1) returns an $\mathbf{x}^* \in C$ that satisfies the constraint $\alpha^\top \mathbf{x}^* \geq \beta$, or (2) correctly determines that the constraint is infeasible for all $\mathbf{x} \in C$, the MWU method either returns an $\mathbf{x}' \in C$ such that $\mathbf{A}^\top \mathbf{x}' \geq \mathbf{b} - \delta$, or correctly determines that the system is infeasible within at most $O((\rho^2 \log m)/\delta^2)$ calls to the said oracle. Here, m is the number of constraints in the system, $\delta > 0$ is the desired approximation factor, and ρ is a quantity known as the *width* of the Linear Program. Thus, the overall running time of this approach is $O(R\rho^2 \log m/\delta^2)$, where R is the runtime of a single call to the said oracle. We will refer the interested reader to (Plotkin et al., 1995) for a comprehensive study of this framework.

In our proposed LP, we are interested in finding an $\mathbf{x}$ in the domain $[0,1]^{|E|}$ such that $\mathbf{A}^\top \mathbf{x} \geq \mathbf{b}$, where $\mathbf{A}^\top$, $\mathbf{b}$ is a compact way of representing our constraints. To be precise, the matrix $\mathbf{A}^\top$ has $|E|$ columns, where $|E|$ is the number of edges in G, and $|\mathbb{C}| + n$ rows, with the first $|\mathbb{C}|$ rows corresponding to the inconsistent cycle constraints, and the subsequent n rows corresponding to the vertex constraints. Furthermore, since this feasible set is provably non-empty, such an $\mathbf{x}$ always exists. For our problem, the oracle is very simple: given a single constraint $\alpha^\top \mathbf{x} \geq \beta$, the constraint is always satisfiable if $\beta \leq 0$, as the all 0 vector $\mathbf{x} = \mathbf{0}^{|E|}$ satisfies it. If on the other hand, $\beta > 0$, then the constraint is satisfiable if and only if the total sum of the positive entries within α exceeds β, i.e. $\sum_{i \in [|E|]} \alpha_i \mathbf{1}(\alpha_i > 0) \geq \beta$. In this case, the solution

$\mathbf{x} = \mathbf{1}(\alpha > 0)$ satisfies the constraint. In all other cases, the constraint is unsatisfiable. As one can observe, this oracle is very efficient, and runs in time $O(|E|)$. Furthermore, constructing the meta-constraint that would be given to the oracle in a single call takes time $O(\texttt{nnz}(\mathbf{A}^\top))$, the number of non-zero entries in our constraint matrix $\mathbf{A}^\top$. This is easy to bound, as the number of non-zero variables in a single inconsistent cycle constraint is at most $4 + \log n/\log(np)$, and there are at most $O(n^{2+o(1)}(np)^4)$ such constraints. Furthermore, there are exactly n vertex constraints, each having $O(np)$ non-zero variables. This gives a bound $\texttt{nnz}(\mathbf{A}^\top) = (4 + \log n/\log(np))O(n^{2+o(1)}(np)^4) + nO(np) = O(n^{2+o(1)}(np)^4)$. Thus, the runtime R of a single call to our oracle is $O(n^{2+o(1)}(np)^4)$.

The key factor in the running time of the MWU method is the *width* ρ of the Linear Program, which is defined as the maximum absolute slack in any constraint in our convex domain C. Specifically, given a system of constraints $\mathbf{A}^\top\mathbf{x} \geq \mathbf{b}$, $\mathbf{x} \in C$ for some convex domain C, $\rho = \max\{1, \max_{i\in[m],\mathbf{x}\in C} |\mathbf{A}_i^\top\mathbf{x} - b_i|\}$. In our problem, in the case of inconsistent cycle constraints, we have that $\mathbf{A}_i^\top$ consists of 0s except for at most $4 + \log n/\log(np)$ locations that have 1s, and $b_i = 1$. The since $\mathbf{x} \in [0,1]^{|E|}$, we have $|\mathbf{A}_i^\top\mathbf{x} - b_i| \leq 3 + \log n/\log(np)$. In the case of vertex constraints, for any vertex $v \in V$, we have that $\mathbf{A}_i^\top$ consists of 0s except for $|E(v)|$ locations that have -1s, and $b_i = -\gamma_{\text{LP}}|E(v)|$, where $|E(v)|$ is the number of edges incident on vertex v. Since $\mathbf{x} \in [0,1]^{|E|}$, we have $|\mathbf{A}_i^\top\mathbf{x} - b_i| \leq \max\{\gamma_{\text{LP}}|E(v)|, (1-\gamma_{\text{LP}})|E(v)|\} \leq O(np)$ from Fact 1. Thus, we have $\rho^2 \leq O((np)^2)$.

The only concern left is that this approach would only return a δ-approximately feasible solution. However, this is also easy to handle. We simply set δ to some small constant, say $1/100$, and slightly lower our rounding threshold by a $1-\delta$ factor, i.e. round all edges with value $x(e) \geq (1-\delta)\log(np)/(5\log n)$. It is straightforward to verify that this slight adjustment would not affect the correctness of our algorithm; our rounding scheme would still delete every corrupted edge, without discarding too many edges per vertex.

From the above discussion, we can conclude that the MWU method will find a feasible fractional solution to our Linear Program in time $O(n^{2+o(1)}(np)^6)$, which is just quadratic in the sparse data regime. □

A.9. Proof of Observation 1

Observation 1. *A solution that assigns $x(e) = 1$ to every edge $e \in E_a$, the set of adversarially corrupted edges, and $x(e) = 0$ to every edge $e \in E_u$, the set of uncorrupted edges is a feasible solution to the above LP.*

Proof. Note that for $b \geq \max_{i,j\in[n]} w_i^*/w_j^*$, an additive error of at most $\epsilon_L/(1+b)$ corresponds to a multiplicative error of at most ϵ_L, i.e. for every pair $(u,v) \in E_u$, $|p_{uv} - p^*_{uv}| \leq \epsilon_L/(1+b) \Rightarrow (1-\epsilon_L)p^*_{uv} \leq p_{uv} \leq (1+\epsilon_L)p^*_{uv}$, and the same holds for (v,u).

We shall prove this observation by showing that no cycle of length at most $4 + \log n/\log(np)$ and consisting of edges only from E_u can be inconsistent. To prove this claim, consider any cycle $C = (v_{c_1}, \ldots v_{c_l}, v_{c_{l+1}} = v_{c_1})$ of length l, where $\forall 1 \leq i \leq l, (v_{c_i}, v_{c_{i+1}}) \in E_u$. We have

$$\begin{aligned}
\prod_{i=1}^{l} \frac{p_{v_{c_i}v_{c_{i+1}}}}{p_{v_{c_{i+1}}v_{c_i}}} &\leq \left(\frac{1+\epsilon_L}{1-\epsilon_L}\right)^l \prod_{i=1}^{l} \frac{p^*_{v_{c_i}v_{c_{i+1}}}}{p^*_{v_{c_{i+1}}v_{c_i}}} \\
&= \left(1 + \frac{2\epsilon_L}{1-\epsilon_L}\right)^l \prod_{i=1}^{l} \frac{w^*_{v_{c_i}}}{w^*_{v_{c_{i+1}}}} \\
&\leq 1 + \frac{2l\epsilon_L}{1-(2l-1)\epsilon_L} \\
&= \frac{1+\epsilon_L}{1-(2l-1)\epsilon_L}
\end{aligned}$$

Where the first inequality follows from the assumption that $\forall 1 \leq i \leq l, (v_{c_i}, v_{c_{i+1}}) \in E_u$, and the final inequality follows from the bound $(1+x)^r \leq 1 + rx/(1-(r-1)x)$ for $x \in [-1, 1/(r-1))$.

Similarly, we have

$$\begin{aligned}
\prod_{i=1}^{l} \frac{p_{v_{c_i}v_{c_{i+1}}}}{1-p_{v_{c_i}v_{c_{i+1}}}} &\geq \left(\frac{1-\epsilon_L}{1+\epsilon_L}\right)^l \prod_{i=1}^{l} \frac{p^*_{v_{c_i}v_{c_{i+1}}}}{p^*_{v_{c_{i+1}}v_{c_i}}} \\
&= \left(1 - \frac{2\epsilon_L}{1+\epsilon_L}\right)^l \prod_{i=1}^{l} \frac{w^*_{v_{c_i}}}{w^*_{v_{c_{i+1}}}} \\
&\geq 1 - \frac{2l\epsilon_L}{1+\epsilon_L} \\
&\geq \frac{1-(2l-1)\epsilon_L}{1+\epsilon_L}
\end{aligned}$$

Where the first inequality follows from the assumption that $\forall 1 \leq i \leq l, (v_{c_i}, v_{c_{i+1}}) \in E_u$, and the final inequality follows from the bound $(1-x)^l \geq 1 - lx$ for $x \in [0,1]$. Thus, every inconsistent cycle must contain at least 1 edge from E_a, and hence the assignment $x(e) = 0 \forall e \in U$ and $x(e) = 1 \forall e \in E_a$ is a feasible assignment. □

A.10. Proof of Observation 2

Observation 2. *For any edge $(u,v) \in E_A$, any path from u to v consisting of edges only from E_u of length at most $4 + \log n/\log(np)$ will induce an inconsistent cycle.*

Proof. Let $l_m = 4 + \log n/\log(np)$. To prove this observation for any edge $(u,v) \in E_A$, consider any cycle $C = (v_{c_1}, v_{c_2}, v_{c_3}, \ldots, v_{c_l}, v_{c_{l+1}} = v_{c_1})$ of length $l \leq l_m$

such that $v_{c_1} = u$, $v_{c_2} = v$, and for all $2 \le i \le l, (v_{c_i}, v_{c_{i+1}}) \in E_u$.

Case 1: Suppose $(p_{uv} - p^*_{uv}) > 2(2l_m - 1)\epsilon_L$

$$\begin{aligned}
\prod_{i=1}^{l} \frac{p_{v_{c_i} v_{c_{i+1}}}}{p_{v_{c_{i+1}} v_{c_i}}} &= \frac{p_{uv}}{p_{vu}} \prod_{i=2}^{l} \frac{p_{v_{c_i} v_{c_{i+1}}}}{p_{v_{c_{i+1}} v_{c_i}}} \\
&> \frac{(1 + 2(2l_m - 1)\epsilon_L) p^*_{uv}}{(1 - 2(2l_m - 1)\epsilon_L) p^*_{vu}} \prod_{i=2}^{l} \frac{(1 - \epsilon_L) p^*_{v_{c_i} v_{c_{i+1}}}}{(1 + \epsilon_L) p^*_{v_{c_{i+1}} v_{c_i}}} \\
&\ge \frac{1 + 2(2l_m - 1)\epsilon_L}{1 - 2(2l_m - 1)\epsilon_L} \left(\frac{1 - \epsilon_L}{1 + \epsilon_L}\right)^{l} \left(\frac{1 + \epsilon_L}{1 - \epsilon_L}\right) \\
&\ge \left(\frac{1 + 2(2l_m - 1)\epsilon_L}{1 - 2(2l_m - 1)\epsilon_L}\right) \left(\frac{1 - (2l - 1)\epsilon_L}{1 - \epsilon_L}\right) \\
&\ge \frac{1 + 2(2l_m - 1)\epsilon_L}{(1 - (2l_m - 1)\epsilon_L)^2} \left(\frac{1 - (2l - 1)\epsilon_L}{1 - \epsilon_L}\right) \\
&\ge \frac{1 + 2(2l_m - 1)\epsilon_L}{1 - \epsilon_L} \frac{1}{1 - (2l_m - 1)\epsilon_L} \\
&\ge \frac{1 + \epsilon_L}{1 - (2l - 1)\epsilon_L}
\end{aligned}$$

where the first inequality follows from the assumption of Case 1, and that for all $2 \le i \le l, (v_{c_i}, v_{c_{i+1}}) \in E_u$ due to which for all $2 \le i \le l, p_{v_{c_i} v_{c_{i+1}}} \ge (1 - \epsilon_L) p^*_{v_{c_i} v_{c_{i+1}}}$ and $p_{v_{c_{i+1}} v_{c_i}} \le (1 + \epsilon_L) p^*_{v_{c_{i+1}} v_{c_i}}$. The third inequality follows from the bound $(1 - x)^l \ge 1 - lx$ for $x \in [0, 1]$. The final three inequalities follow from the fact that $(1-x)^2 > 1-2x$, and $l_m \ge l$.

Case 2: Suppose $(p^*_{uv} - p_{uv}) > 2(2l_m - 1)\epsilon_L$

$$\begin{aligned}
\prod_{i=1}^{l} \frac{p_{v_{c_i} v_{c_{i+1}}}}{p_{v_{c_{i+1}} v_{c_i}}} &= \frac{p_{uv}}{p_{vu}} \prod_{i=2}^{l} \frac{p_{v_{c_i} v_{c_{i+1}}}}{p_{v_{c_{i+1}} v_{c_i}}} \\
&< \frac{(1 - 2(2l_m - 1)\epsilon_L) p^*_{uv}}{(1 + 2(2l_m - 1)\epsilon_L) p^*_{vu}} \prod_{i=2}^{l} \frac{(1 + \epsilon_L) p^*_{v_{c_i} v_{c_{i+1}}}}{(1 - \epsilon_L) p^*_{v_{c_{i+1}} v_{c_i}}} \\
&\le \frac{1 - 2(2l_m - 1)\epsilon_L}{1 + 2(2l_m - 1)\epsilon_L} \left(\frac{1 + \epsilon_L}{1 - \epsilon_L}\right)^{l-1} \left(\frac{1 - \epsilon_L}{1 + \epsilon_L}\right) \\
&\le \left(\frac{1 - 2(2l_m - 1)\epsilon_L}{1 + 2(2l_m - 1)\epsilon_L}\right) \left(\frac{1 - \epsilon_L}{1 - (2l - 1)\epsilon_L}\right) \\
&\le \frac{1 - \epsilon_L}{1 + 2(2l_m - 1)\epsilon_L} (1 - (2l_m - 1)\epsilon_L) \\
&\le \frac{1 - (2l - 1)\epsilon_L}{1 + \epsilon_L}
\end{aligned}$$

where the first inequality follows from the assumption of Case 2, and that for all $2 \le i \le l, (v_{c_i}, v_{c_{i+1}}) \in E_u$ due to which for all $2 \le i \le l, p_{v_{c_i} v_{c_{i+1}}} \le (1 + \epsilon_L) p^*_{v_{c_i} v_{c_{i+1}}}$ and $p_{v_{c_{i+1}} v_{c_i}} \ge (1 - \epsilon_L) p^*_{v_{c_{i+1}} v_{c_i}}$. The third inequality follows from the bound $(1 + x)^r \le 1 + rx/(1 - (r - 1)x)$ for $x \in [-1, 1/(r - 1))$.The final three inequalities follow from the fact that $(1 - x)^2 > 1 - 2x$, and $l_m \ge l$.

In either case, we can observe that cycle C is inconsistent, proving the claim. □

A.11. Proof of Lemma 3

Lemma 3. *In the setting of Theorem 3, let $\mathbf{w}^*$ be the set of true BTL weights, and let $\mathbf{w}$ be the estimate returned by the ASR algorithm with input $\widetilde{G} = (V, \widetilde{E})$. Then we have that*

$$||\mathbf{w} - \mathbf{w}^*||_1 \le (Cb \log b)\epsilon_L$$

where b is an upper bound on $\max_{i,j \in [n]} w^*_i / w^*_j$. This claim holds with probability at least $1 - 1/\mathrm{poly}(n)$.

For any fixed arbitrarily small constant $0 < \epsilon \le 1/30$, let the initial comparison graph $G^* = (V, E^*) \sim G_{n,p}$ satisfy the condition specified by Fact 1, and let $G = (V, E)$ the subsequent adversarially contaminated graph conforming to the contamination model specified in Section 3.1 that is received as input to our algorithm. The subsequent results follow deterministically, assuming G^* satisfies the condition specified by Fact 1, which itself holds with probability at least $1 - 1/\mathrm{poly}(n)$.

Before proving this lemma, we shall first set up some notation. As stated earlier in the main paper, given comparison graph $\widetilde{G} = (V, \widetilde{E})$, the ASR algorithm defines a lazy random walk over $\widetilde{G}$ with probability of transition $\widetilde{P}_{uv}$ from vertex u to vertex v given by

$$\widetilde{P}_{uv} = \begin{cases} \frac{1}{\widetilde{d}_u} p_{vu} & \text{if } u \ne v,\ (u, v) \in \widetilde{E}, \\ \frac{1}{\widetilde{d}_u} \sum_{v \in \delta_{\widetilde{E}}(u)} p_{uv} & \text{if } u = v, \\ 0 & \text{otherwise.} \end{cases} \tag{2}$$

where $\widetilde{d}_u$ is the degree $|\delta_{\widetilde{E}}(u)|$ of vertex u in the graph $\widetilde{G} = (V, \widetilde{E})$. Let $\widetilde{\mathbf{P}} = [\widetilde{P}_{uv}]$ be the transition probability matrix of the corresponding Markov chain. The estimate $\mathbf{w}$ returned by the ASR algorithm is a linear transformation $\mathbf{w} = \widetilde{\mathbf{D}}^{-1}\boldsymbol{\pi}$, where $\boldsymbol{\pi} = \widetilde{\mathbf{P}}^\top \boldsymbol{\pi}$ the stationary distribution of this Markov chain, and $\widetilde{\mathbf{D}}$ is the diagonal matrix of degrees $\widetilde{D}_{uu} = \widetilde{d}_u$.

Furthermore, given comparison graph $\widetilde{G} = (V, \widetilde{E})$, and true BTL weights $\mathbf{w}^*$, we define the true random walk over $\widetilde{G}$ with probability of transition $\widetilde{P}^*_{uv}$ from vertex u to vertex v given by

$$\widetilde{P}^*_{uv} = \begin{cases} \frac{1}{\widetilde{d}_u} p^*_{vu} & \text{if } u \ne v,\ (u, v) \in \widetilde{E}, \\ \frac{1}{\widetilde{d}_u} \sum_{v \in \delta_{\widetilde{E}}(u)} p^*_{uv} & \text{if } u = v, \\ 0 & \text{otherwise.} \end{cases} \tag{3}$$

Let $\widetilde{\mathbf{P}}^* = [\widetilde{P}^*_{uv}]$ be the transition probability matrix of the corresponding true Markov chain. One can verify that given $\widetilde{\mathbf{P}}^*$, one can recover the true BTL weights $\mathbf{w}^*$ from a linear transformation of the stationary distribution of $\widetilde{\mathbf{P}}^*$.

Specifically, we have $\mathbf{w}^* = \widetilde{\mathbf{D}}^{-1}\boldsymbol{\pi}^*$, where $\boldsymbol{\pi}^* = \widetilde{\mathbf{P}^*}^\top \boldsymbol{\pi}^*$ is the stationary distribution of this true Markov Chain.

The key result within (Agarwal et al., 2018) essentially relates the deviation between the estimated weights $\mathbf{w}$ and the true weights $\mathbf{w}^*$ to the deviation between the transition probability matrix $\widetilde{\mathbf{P}}$ constructed from input data, and the true transition probability matrix $\widetilde{\mathbf{P}}^*$, where $\widetilde{\mathbf{P}}$, and $\widetilde{\mathbf{P}}^*$ are as defined in Equations 2, 3, respectively. From Theorems 2,3 and Lemma 6 from (Agarwal et al., 2018), we have

Theorem 4. *Given the true BTL weights $\mathbf{w}^*$, and given a $\widetilde{\mathbf{P}}$ and corresponding $\widetilde{\mathbf{P}}^*$ defined according to Equations 2, 3, respectively, let $\mathbf{w} = \widetilde{\mathbf{D}}^{-1}\boldsymbol{\pi}$, where $\boldsymbol{\pi} = \widetilde{\mathbf{P}}^\top\boldsymbol{\pi}$ the stationary distribution of $\widetilde{\mathbf{P}}$. Then we have that*

$$||\mathbf{w} - \mathbf{w}^*||_{TV} \leq \frac{C}{\mu(\widetilde{\mathbf{P}}^*)} b\delta \log(b\delta)\, ||\widetilde{\mathbf{P}} - \widetilde{\mathbf{P}}^*||_\infty$$

where C is an absolute constant, b is an upper bound on $\max_{i,j\in[n]} w_i^*/w_j^*$, $\delta = \max_{u,v\in V} \widetilde{d}_u/\widetilde{d}_v$, and $\mu(\widetilde{\mathbf{P}}^*) := 1 - \lambda_2(\widetilde{\mathbf{P}}^*)$ is the spectral gap of the Markov Chain defined by $\widetilde{\mathbf{P}}^*$

We shall first prove a bound on $||\widetilde{\mathbf{P}} - \widetilde{\mathbf{P}}^*||_\infty$.

Lemma 6. *Consider the setting of Theorem 3. Let $\widetilde{\mathbf{P}}$ and $\widetilde{\mathbf{P}}^*$ be as defined in Equation 2, and Equation 3. Then we have that*

$$||\widetilde{\mathbf{P}} - \widetilde{\mathbf{P}}^*||_\infty \leq (1 + 1/15 + 31/120)\epsilon_L$$

Proof. By definition of the infinity norm, we have

$$\begin{aligned}||\widetilde{\mathbf{P}} - \widetilde{\mathbf{P}}^*||_\infty &= \max_{u\in[n]} \sum_{v\in[n]} |\widetilde{P}_{uv} - \widetilde{P}^*_{uv}| \\ &\leq \max_{u\in[n]} \frac{1}{\widetilde{d}_u} \sum_{v\in\delta_{\widetilde{E}}(u)} (|p_{uv} - p^*_{uv}| + |p_{vu} - p^*_{vu}|)\end{aligned}$$

Consider any fixed $u \in [n]$. From the proof of Lemma 5, we know that for every vertex $u \in V$, we have that for every vertex $u \in V$, $\widetilde{d}_u \geq (1 - 1/25)d_u$, where $d_u = |\delta_E(u)|$ is the degree of u in $G = (V, E)$. Furthermore, from Equation 1, we have that $|\delta_{E_a}(u)| \leq \gamma d_u$. For this vertex u, let $\delta_{\widetilde{E}\cap E_a}(u)$ be the set of neighbours of u in $\widetilde{G}$ that are connected through a corrupted edge, and let $\delta_{\widetilde{E}\cap E_u}(u)$ be the set of neighbours of u in $\widetilde{G}$ that are connected through an uncorrupted edge. From Lemma 2, we have that for any edge $(u, v) \in \widetilde{E} \cap E_a$, $|p_{uv} - p^*_{uv}| \leq 4l_m\epsilon_L$, where $l_m = 4 + \log n/\log(np)$. Thus, we have that

$$\begin{aligned}&\frac{1}{\widetilde{d}_u} \sum_{v\in\delta_{\widetilde{E}}(u)} (|p_{uv} - p^*_{uv}| + |p_{vu} - p^*_{vu}|) \\ &= \frac{1}{\widetilde{d}_u} \left(\sum_{v\in\delta_{\widetilde{E}\cap E_u}(u)} (|p_{uv} - p^*_{uv}| + |p_{vu} - p^*_{vu}|) \right) \\ &+ \frac{1}{\widetilde{d}_u} \left(\sum_{v\in\delta_{\widetilde{E}\cap E_a}(u)} (|p_{uv} - p^*_{uv}| + |p_{vu} - p^*_{vu}|) \right) \\ &\leq \frac{1}{\widetilde{d}_u} \left(\sum_{v\in\delta_{\widetilde{E}\cap E_u}(u)} \epsilon_L(p_{uv} + p_{vu}) + \sum_{v\in\delta_{\widetilde{E}\cap E_a}(u)} 8l_m\epsilon_L \right) \\ &\leq \frac{25}{24d_u} \left(\left(1 - \gamma - \frac{1}{25}\right) d_u\epsilon_L + 8\gamma d_u l_m \epsilon_L \right) \\ &\leq \epsilon_L \left(1 + \frac{1}{15} + \frac{31}{120}\frac{\log(np)}{\log n}\right)\end{aligned}$$

Where the final upper bound follows by substituting $l_m = 4 + \log n/\log(np)$, and $\gamma \leq \log(np)/(125\log n)$. Clearly, $np \leq n$, and this proves the lemma. □

We shall next prove that $\delta = \max_{u,v\in V} \widetilde{d}_u/\widetilde{d}_v$ is bounded by a constant. From our assumption about G^*, and Equation 1, we have that $(1-\epsilon)(1-\gamma)np \leq |E(u)| \leq (1+\epsilon)(1+\gamma)np$ for every vertex $u \in V$. From Lemma 5, we have that for any vertex $u \in V$, $|\widetilde{E}(u)| = |E(u) \setminus E_{lpr}(u)| \geq 24|E(u)|/25$. The threshold pruning step only deletes edges, due to which for any vertex $u \in V$, $|\widetilde{E}(u)| \leq |E(u)|$. Thus, we have

$$\delta \leq \frac{25(1+\epsilon)(1+\gamma)np}{24(1-\epsilon)(1-\gamma)np} \leq C$$

for some constant C.

To prove a bound on the spectral gap $\mu(\widetilde{\mathbf{P}}^*)$, we shall leverage Lemma 7 of ASR, which, for the case of pairwise comparison data gives that

$$\mu(\widetilde{\mathbf{P}}^*) \geq \frac{\xi}{b}$$

where b is an upper bound on $\max_{i,j\in[n]} w_i^*/w_j^*$, and $\xi := 1 - \lambda_2(\mathbf{L})$ is the spectral gap of the Laplacian $\mathbf{L}$, which is the unweighted random walk on the graph $\widetilde{G}$. Formally, it is defined as $\mathbf{L} := \widetilde{\mathbf{D}}^{-1}\widetilde{\mathbf{A}}$, where $\widetilde{\mathbf{A}}$ is the adjacency matrix of $\widetilde{G}$. We shall prove that the spectral gap ξ of the Laplacian $\mathbf{L}$ is lower bounded by a constant

Lemma 7. *Let $\widetilde{G} = (V, \widetilde{E})$ be the subgraph that survives the threshold pruning step, and let $\mathbf{L} = \widetilde{\mathbf{D}}^{-1}\widetilde{\mathbf{A}}$ be its Laplacian, where $\widetilde{\mathbf{D}}$ is the diagonal matrix where each diagonal entry corresponds to the degree of the corresponding vertex*

in $\widetilde{G}$, and $\widetilde{A}$ is the adjacency matrix of $\widetilde{G}$. Then we have that the spectral gap $\xi = 1 - \lambda_2(\mathbf{L})$

$$\xi \geq C$$

where C is some constant. This statement holds with probability at least $1 - 1/\text{poly}(n)$

Proof. Let $\widetilde{m} = |\widetilde{E}|$, and let ν be the stationary state of $\mathbf{L}$. By definition of the Laplacian, we have for all $(u, v) \in \widetilde{E}$, $\nu_u L_{uv} = 1/2\widetilde{m}$, which follows from the fact that for all $u \in V, \pi_u = \widetilde{d}_u/2\widetilde{m}$. Thus, for any cut $(S, V \setminus S)$ where $s = |S| \leq n/2$, we have the conductance of the cut

$$\Phi(S) = \frac{\sum_{u \in S, v \in V \setminus S} \nu_u L_{uv}}{\sum_{u \in S} \nu_u} = \frac{|\widetilde{E}(S, V \setminus S)|}{\sum_{u \in S} \widetilde{d}_u}$$

Let $E_{lpr} \subset E$ be the subset of edges deleted by our threshold pruning step, and let $E_{lpr}(S) \subseteq E_{lpr}$ be the set of edges in E_{lpr} for which at least one of the end points was in S. Thus, we have

$$\begin{aligned}
\Phi(S) &= \frac{|\widetilde{E}(S, V \setminus S)|}{\sum_{u \in S} \widetilde{d}_u} \\
&\geq \frac{|E(S, V \setminus S)| - |E_{lpr}(S)|}{(\sum_{u \in S} |E(u)|) - |E_{lpr}(S)|} \\
&\geq \frac{25}{24} \frac{|E(S, V \setminus S)|}{\sum_{u \in S} |E(u)|} - \frac{1}{24}
\end{aligned}$$

Where the final inequality follows from the fact that for any vertex $i \in V$, the rounding scheme can delete at most $E(u)/25$ edges from $E(u)$.

From our assumption about G^*, we have that $|E^*(S, V \setminus S)| \geq (1 - \epsilon)s(n - s)p$, and for every vertex $u \in V$, $|E^*(u)| \leq (1 + \epsilon)np$. Due to Equation 1, we have

$$|E(S, V \setminus S)| \geq |E^*(S, V \setminus S)| - \gamma \left(\sum_{u \in S} |E^*(u)| \right)$$

From the same bound, we have for any vertex $u \in V$,

$$|E(u)| \leq (1 + \gamma)|E^*(u)|$$

Thus, we have

$$\begin{aligned}
\Phi(S) &\geq \frac{25}{24} \frac{|E^*(S, V \setminus S)| - \sum_{u \in S} \gamma |E^*(u)|}{\sum_{u \in S} (1 + \gamma)|E^*(u)|} - \frac{1}{24} \\
&\geq \frac{25}{24(1 + \gamma)} \frac{|E^*(S, V \setminus S)|}{\sum_{u \in S} |E^*(u)|} - \frac{\gamma}{1 + \gamma} - \frac{1}{24} \\
&\geq \frac{25}{24} \frac{(1 - \epsilon)(n - s)}{(1 + \epsilon)(1 + \gamma)n} - \frac{\gamma}{1 + \gamma} - \frac{1}{24}
\end{aligned}$$

We have that $\gamma \leq \log(np)/(125 \log n) \leq 1/125$, and $s \leq n/2$. Thus, after bounding the constants, we have

$$\Phi(S) > \frac{31}{60} \frac{(1 - \epsilon)}{(1 + \epsilon)} - \frac{3}{60}$$

By assumption, $\epsilon \leq 1/30$, we have

$$\Phi(S) > \frac{29}{60} - \frac{3}{60} > \frac{13}{30}$$

which is a constant. □

Combining all of these results, we have the claim of Lemma 3

$$||\hat{\mathbf{w}} - \mathbf{w}^*||_{TV} \leq (Cb \log b)\epsilon_L$$

A.12. Concentration Inequalities

Theorem 5 (Chernoff Bounds for Bernoulli Random Variables)**.** *Suppose $X_1, \ldots, X_n$ are i.i.d Bernoulli Random Variables, each with parameter p, then for any $0 \leq \delta \leq 1$, we have*

$$\Pr\left(\sum_{i=1}^{n} X_i \leq (1 - \delta)np \right) \leq \exp\left(-\frac{\delta^2 np}{2} \right),$$

and

$$\Pr\left(\sum_{i=1}^{n} X_i \geq (1 + \delta)np \right) \leq \exp\left(-\frac{\delta^2 np}{3} \right).$$

Theorem 6 (Hoeffding's Inequality for Bernoulli Random Variables)**.** *Suppose $X_1, \ldots, X_n$ are i.i.d Bernoulli random variables, each with parameter p, then for any $\delta \geq 0$, we have*

$$\Pr\left(\left| \sum_{i=1}^{n} X_i - np \right| \geq n\delta \right) \leq 2 \exp\left(-2n\delta^2 \right).$$