Symmetric Perceptrons, Number Partitioning and Lattices

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Abstract

The symmetric binary perceptron (SBP_{κ}) problem with parameter $\kappa: \mathbb{R}_{\geq 1} \to [0,1]$ is an average-case search problem defined as follows: given a random Gaussian matrix $\mathbf{A} \sim \mathcal{N}(0,1)^{n \times m}$ as input where $m \geq n$, output a vector $\mathbf{x} \in \{-1,1\}^m$ such that

$$||\mathbf{A}\mathbf{x}||_{\infty} \leq \kappa(m/n) \cdot \sqrt{m}$$
.

The *number partitioning problem* (NPP $_{\kappa}$) corresponds to the special case of setting n=1. There is considerable evidence that both problems exhibit large computational-statistical gaps.

In this work, we show (nearly) tight *average-case* hardness for these problems, assuming the *worst-case* hardness of standard approximate shortest vector problems on lattices.

- For SBP $_{\kappa}$, statistically, solutions exist with $\kappa(x)=2^{-\Theta(x)}$ (Aubin, Perkins and Zdeborová, Journal of Physics 2019). For large n, the best that efficient algorithms have been able to achieve is a far cry from the statistical bound, namely $\kappa(x)=\Theta(1/\sqrt{x})$ (Bansal and Spencer, Random Structures and Algorithms 2020). The problem has been extensively studied in the TCS and statistics communities, and Gamarnik, Kızıldağ, Perkins and Xu (FOCS 2022) conjecture that Bansal-Spencer is tight: namely, $\kappa(x)=\widetilde{\Theta}(1/\sqrt{x})$ is the optimal value achieved by computationally efficient algorithms.
 - We prove their conjecture assuming the worst-case hardness of approximating the shortest vector problem on lattices.
- For NPP_{κ}, statistically, solutions exist with $\kappa(m) = \Theta(2^{-m})$ (Karmarkar, Karp, Lueker and Odlyzko, Journal of Applied Probability 1986). Karmarkar and Karp's classical differencing algorithm achieves $\kappa(m) = 2^{-O(\log^2 m)}$.
 - We prove that Karmarkar-Karp is nearly tight: namely, no polynomial-time algorithm can achieve $\kappa(m) = 2^{-\Omega(\log^3 m)}$, once again assuming the worst-case subexponential hardness of approximating the shortest vector problem on lattices to within a subexponential factor.

Our hardness results are versatile, and hold with respect to different distributions of the matrix A (e.g., i.i.d. uniform entries from [0, 1]) and weaker requirements on the solution vector \mathbf{x} .

1 Introduction

Symmetric Binary Perceptrons. The *symmetric binary perceptron* (SBP_{κ}) problem, also called the symmetric Ising perceptron problem [JH60, Win61, Cov65, APZ19, BDVLZ20, PX21, ALS21, ALS22, GKPX22, GKPX23, BEAKZ24], is a search problem defined as follows: given a random matrix $\mathbf{A} \sim \mathcal{N}(0,1)^{n\times m}$ with entries chosen i.i.d. from the normal distribution where $m \geq n$, find a binary vector $\mathbf{x} \in \{-1,1\}^m$ such that $\|\mathbf{A}\mathbf{x}\|_{\infty}$ is minimized. More formally, SBP with parameter $\kappa: \mathbb{R}_{\geq 1} \to [0,1]$ asks us to find an $\mathbf{x} \in \{-1,1\}^m$ such that

$$\|\mathbf{A}\mathbf{x}\|_{\infty} < \kappa(m/n) \cdot \sqrt{m}$$
.

Here, the quality of the solution is parameterized as a function of m/n, the so-called (inverse) aspect ratio of the problem.¹

The problem is versatile, and can be defined with respect to different distributions of the matrix A (i.i.d. uniform [0,1] entries is another popular choice) and different requirements on the solution vector \mathbf{x} . The problem also has rich connections to several other fields including the classical subset sum problem and its variants and the problem of discrepancy minimization.

Two natural questions arise: a statistical question and a computational one. The statistical question asks for which κ do solutions exist (with high probability over the choice of the matrix A). Recent works by Aubin, Perkins and Zdeborová [APZ19], Perkins and Xu [PX21], and Abbe, Li and Sly [ALS21] showed sharp statistical thresholds for this problem. In particular, they showed that the threshold for the existence of solutions is

$$\kappa_{\rm stat}(x) = O(2^{-x}).$$

On the other hand, the best solutions found by polynomial-time algorithms satisfy

$$\kappa_{\mathsf{comp}}(x) = \Omega\left(\frac{1}{\sqrt{x}}\right)^{2}$$

This comes from the breakthrough work of Bansal [Ban10] and Bansal and Spencer [BS20] from the closely related field of combinational discrepancy theory.³ Gamarnik, Kızıldağ, Perkins and Xu [GKPX22, GKPX23] studied the large statistical-computational gap scenario in detail and conjectured that no polynomial-time algorithms can achieve a guarantee much better than $\kappa_{\text{comp}}(x)$. In particular, they show that the so-called overlap gap property [MMZ05, AR06, Gl16], which rules out a class of stable algorithms, sets in at

$$\kappa_{\text{overlap}}(x) = O\left(\frac{1}{\sqrt{x \log x}}\right).$$

¹The notation commonly used in the literature to parameterize SBP (and NPP) is slightly different than the notation we choose to use. The aspect ratio is given by $\alpha = n/m$, and instead of writing κ as a function of α (or really, $\kappa = 1/\alpha$, as we do), the roles are flipped, where α is a function of κ . For example, $\kappa(x) = 2^{-x}$ and $\kappa(x) = 1/\sqrt{x}$, in our notation, correspond to $\alpha(\kappa) = 1/\log_2(1/\kappa)$ and $\alpha(\kappa) = \kappa^2$ in the notation of [GKPX22], respectively.

²In an extreme parameter regime where $n = O\left(\sqrt{\log m}\right)$, [TMR20] gives a polynomial-time algorithm achieving discrepancy $2^{-\Omega(\log^2(m)/n)}$, but throughout our paper, for SBP, we will only consider the regime where $n = \omega(\log m)$.

³Their result is established for the case of matrices A with i.i.d. Rademacher entries. Nevertheless, [GKPX22] conjecture that the same guarantee remains true for the case of i.i.d. standard normal entries.

We refer the reader to [GKPX22] for an extensive discussion of the rationale behind their conjecture.

A survey of Gamarnik on the overlap gap property [Gam21] points to the question of whether average-case hardness of perceptron problems (and more) can be based on *worst-case* hardness assumptions. Gamarnik explicitly states that worst-case to average-case reductions for these problems "would be ideal for our setting, as they would provide the most compelling evidence of hardness of these problems" [Gam21].

The first contribution of this work is a proof of the conjecture of [GKPX22] upto lower order terms, under the assumption that standard, well-studied, lattice problems are hard to approximate in the worst case.

Theorem 1 (Informal version of Corollary 2). Let $\varepsilon > 0$ be any constant. Assuming that γ -approximate lattice problems in n dimensions with $\gamma(n) = n^{O(1/\varepsilon)}$ are worst-case hard for polynomial-time algorithms, SBP $_{\kappa}$ with

$$\kappa(x) = \frac{1}{x^{1/2+\varepsilon}}$$

is hard for polynomial-time algorithms. Additionally, assuming near-optimal worst-case hardness of lattice problems, we obtain near-optimal hardness of SBP. In particular, assuming that $\gamma(n)$ -approximate lattice problems require $2^{\omega(n^{1/2-\varepsilon})}$ time to solve in the worst case with $\gamma(n)=2^{n^{1/2-\varepsilon}}$, we have that SBP $_\kappa$ with

$$\kappa(x) = \frac{1}{\sqrt{x} \cdot (\log x)^c}$$

is hard for some constant c > 1.

Lattice problems such as the shortest vector problem and the shortest independent vectors problem have been extensively studied for decades largely for their implications to combinatorial optimization [LLL82, Kan83, Kan87, RR23] and even more so, to cryptography [Ajt96, MR07, Reg09]. Time-approximation tradeoffs for lattice problems are well-known [Sch87]. The best known algorithms for these lattice problems employ *lattice reduction* techniques, based on the LLL and BKZ algorithms [LLL82, Sch87]. They currently all have the following time-approximation trade-off: For a tunable parameter k, in dimension n, it is posible to solve lattice problems with approximation factor $\gamma = 2^{\widetilde{O}(n/k)}$ in time $2^{\widetilde{O}(k)}$, where $\widetilde{O}(\cdot)$ hides polylog factors in n. Equalizing these terms gives a $2^{\widetilde{O}(\sqrt{n})}$ -time algorithm to solve $2^{\widetilde{O}(\sqrt{n})}$ -approximate lattice problems. Despite extensive study in the lattice literature, no better algorithms are known (that improve the exponent by more than a polylog factor). This motivates the assumptions in our theorem statement, both the conservative one and the near-optimal one. (For more discussion on this, see Section 2.2).

Our reduction has several additional features: first, our reduction is versatile and works with respect to different distributions of the matrix A (e.g., i.i.d. uniform entries from [0,1]); and secondly, it shows the hardness not just of computing an SBP solution with 1 - o(1) probability, but indeed with any non-trivial inverse polynomial probability. For more implications of our reduction, we refer the reader to Section 3.2. Moreover, our full reduction is conceptually simple and direct. We discuss the possibility of using other intermediate problems to establish these same results in Section 1.1.3.

Number Partitioning (or Number Balancing). The (average-case) *number partitioning* problem is a special case of SBP and corresponds to setting n = 1 in SBP. Given m random numbers $a_1, \ldots, a_m \sim \mathcal{N}(0, 1)$ with entries chosen i.i.d. from the standard normal distribution, the goal is to find a binary vector $\mathbf{x} \in \{-1, 1\}^m$ such that $|\sum_{i=1}^m x_i a_i|$ is minimized. More formally, NPP with parameter $\kappa : \mathbb{Z}_{>1} \to [0, 1]$ asks us to find an $\mathbf{x} \in \{-1, 1\}^m$ such that

$$\left|\sum_{i=1}^m x_i a_i\right| \le \kappa(m) \cdot \sqrt{m}$$

Number partitioning, as a worst-case problem, was one of the six original NP-complete problems in the classic book of Garey and Johnson [GJ79]. The worst-case version of the problem, where a_1, \ldots, a_m are arbitrary, is closely related to discrepancy problems and has been extensively studied; see [Spe85, Ban10, LM15, LRR17, Rot14, HRRY17]. In particular, [HRRY17] show that there are (worst-case to worst-case) reductions between NPP and worst-case lattice problems.⁴ Their reduction from worst-case lattice problems to NPP [HRRY17, Theorem 9] shows hardness for $\kappa(m) \leq 2^{-m/2}$. Looking ahead, we show computational hardness for the average-case version of NPP and for the much tighter range of $\kappa(m) = 2^{-\text{polylog}(m)}$.

An application of the pigeonhole principle shows that solutions exist, both in the worst case and on average (even with high probability [KKLO86]), for

$$\kappa_{\text{stat}}(m) = 2^{-m}$$
.

However, the best solutions found by polynomial-time algorithms satisfy

$$\kappa_{\mathsf{comp}}(m) = 2^{-O(\log^2 m)}.$$

This comes from the beautiful "differencing algorithm" of Karmarkar and Karp [KK82] which starts with a list; sorts it; replaces the largest and second largest element with their absolute difference; and repeats until a single element is left which the algorithm outputs as the discrepancy of the set of numbers. (An informed reader might already have observed the analogy of Karmarkar-Karp with the Blum-Kalai-Wasserman [BKW03] algorithm, which came much later.) A subsequent work of Yakir [Yak96] proved that their algorithm indeed achieves the claimed discrepancy of $\kappa_{\text{comp}}(m) = 2^{-O(\log^2 m)}$. This remains the best algorithm known to date.

Gamarnik and Kızıldağ [GK21] studied the statistical-computational gap in depth, demonstrated an overlap gap property at

$$\kappa^*(m) = 2^{-\omega \left(\sqrt{m\log m}\right)},$$

implying that the class of stable algorithms will fail to find solutions for such small κ . This leaves open the question of the ground truth: are there improvements to Karmarkar-Karp, stable or not, that efficiently solve NPP $_{\kappa}$ with $\kappa(m) = 2^{-m^{\Omega(1)}}$?

⁴Technically, the problem they consider is the number balancing problem, where they require the weaker condition on the solution \mathbf{x} that $\mathbf{x} \in \{-1, 0, 1\}^m \setminus \{\mathbf{0}\}$ instead of $\mathbf{x} \in \{-1, 1\}^m$. As we explain later (Section 3.2), our reduction works in this setting as well.

Our second contribution is proving that the answer to this question is *no*, in a quantitatively strong way. In particular, under the assumption that standard, well-studied, lattice problems are sub-exponentially hard to approximate in the worst case, we prove that Karmarkar-Karp is tight, up to a logarithmic factor in the exponent.

Theorem 2 (Informal version of Corollary 3). Suppose $\kappa(m) = 2^{-\log^{3+\varepsilon} m}$ for some constant $\varepsilon > 0$. Assuming near-optimal hardness of worst-case lattice problems, then NPP_{κ} is hard for polynomial-time algorithms. In particular, assuming that $\gamma(n)$ -approximate lattice problems in dimension n require $2^{\omega(n^{1/2-\varepsilon})}$ time to solve in the worst case with $\gamma(n) = 2^{n^{1/2-\varepsilon}}$, we have that NPP_{κ} (in dimension m) is hard for poly(m)-time algorithms.

Similar to the case of our SBP result, this theorem is quite versatile. We refer the reader to the technical overview and Section 4 for more details. Like in the case SBP, our full reduction is conceptually simple and direct, and we discuss the possibility of using other intermediate problems to establish these same results in Section 1.1.3.

Adaptive Robustness of Johnson-Lindenstrauss. The Johnson-Lindenstrauss lemma [JL84] states that for all fixed, small, finite sets $S \subseteq \mathbb{R}^m$, for random $\mathbf{A} \sim \mathcal{N}(0,1)^{n \times m}$, the linear map given by \mathbf{A} embeds S into \mathbb{R}^n in a way that approximately preserves all ℓ_2 norms (up to a \sqrt{n} normalization term). Slightly more concretely,

$$\forall \text{ small, finite } S \subseteq \mathbb{R}^m, \quad \Pr_{\mathbf{A} \sim \mathcal{N}(0,1)^{n \times m}} \left[\forall \mathbf{x} \in S, \ \|\mathbf{A}\mathbf{x}\|_2 \approx \sqrt{n} \, \|\mathbf{x}\|_2 \right] = 1 - o(1).$$

This statement crucially relies on the fact the set S that is defined independently of A. For example, even considering only singleton sets S, one can ask whether the order of quantifiers can be switched so that x can be chosen adaptively based on A, in the following sense:

$$\Pr_{\mathbf{A} \sim \mathcal{N}(0,1)^{n \times m}} \left[\forall \mathbf{x} \in \mathbb{R}^m, \|\mathbf{A}\mathbf{x}\|_2 \approx \sqrt{n} \|\mathbf{x}\|_2 \right] \stackrel{?}{=} 1 - o(1).$$

Or even weaker, for a function $\kappa: \mathbb{R}_{>1} \to (0,1]$,

$$\Pr_{\mathbf{A} \sim \mathcal{N}(0,1)^{n \times m}} \left[\forall \mathbf{x} \in \mathbb{R}^m, \|\mathbf{A}\mathbf{x}\|_2 \ge \|\mathbf{x}\|_2 \kappa(m/n) \sqrt{n} \right] \stackrel{?}{=} 1 - o(1).$$

However, this is impossible. For m > n, one can find a vector $\mathbf{x} \in \ker(\mathbf{A})$ (so $S = \{\mathbf{x}\}$), making $\|\mathbf{A}\mathbf{x}\|_2 = 0$ while $\|\mathbf{x}\|_2$ can be arbitrarily large.

A natural question is whether this phenomenon could hold if we constrain $\mathbf{x} \in \mathbb{R}^m$ to some structured set, e.g., $\{-1, 1\}^m$:

$$\Pr_{\mathbf{A} \sim \mathcal{N}(0,1)^{n \times m}} \left[\forall \mathbf{x} \in \{-1,1\}^m, \ \|\mathbf{A}\mathbf{x}\|_2 \ge \kappa(m/n) \sqrt{nm} \right] \stackrel{?}{=} 1 - o(1).$$

This question is exactly the statistical capacity of SBP_{κ}, with the exception that the norm on **Ax** has changed from ℓ_{∞} to ℓ_{2} (which differ by only a \sqrt{n} factor at most).

Therefore, we view SBP_{κ} as defining a natural adaptive robustness question about a certain discretized version of the Johnson-Lindenstrauss lemma. In particular, since SBP_{κ} exhibits a computational-statistical gap, so does this variant of the Johnson-Lindenstrauss lemma. Given the utility of the Johnson-Lindenstrauss lemma in compressed sensing, dimensionality reduction, and more, we believe that this interpretation may inspire connections between worst-case lattice problems and adaptive robustness of downstream applications of the Johnson-Lindenstrauss lemma.

Open Questions and Future Directions. A direct open question raised by our results is to better understand the gap between Theorem 2 and Karmarkar-Karp [KK82]. Specifically, our result shows hardness for $\kappa(m) = 2^{-\log^{3+\epsilon} m}$, but Karmarkar-Karp gives a polynomial time algorithm that achieves $\kappa(m) = 2^{-\Theta(\log^2 m)}$. Can Karmarkar-Karp be improved to $\kappa(m) = 2^{-\Theta(\log^3 m)}$, can the hardness shown in Theorem 2 be improved, or is the truth somewhere in the middle?

Another question one can ask is related to the *asymmetric* binary perceptron (ABP_{κ}) problem, also called the asymmetric Ising perceptron problem, which for $\mathbf{A} \sim \mathcal{N}(0,1)^{n\times m}$, asks to find $\mathbf{x} \in \{-1,1\}^m$ such that

$$Ax \ge \kappa(m/n)\sqrt{m} \cdot 1$$
,

in the sense that for every row $\mathbf{a}_j \in \mathbb{R}^m$ for $j \in [n]$, we want $\mathbf{a}_j^\top \mathbf{x} \ge \kappa(m/n)\sqrt{m}$. (For more details, see e.g., [GKPX22].) Note that unlike SBP_{κ}, setting $\kappa(x) = 0$ still defines a meaningful problem. Does ABP share similar hardness results (from worst-case problems)? If so, are lattice problems the source of such hardness?

More generally, we can ask the *converse* questions of the ones raised in our paper. Can lattice algorithms be used to improve algorithms for NPP or ABP?

We leave all of these as fascinating open questions and future directions of our work.

1.1 Technical Overview

For both of our reductions, we take direct inspiration from worst-case to average-case reductions in the lattice literature [Ajt96, MR07]. Our exposition and proofs closely follow [MR07]. We view our work as bringing techniques from worst-case to average-case reductions in the lattice literature to closely related problems in statistics and discrete optimization.

1.1.1 Reduction to SBP

We outline the main ideas behind the reduction from the worst-case lattice problems to SBP. For an invertible matrix $\mathbf{B} \in \mathbb{R}^{n \times n}$, let $\mathcal{L}(\mathbf{B})$ denote the lattice generated by (the columns of) \mathbf{B} , i.e.,

$$\mathcal{L}(\mathbf{B}) = \{ \mathbf{B}\mathbf{z} : \mathbf{z} \in \mathbb{Z}^n \}.$$

Let $\mathcal{P}(\mathbf{B})$ denote the fundamental parallelepiped of B, given by the set $\mathbf{B}[0,1)^n$.

To illustrate the ideas in our reduction, consider the following (informal) worst-case lattice problem: given some invertible lattice basis $\mathbf{B} \in \mathbb{R}^{n \times n}$, output a (non-zero) vector $\mathbf{s} \in \mathcal{L}(\mathbf{B})$ slightly smaller than the columns in \mathbf{B} . (While this is not quite the worst-case lattice problem we reduce

from, it is simpler to describe this way and captures all of the main ideas. See Definition 4 for more precise details.)

The key idea we employ is *smoothing* using Gaussian measures, as introduced by Micciancio and Regev [MR07]. Specifically, here is the critical point: For an *arbitrary* invertible $\mathbf{B} \in \mathbb{R}^{n \times n}$ and for $\mathbf{u} \sim \mathcal{N}(\mathbf{0}, \sigma^2 \mathbf{I}_n)$, as long as σ is large enough, $\mathbf{u} \pmod{\mathcal{L}(\mathbf{B})}$ looks *statistically* close to uniform over $\mathcal{P}(\mathbf{B})$. This holds as long as σ is larger than the *smoothing parameter* of the lattice $\mathcal{L}(\mathbf{B})$, defined by [MR07], which is basis-independent. Transforming by \mathbf{B}^{-1} , we see that

$$\mathbf{B}^{-1}\mathbf{u} \pmod{\mathbb{Z}^n} \approx U([0,1)^n),$$

where \approx denotes small total variation distance, and $U(\cdot)$ denotes the uniform distribution. Thus, we have converted worst-case structure into average-case structure (that is independent of **B**).

We can repeat this process m times as follows: sample $U \sim \mathcal{N}(\mathbf{0}, \sigma^2 \mathbf{I}_n)^m$, where now

$$\mathbf{B}^{-1}\mathbf{U} \pmod{\mathbb{Z}^{n\times m}} \approx U([0,1)^{n\times m}).$$

We then set $A = B^{-1}U \pmod{\mathbb{Z}^{n \times m}}$ and feed A into the SBP solver. (For simplicity, in this overview, we assume that SBP allows input matrices that are uniform over $[0,1)^{n \times m}$ instead of Gaussian, but in the proof, we resolve this distinction by sampling from an appropriate discrete Gaussian distribution.) We get back some $\mathbf{x} \in \{\pm 1\}^m$ so that $\|\mathbf{A}\mathbf{x}\|_{\infty} \leq \kappa(m/n)\sqrt{m}$, or in other words,

$$Ax - e = 0$$

where $\|\mathbf{e}\|_{\infty} \leq \kappa(m/n)\sqrt{m}$. Since $\mathbf{x} \in \mathbb{Z}^m$, this implies

$$\mathbf{B}^{-1}\mathbf{U}\mathbf{x} - \mathbf{e} = \mathbf{0} \mod \mathbb{Z}^{n \times m}$$

Multiplying by **B** on the left gives

$$Ux - Be = 0 \mod \mathcal{L}(B),$$

or in other words, $\mathbf{U}\mathbf{x} - \mathbf{B}\mathbf{e} \in \mathcal{L}(\mathbf{B})$. To see that $\mathbf{U}\mathbf{x} - \mathbf{B}\mathbf{e}$ is slightly smaller than the vectors in \mathbf{B} , note that $\|\mathbf{U}\mathbf{x}\|_2$ can be upper-bounded by a basis-independent quantity, since σ was chosen basis-independently. The bottleneck term is $\|\mathbf{B}\mathbf{e}\|_2$, which is smaller than columns of \mathbf{B} as long as $\|\mathbf{e}\|_2$ is sufficiently small. If $\kappa(x) = 1/x^{1/2+\varepsilon}$ for $\varepsilon > 0$, then

$$\|\mathbf{e}\|_{\infty} \leq \kappa(m/n)\sqrt{m} = \frac{n^{1/2+\varepsilon}}{m^{\varepsilon}},$$

so as long as we set m large enough so that $m^{\varepsilon} \gg n^{1/2+\varepsilon}$, we can force $\|\mathbf{Be}\|_2$ to be small enough to produce a smaller lattice vector than anything in \mathbf{B} , as desired.

Allowing $\mathbf{x} \in \{-1, 0, 1\}^m$ **instead of** $\mathbf{x} \in \{\pm 1\}^m$. One could define a variant of SBP where we just need $\mathbf{x} \in \{-1, 0, 1\}^m \setminus \{\mathbf{0}\}$ instead of $\mathbf{x} \in \{\pm 1\}^m$. This is an easier problem, as including zero entries in \mathbf{x} is a simple way to decrease $\|\mathbf{A}\mathbf{x}\|_{\infty}$. However, our reduction idea above actually *still* works in this setting. For more details, see Section 3.2.

1.1.2 Reduction to NPP

The reduction to NPP is quite similar to the reduction to SBP, but with one additional trick that converts between vectors and scalars. This trick has been used in [Reg04, BV15] for a similar reason, and we follow their footsteps. Explicitly, it is the *Chinese remainder theorem*: for distinct primes p_1, \ldots, p_n and $q = \prod_{i \in [n]} p_i$, there is a group isomorphism

$$\varphi: \bigoplus_{i\in[n]} \mathbb{Z}/p_i\mathbb{Z} \longrightarrow \mathbb{Z}/q\mathbb{Z}.$$

We will scale this isomorphism so that

$$\widetilde{\varphi}: \bigoplus_{i \in [n]} 1/p_i \cdot \mathbb{Z}/p_i \mathbb{Z} \longrightarrow 1/q \cdot \mathbb{Z}/q \mathbb{Z}$$

is (a) invariant to integer shifts in the input and (b) \mathbb{Z} -linear, in the sense that $\widetilde{\varphi}(\mathbf{x}) = \mathbf{c}^{\top}\mathbf{x} \pmod{1}$ for some $\mathbf{c} \in \mathbb{Z}^n$.

Following the SBP reduction above, we can set $\mathbf{A} = \mathbf{B}^{-1}\mathbf{U} \pmod{\mathbb{Z}^{n\times m}}$ and then appropriately "round" \mathbf{A} to get uniformly random $\mathbf{A}' = [\mathbf{A}]_{\mathbf{p}} \in \left(\bigoplus_{i \in [n]} 1/p_i \cdot \mathbb{Z}/p_i\mathbb{Z}\right)^m$. We then apply $\widetilde{\varphi}$ columnwise to \mathbf{A}' to get $\mathbf{a}' \in (1/q \cdot \mathbb{Z}/q\mathbb{Z})^m \subseteq [0,1)^m$. By adding small uniform noise to \mathbf{a}' , we can get some $\mathbf{a} \sim [0,1)^m$. We feed this into our NPP solver to get some $\mathbf{x} \in \{\pm 1\}^m$ such that $|\mathbf{a}^{\mathsf{T}}\mathbf{x}| \leq \kappa(m)\sqrt{m}$. By using \mathbb{Z} -linearity of $\widetilde{\varphi}$, and assuming κ is sufficiently small so that there are no "wraparound" errors in $\widetilde{\varphi}^{-1}$, we can recover a somewhat smaller lattice vector $\mathbf{s} \in \mathcal{L}(\mathbf{B})$ than we started with, just like in the SBP reduction.

1.1.3 Alternate Reduction Paths

We briefly mention other paths to reduce worst-case lattice problems to SBP and NPP, using existing intermediate problems and known worst-case to average-case reductions.

SBP. Instead of starting from worst-case lattice problems, we could start with a noisy version of the *short integer solutions* (SIS) problem in the ℓ_{∞} norm, defined roughly as follows: Given random $\mathbf{A} \sim (\mathbb{Z}/q\mathbb{Z})^{n\times m}$, output $\mathbf{x} \in \mathbb{Z}^m \setminus \{\mathbf{0}\}$ such that $\|\mathbf{x}\|_{\infty}$ and $\|\mathbf{A}\mathbf{x} \pmod q\|_{\infty}$ are small. The original worst-case to average-case reductions for lattice problems indeed reduce worst-case lattice problems to (average-case) SIS [Ajt96, MR07]. One can reduce (noisy) SIS to SBP by adding small noise $\mathbf{E} \sim U\left([0,1/q)^{n\times m}\right)$ and setting

$$\mathbf{A}' = \frac{1}{q}\mathbf{A} + \mathbf{E} \sim U([0,1)^{n \times m}).$$

(For simplicity, we assume here that SBP allows matrices with i.i.d. U([0,1)) entries instead of standard normal, but we can remove this assumption by sampling from the appropriate discrete Gaussian and scaling appropriately.) Feeding \mathbf{A}' into the SBP solver yields $\mathbf{x} \in \{\pm 1\}^m$ such that $\|\mathbf{A}\mathbf{x}\|_{\infty}$ is small, assuming $q \|\mathbf{E}\mathbf{x}\|_{\infty}$ is also sufficiently small.

NPP. Brakerski and Vaikuntanathan [BV15] define a problem called the *one-dimensional short* integer solutions problem (1D-SIS), defined roughly as follows: Given a $\sim (\mathbb{Z}/q\mathbb{Z})^m$, output $\mathbf{x} \in \mathbb{Z}^m \setminus \{\mathbf{0}\}$ such that $\|\mathbf{x}\|_{\infty}$ and $|\mathbf{a}^{\top}\mathbf{x}| \pmod{q}$ are small, where q is a product of n primes. [BV15] shows a worst-case to average-case reduction from worst-case lattice problems in dimension n to 1D-SIS. One can reduce 1D-SIS to NPP by adding small noise $\mathbf{e} \sim U\left([0,1/q)^m\right)$ and setting

$$\mathbf{a}' = \frac{1}{q}\mathbf{a} + \mathbf{e} \sim U([0,1)^m).$$

(For simplicity, we similarly assume here that NPP works for vectors with i.i.d. U([0,1)) entries.) Feeding \mathbf{a}' into the NPP solver yields $\mathbf{x} \in \{\pm 1\}^m$ such that $|\mathbf{a}^\top \mathbf{x}|$ is small, assuming $q|\mathbf{e}^\top \mathbf{x}|$ is also sufficiently small.

While these approaches would work, we view this introduction of parameter $q \in \mathbb{Z}$ as extraneous and misleading. To show hardness of a discrete optimization problem in continuous Euclidean space (SBP and NPP), we should ideally start from a problem that is itself a discrete optimization problem in continuous Euclidean space (worst-case lattice problems). There is no need to add premature discretization by introducing the parameter q.

2 Preliminaries

For a predicate φ , we use the notation $\mathbf{1}[\varphi] \in \{0, 1\}$ to denote the indicator variable of whether φ is true (1) or false (0). We use $\ln(\cdot)$ to denote the natural logarithm (base e) and $\log(\cdot)$ to denote $\log_2(\cdot)$. We use $\mathbb{R}_{>0}$ to refer to the set of all positive real numbers, and we use $\mathbb{R}_{\geq 1}$ to refer to the set of all real numbers that are at least 1. We say a function $f: \mathbb{N} \to \mathbb{R}_{>0}$ is *negligible* if for all $c \in \mathbb{N}$,

$$\lim_{n\to\infty} n^c \cdot f(n) = 0.$$

We say a function $f : \mathbb{N} \to \mathbb{R}$ is *non-negligible* if there exists some $c \in \mathbb{N}$ such that $f(n) \ge 1/n^c$ for all sufficiently large n.

For two distributions \mathcal{D}_1 , \mathcal{D}_2 , we use the notation $\Delta(\mathcal{D}_1, \mathcal{D}_2)$ to denote the total variation distance between \mathcal{D}_1 and \mathcal{D}_2 , which we refer to simply as the *statistical distance* between the two distributions. We say that two distributions are *statistically close* if their statistical distance is negligible (in some implicit parameter, typically n or m for us).

For a set S, we let U(S) denote the uniform distribution over S. (If S is not finite and $S \subseteq \mathbb{R}^n$ is Lebesgue measurable, this will be uniform with respect to the standard Lebesgue measure.)

2.1 Norms & Matrices

We use the notation $\mathbf{0}$ to denote the all 0s vector in \mathbb{R}^n , where the dimension n is clear from context. We use $\mathbf{I}_n \in \mathbb{R}^{n \times n}$ to denote the n-dimensional identity matrix. We use the standard $\ell_1, \ell_2, \ell_\infty$ norms on \mathbb{R}^n . For a matrix $\mathbf{A} \in \mathbb{R}^{n \times m}$, we use the notation $\sigma_{\max}(\mathbf{A})$ to denote the spectral

norm, or maximum singular value, of A. Explicitly,

$$\sigma_{\max}(\mathbf{A}) = \max_{\mathbf{v} \neq \mathbf{0}} \frac{\|\mathbf{A}\mathbf{v}\|_2}{\|\mathbf{v}\|_2}.$$

For a matrix $A \in \mathbb{R}^{n \times m}$, we often write A by its columns, as in $A = [\mathbf{a}_1, \mathbf{a}_2, \dots, \mathbf{a}_m]$ for $\mathbf{a}_i \in \mathbb{R}^n$. We sometimes abuse notation and move interchangeably between matrices $A \in \mathbb{R}^{n \times m}$ and tuples of m vectors in \mathbb{R}^n (as defined by the columns of A). For a matrix $A \in \mathbb{R}^{n \times m}$ given by $A = [\mathbf{a}_1, \mathbf{a}_2, \dots, \mathbf{a}_m]$, we use the notation

$$\|\mathbf{A}\| = \max_{j \in [m]} \|\mathbf{a}_j\|_2.$$

We emphasize that $\|A\|$ does *not* refer to the standard spectral norm on matrices.

Lemma 1. For $A \in \mathbb{R}^{n \times m}$ and $\mathbf{v} \in \mathbb{R}^m$, we have the inequality $\|A\mathbf{v}\|_2 \leq \|A\| \|\mathbf{v}\|_1$.

Proof. Let $A = [a_1, \dots, a_m]$, and let v have entries $v_i \in \mathbb{R}$. We have

$$\|\mathbf{A}\mathbf{v}\|_{2} = \left\| \sum_{j=1}^{m} v_{j} \mathbf{a}_{j} \right\|_{2} \leq \sum_{j=1}^{m} \|v_{j} \mathbf{a}_{j}\|_{2} = \sum_{j=1}^{m} |v_{j}| \|\mathbf{a}_{j}\|_{2} \leq \sum_{j=1}^{m} |v_{j}| \|\mathbf{A}\| = \|\mathbf{A}\| \|\mathbf{v}\|_{1}.$$

We also use the following basic fact.

Lemma 2. For all $\mathbf{v} \in \mathbb{R}^n$, $\|\mathbf{v}\|_1 \le n \|\mathbf{v}\|_{\infty}$.

2.2 Lattices

For an invertible matrix $\mathbf{B} \in \mathbb{R}^{n \times n}$, an n-dimensional lattice generated by basis \mathbf{B} , denoted $\mathcal{L}(\mathbf{B})$, is given by all integer linear combinations of columns of \mathbf{B} . That is,

$$\mathcal{L}(\mathbf{B}) := \{ \mathbf{B}\mathbf{z} : \mathbf{z} \in \mathbb{Z}^n \}.$$

We define the (half-open) parallelepiped $\mathcal{P}(B)$ to be the set

$$\mathcal{P}(\mathbf{B}) := \{ \mathbf{B}\mathbf{v} : \mathbf{v} \in [0, 1)^n \}.$$

Note that for all $x \in \mathbb{R}^n$, there exists unique $y \in \mathcal{P}(B)$ such that $x - y \in \mathcal{L}(B)$. We use the notation $y = x \pmod{B}$ to denote the corresponding $y \in \mathcal{P}(B)$ for a given $x \in \mathbb{R}^n$. Note that y is computable in polynomial time given B and x. For a lattice Λ , we denote the *dual lattice* of Λ as Λ^* , defined by

$$\Lambda^* = \{ \mathbf{x} \in \mathbb{R}^n : \forall \mathbf{v} \in \Lambda, \langle \mathbf{x}, \mathbf{v} \rangle \in \mathbb{Z} \}.$$

For $i \in [n]$, we can define the *i*th successive minimum of a lattice Λ to be the smallest λ_i such that there exist *i* linearly independent lattice points of ℓ_2 norm at most λ_i . Letting *B* denote the unit ball, this can be phrased as

$$\lambda_i(\Lambda) := \min\{r : \dim(\operatorname{span}(\Lambda \cap rB)) \ge i\}.$$

Note that $\lambda_1(\Lambda)$ is the *minimum distance* of Λ .

We also define the *covering radius* $v(\Lambda)$ of a lattice Λ , defined by

$$\nu(\Lambda) = \max_{\mathbf{x} \in \mathbb{R}^n} \min_{\mathbf{v} \in \Lambda} \|\mathbf{x} - \mathbf{v}\|_2.$$

That is, $\nu(\Lambda)$ is the smallest real number such that every element of \mathbb{R}^n has distance at most $\nu(\Lambda)$ from (some point in) Λ .

We now define some fundamental (worst-case) lattice problems.

Definition 1 (Shortest Independent Vectors Problem). For a parameter $\gamma : \mathbb{N} \to \mathbb{R}_{\geq 1}$, the shortest independent vectors problem (SIVP $_{\gamma}$) is a (worst-case) search problem defined as follows. Given an invertible basis $\mathbf{B} \in \mathbb{R}^{n \times n}$ as input, output n vectors $\mathbf{S} = [\mathbf{s}_1, \cdots, \mathbf{s}_n] \in \mathbb{R}^{n \times n}$ such that the following hold:

- S is linearly independent.
- For all $i \in [n]$, $s_i \in \mathcal{L}(B)$;
- $\|\mathbf{S}\| \leq \gamma(n) \cdot \lambda_n(\mathcal{L}(\mathbf{B})).$

Definition 2 (Covering Radius Problem). For a parameter $\gamma : \mathbb{N} \to \mathbb{R}_{\geq 1}$, the gap covering radius problem (GapCRP $_{\gamma}$) is a (worst-case) decision problem defined as follows. Given an invertible basis $\mathbf{B} \in \mathbb{R}^{n \times n}$ and threshold $\theta \in \mathbb{R}_{>0}$ as input, output 1 if $\nu(\mathcal{L}(\mathbf{B})) \leq \theta$ and 0 if $\nu(\mathcal{L}(\mathbf{B})) > \gamma(n) \cdot \theta$.

Definition 3 (Guaranteed Distance Decoding). For a parameter $\gamma : \mathbb{N} \to \mathbb{R}_{\geq 1}$, the guaranteed distance decoding (GDD $_{\gamma}$) is a (worst-case) search problem defined as follows. Given an invertible basis $\mathbf{B} \in \mathbb{R}^{n \times n}$ and target vector $\mathbf{t} \in \mathbb{R}^n$ as input, output a lattice point $\mathbf{x} \in \mathcal{L}(\mathbf{B})$ such that $\|\mathbf{t} - \mathbf{x}\|_2 \leq \gamma(n) \cdot \lambda_n(\mathcal{L}(\mathbf{B}))$.

We note that GDD is often defined using $v(\cdot)$ instead of $\lambda_n(\cdot)$, as using $v(\cdot)$ guarantees solutions for all $\gamma \geq 1$. However, throughout this paper, we will be in the regime where $\gamma(n) \geq \sqrt{n}/2$, which guarantees a solution even with $\lambda_n(\cdot)$ because $\lambda_n(\Lambda) \geq \frac{2}{\sqrt{n}}v(\Lambda)$ for all full-rank lattices Λ [GMR04, Lemma 4.3].

We will also use an intermediate problem called Incremental Guaranteed Distance Decoding (IncGDD) defined by [MR07].

Definition 4 (Incremental GDD, Definition 5.6 of [MR07]). For a parameter $\gamma : \mathbb{N} \to \mathbb{R}_{\geq 1}$, the incremental guaranteed distance decoding (IncGDD $_{\gamma}$) problem is a (worst-case) search problem defined as follows. Given as input:

- An invertible basis $\mathbf{B} \in \mathbb{R}^{n \times n}$,
- A set S of n linearly independent vectors in $\mathcal{L}(B)$ (represented as columns of $S \in \mathbb{R}^{n \times n}$),
- A target point $\mathbf{t} \in \mathbb{R}^n$, and
- A parameter $r > \gamma(n) \cdot \lambda_n(\mathcal{L}(\mathbf{B}))$,

output some vector $\mathbf{s} \in \mathcal{L}(\mathbf{B})$ such that $\|\mathbf{s} - \mathbf{t}\|_2 \le r + \frac{\|\mathbf{S}\|}{8}$.

Comparison to [MR07]. Definition 4 differs from [MR07, Definition 5.6] in two ways. First, instead of using $\lambda_n(\mathcal{L}(\mathbf{B}))$, [MR07] considers general functions ϕ mapping n-dimensional lattices to $\mathbb{R}_{>0}$. Second, instead of fixing the constant 8, [MR07] parametrizes this more generally by some constant g.

As shown in [MR07], there are reductions from these worst-case lattice problems to IncGDD.

Lemma 3 (Lemma 5.10 in [MR07]). For any $\gamma(n) \ge 1$, there is a reduction from SIVP_{8 γ} to IncGDD_{γ}.

Lemma 4 (Combining Lemmas 5.11, 5.12 in [MR07]). For any $\gamma(n) \ge 1$, there is a (randomized) reduction from GapCRP_{12 γ} to IncGDD_{γ}.

Lemma 5 (Lemma 5.11 in [MR07]). For any $\gamma(n) \ge 1$, there is a reduction from $GDD_{3\gamma}$ to $IncGDD_{\gamma}$.

The best known algorithms for these lattice problems employ *lattice reduction* techniques, based on LLL and BKZ [LLL82, Sch87]. They currently all have the following time-approximation trade-off: For a tunable parameter k, in dimension n, it is posible to solve lattice problems with approximation factor $\gamma = 2^{\widetilde{O}(n/k)}$ in time $2^{\widetilde{O}(k)}$, where $\widetilde{O}(\cdot)$ hides polylog factors in n. Equalizing these terms gives a $2^{\widetilde{O}(\sqrt{n})}$ -time algorithm to solve $2^{\widetilde{O}(\sqrt{n})}$ -approximate lattice problems. Despite extensive study in the lattice literature, no better algorithms are known (that improve the exponent by more than a polylog factor). In particular, the following two assumptions stand:

Assumption 1 (Polynomial Hardness of Approximate Worst-case Lattice Problems). For every polynomial $\gamma(n)$, at least one of SIVP $_{\gamma}$, GapCRP $_{\gamma}$, or GDD $_{\gamma}$ requires super-polynomial time to solve.

Assumption 2 (Subexponential Hardness of Approximate Worst-case Lattice Problems). For all constants $\varepsilon > 0$, at least one of SIVP $_{\gamma}$, GapCRP $_{\gamma}$, or GDD $_{\gamma}$ requires time $2^{\omega(n^{1/2-\varepsilon})}$ to solve, where $\gamma(n) = 2^{n^{1/2-\varepsilon}}$.

2.3 Continuous and Discrete Gaussian Measures

For $\mu \in \mathbb{R}$ and $\sigma \in \mathbb{R}_{>0}$, we use the notation $\mathcal{N}(\mu, \sigma^2)$ to denote the (univariate) Normal distribution with mean μ and standard deviation σ . More generally, in $n \in \mathbb{N}$ variables, for $\mu \in \mathbb{R}^n$ and positive semi-definite $\Sigma \in \mathbb{R}^{n \times n}$, we use the notation $\mathcal{N}(\mu, \Sigma)$ to denote the multivariate Gaussian distribution with mean μ and covariance matrix Σ . For the special case where $\Sigma = \sigma^2 \mathbf{I}_n$ for some $\sigma \in \mathbb{R}_{>0}$, let $\varphi_{\sigma,\mu}(\cdot)$ denote the probability density function of $\mathcal{N}(\mu, \sigma^2 \mathbf{I}_n)$, i.e.,

$$\varphi_{\sigma,\boldsymbol{\mu}}(\mathbf{x}) = \frac{1}{(2\pi\sigma^2)^{n/2}} \cdot \exp\left(-\frac{\|\mathbf{x} - \boldsymbol{\mu}\|_2^2}{2\sigma^2}\right).$$

We now recall standard Gaussian measure notions from the lattice literature (specifically, [MR07]). For an input $\mathbf{x} \in \mathbb{R}^n$, we define the Gaussian function $\rho_{s,\mu}(\cdot)$ centered at $\mu \in \mathbb{R}^n$ with scale parameter $s \in \mathbb{R}_{>0}$ to be

$$\rho_{s,\boldsymbol{\mu}}(\mathbf{x}) = \exp\left(-\frac{\pi \|\mathbf{x} - \boldsymbol{\mu}\|_2^2}{s^2}\right) = s^n \cdot \varphi_{s/\sqrt{2\pi},\boldsymbol{\mu}}(\mathbf{x}).$$

Therefore, $\rho_{s,\mu}(\mathbf{x})/s^n$ is the density function for the probability distribution $\mathcal{N}(\boldsymbol{\mu}, \frac{s^2}{2\pi}\mathbf{I}_n)$. For any discrete $S \subseteq \mathbb{R}^n$, let $\rho_{s,\mu}(S)$ denote the sum

$$\sum_{\mathbf{v}\in S}\rho_{s,\mu}(\mathbf{v})\in\mathbb{R}.$$

Let Λ be a full-rank *n*-dimensional lattice. We define the *discrete Gaussian distribution* $D_{\Lambda+\mu,s}$ shifted by $\mu \in \mathbb{R}^n$ and scale parameter $s \in \mathbb{R}_{>0}$ to be the distribution with probability mass function

$$D_{\Lambda+\mu,s}(\mathbf{x}) = \begin{cases} \frac{\rho_{s,0}(\mathbf{x})}{\rho_{s,0}(\Lambda+\mu)} = \frac{\rho_{s,0}(\mathbf{x})}{\rho_{s,-\mu}(\Lambda)} & \text{if } \mathbf{x} \in \Lambda, \\ 0 & \text{otherwise,} \end{cases}$$
$$= \mathbf{1}[\mathbf{x} \in \Lambda] \cdot \frac{\exp\left(-\pi \|\mathbf{x} - \mu\|_{2}^{2}/s^{2}\right)}{\sum_{\mathbf{y} \in \Lambda} \exp\left(-\pi \|\mathbf{x} - \mu\|_{2}^{2}/s^{2}\right)}.$$

We will use the fact that there is an efficient sampler for the discrete Gaussian for the special case where $\Lambda = \mathbb{Z}^n$ (see, e.g., [BLP+13, Lemma 2.3]).

For an *n*-dimensional lattice Λ and $\epsilon > 0$, we define the *smoothing parameter* $\eta_{\epsilon}(\Lambda)$ to be the smallest $s \in \mathbb{R}_{>0}$ such that $\rho_{1/s,0}(\Lambda^* \setminus \{\mathbf{0}\}) \le \epsilon$. We now recall standard results about the smoothing parameter.

Lemma 6 (Lemma 3.3 of [MR07]). For any n-dimensional lattice Λ and $\epsilon > 0$,

$$\eta_{\epsilon}(\Lambda) \leq \sqrt{\frac{\ln(2n(1+1/\epsilon))}{\pi}} \cdot \lambda_n(\Lambda).$$

Lemma 7 (Lemma 4.1 of [MR07]). For any lattice $\mathcal{L}(\mathbf{B})$ and any $\epsilon > 0$, $s \geq \eta_{\epsilon}(\mathcal{L}(\mathbf{B}))$, and $\mu \in \mathbb{R}^{n}$, we have the statistical distance bound

$$\Delta\left(\mathcal{N}\left(\boldsymbol{\mu}, \frac{s^2}{2\pi}\mathbf{I}_n\right) \mod \mathcal{P}(\mathbf{B}), U(\mathcal{P}(\mathbf{B}))\right) \leq \frac{\epsilon}{2}.$$

Corollary 1. For any full-rank $\mathbf{B} \in \mathbb{R}^{n \times n}$, $\epsilon > 0$, and $\boldsymbol{\mu} \in \mathbb{R}^n$, if

$$\sigma \geq \sqrt{\frac{\ln(2n(1+1/\epsilon))}{2\pi^2}} \cdot \lambda_n(\mathcal{L}(\mathbf{B})),$$

then we have the statistical distance bound

$$\Delta\left(\mathcal{N}\left(\boldsymbol{\mu},\sigma^2\mathbf{I}_n\right) \mod \mathcal{P}(\mathbf{B}), U(\mathcal{P}(\mathbf{B}))\right) \leq \frac{\epsilon}{2}.$$

Proof. This follows by combining Lemmas 6 and 7, where we set $\sigma = s/\sqrt{2\pi}$.

Lemma 8. For $\sigma > 0$, let \mathcal{D}_{σ} denote the distribution of outputs when first sampling $\mathbf{v} \sim U([0,1)^n)$ and then outputting a sample from $D_{\mathbb{Z}^n+\mathbf{v},\sigma\sqrt{2\pi}}$. For any $\epsilon \in (0,1/2)$, if

$$\sigma \geq \sqrt{\frac{\ln(2n(1+1/\epsilon))}{2\pi^2}},$$

then we have the statistical distance bound

$$\Delta\left(\mathcal{N}\left(\mathbf{0},\sigma^{2}\mathbf{I}_{n}\right),\mathcal{D}_{\sigma}\right)\leq4\epsilon.$$

Proof sketch of Lemma 8. Let $s = \sigma \sqrt{2\pi}$. The analysis in [GVV22, Lemma 17] shows that the distributions have statistical distance at most

$$\sup_{\mathbf{v}\in[0,1)^n}\frac{\rho_{s,0}(\mathbb{Z}^n)}{\rho_{s,0}(\mathbb{Z}^n+\mathbf{v})}-1=\sup_{\mathbf{v}\in[0,1)^n}\frac{\rho_{s,0}(\mathbb{Z}^n)}{\rho_{s,-\mathbf{v}}(\mathbb{Z}^n)}-1.$$

Implicit in [MR07, Lemma 4.4] is that as long as $s \ge \eta_{\epsilon}(\mathbb{Z}^n)$, then

$$\frac{\rho_{s,0}(\mathbb{Z}^n)}{\rho_{s,-\mathbf{v}}(\mathbb{Z}^n)} \in \left[1, \frac{1+\epsilon}{1-\epsilon}\right] \subseteq [1, 1+4\epsilon],$$

where the last inclusion comes from the bound $\epsilon < 1/2$. Subtracting by 1 and invoking Lemma 6 with $\Lambda = \mathbb{Z}^n$ gives the desired bound.

We now recall some standard spectral bounds for Gaussian matrices.

Lemma 9 (As in [RV10]). Let $A \in \mathbb{R}^{n \times m}$ be such that $A_{i,j} \sim_{i.i.d.} \mathcal{N}(0,1)$. For all t > 0, we have

$$\Pr\left[\sigma_{\max}(\mathbf{A}) \leq \sqrt{n} + \sqrt{m} + t\right] \geq 1 - 2e^{-t^2/2}.$$

In particular, for $m \ge 16n$, by setting $t = \sqrt{n}$, we have

$$\Pr\left[\sigma_{\max}(\mathbf{A}) \leq \frac{3\sqrt{m}}{2}\right] \geq 1 - 2e^{-n/2}.$$

We now recall standard tail bounds for the $\chi^2(n)$ distribution, corresponding to ℓ_2 -norm bounds on Gaussian vectors.

Lemma 10 (As in [LM00], Corollary of Lemma 1). For any $t \ge 0$, we have

$$\Pr_{\mathbf{v} \sim \mathcal{N}(\mathbf{0}, \mathbf{I}_n)} \left[\|\mathbf{v}\|_2^2 \ge n + 2\sqrt{tn} + 2t \right] \le e^{-t}.$$

In particular, setting t = n/4, we get

$$\Pr_{\mathbf{v} \sim \mathcal{N}(\mathbf{0}, \mathbf{I}_n)} \left[\|\mathbf{v}\|_2 \ge \sqrt{5/2} \cdot \sqrt{n} \right] \le e^{-n/4}.$$

2.4 Symmetric Perceptrons and Number Partitioning

Definition 5 (Symmetric Binary Perceptron Problem). For a parameter $\kappa: \mathbb{R}_{\geq 1} \to [0, 1]$, the symmetric binary perceptron (SBP $_{\kappa}$) problem is an average-case search problem defined as follows. Given a random Gaussian matrix $\mathbf{A} \sim \mathcal{N}(0, 1)^{n \times m}$ as input where $m \geq n$, output a vector $\mathbf{x} \in \{-1, 1\}^m$ such that $\|\mathbf{A}\mathbf{x}\|_{\infty} \leq \kappa(m/n) \cdot \sqrt{m}$.

Definition 6 (Number Partitioning Problem). For a parameter $\kappa : \mathbb{N} \to [0,1]$, the number partitioning problem (NPP_{\kappa}) is an average-case search problem defined as follows. Given a random Gaussian vector $\mathbf{a} \sim \mathcal{N}(\mathbf{0}, \mathbf{I}_m)$, output a vector $\mathbf{x} \in \{-1, 1\}^m$ such that $|\mathbf{a}^\top \mathbf{x}| \leq \kappa(m) \cdot \sqrt{m}$.

We emphasize that NPP_{κ} is exactly a special case of SBP_{κ} (when setting n=1).

2.5 Worst-case to Average-case Reductions

We recall basic notions of reductions between (search) worst-case and average-case computational problems. Let $A \in \mathsf{FNP}$ be a worst-case search problem, and let $B \in \mathsf{FNP}$ be an average-case search problem defined over some distribution family $\{\mathcal{D}_n\}_{n\in\mathbb{N}}$. We say that there is a T(n)-time reduction from A to B if for all non-negligible functions μ , there exists a randomized poly(T)-time oracle Turing machine $M^{(\cdot)}$ such that for all (possibly randomized) \mathcal{O} such that

$$\Pr_{y \leftarrow D_n}[(y, \mathcal{O}(y)) \in B] \ge \mu(n),$$

for all $n \in \mathbb{N}$, it holds that for all $x \in \{0, 1\}^*$,

$$\Pr_{r}\left[\left(x, M^{\mathcal{O}}(x; r)\right) \in A\right] \ge \frac{2}{3},$$

where r denotes the internal randomness of M. We emphasize that the polynomial in the poly(T) run-time of $M^{(\cdot)}$ may depend on the non-negligible function μ . Since x is worst-case and $A \in \mathsf{FNP}$, standard amplification applies to make the success probability of M exponentially close to 1.

3 Reduction to Symmetric Binary Perceptrons

Theorem 3. Suppose $\kappa(x) = 1/x^{1/2+\varepsilon}$ for some constant $\varepsilon > 0$. There exists some polynomial $\gamma(n) = n^{O(1/\varepsilon)}$ such that there is a polynomial-time reduction from $\operatorname{IncGDD}_{\gamma}$ to $\operatorname{SBP}_{\kappa}$.

We now state our main corollary for symmetric binary perceptrons.

Corollary 2. Suppose there is a polynomial time algorithm for SBP_{κ} (on average) for $\kappa(x) = 1/x^{1/2+\epsilon}$ for some constant $\epsilon > 0$ that succeeds with non-negligible probability. Then, there are randomized polynomial time algorithms for the (worst-case) lattice problems SIVP_{γ}, GapCRP_{γ}, and GDD_{γ} for some polynomial $\gamma(n)$. In particular, Assumption 1 implies that there is no polynomial time algorithm for SBP_{κ} (on average) for $\kappa(x) = 1/x^{1/2+\epsilon}$ for some constant $\epsilon > 0$ that succeeds with non-negligible probability.

More generally, if there is a T(n,m)-time algorithm for SBP_{κ} (on average) for $\kappa(x)=1/x^{1/2+\epsilon}$ for some constant $\epsilon>0$ that succeeds with non-negligible probability, then there are randomized poly $(n,T(n,\operatorname{poly}(n)))$ -time algorithms for the (worst-case) lattice problems $SIVP_{\gamma}$, $GapCRP_{\gamma}$, and GDD_{γ} for some polynomial $\gamma(n)$.

Proof of Corollary 2. This follows by directly composing Lemmas 3 to 5 and Theorem 3. \Box

Remark 1. For a sharper bound on κ in Theorem 3 and Corollary 2, we can instead assume subexponential hardness of approximate worst-case lattice problems. Specifically, for $\kappa(x) = \frac{1}{\sqrt{x} \log^{1+c}(x)}$ where c > 0, we can set $m = 2^{O\left(n^{3/(2c)}\right)}$ and $\gamma(n) = 2^{O\left(n^{3/(2c)}\right)}$ in the reduction from IncGDD. In particular, for c > 3, Assumption 2 implies that there is no polynomial time algorithm for SBP $_{\kappa}$ (on average) for $\kappa(x) = \frac{1}{\sqrt{x} \log^{1+c}(x)}$ that succeeds with non-negligible probability.

3.1 Proof of Theorem 3

We now prove Theorem 3.

Proof of Theorem 3. Let $(\mathbf{B} \in \mathbb{R}^{n \times n}, \mathbf{S} \in \mathbb{R}^{n \times n}, \mathbf{t} \in \mathbb{R}^n, r \in \mathbb{R})$ be the given IncGDD instance. Let

$$\sigma_2 = \ln n,$$

$$m = \left[\left(8\sigma_2 n^{3/2 + \varepsilon} \right)^{1/\varepsilon} \right] = n^{1 + \Theta(1)},$$

$$\sigma_1 = \frac{r}{4m},$$

$$\gamma(n) = 4m \ln n = n^{1 + \Theta(1)}.$$

Sample $\mathbf{u}_1 \sim \mathcal{N}(\mathbf{t}, \sigma_1^2 \mathbf{I}_n)$, $\mathbf{u}_2, \cdots, \mathbf{u}_m \sim_{\text{i.i.d.}} \mathcal{N}(\mathbf{0}, \sigma_1^2 \mathbf{I}_n)$. Let $\mathbf{U} = [\mathbf{u}_1, \mathbf{u}_2, \cdots, \mathbf{u}_m] \in \mathbb{R}^{n \times m}$. Sample m uniformly random lattice vectors $\mathbf{v}_i \in \mathcal{L}(\mathbf{B}) \mod \mathcal{P}(\mathbf{S})$ (see [Mic04, Proposition 2.9]), and let $\mathbf{V} = [\mathbf{v}_1, \mathbf{v}_2, \cdots, \mathbf{v}_m] \in \mathbb{R}^{n \times m}$. Define

$$\widetilde{\mathbf{A}} = \mathbf{S}^{-1}(\mathbf{V} + \mathbf{U}) \mod \mathbb{Z}^n \in [0, 1)^{n \times m}$$

Proposition 1. The distribution of $\widetilde{\mathbf{A}}$ is statistically close to $U([0,1)^{n\times m})$.

Proof of Proposition 1. Recall that by definition of IncGDD, we know

$$r > \gamma(n) \cdot \lambda_n(\mathcal{L}(\mathbf{B})) = 4m \ln n \cdot \lambda_n(\mathcal{L}(\mathbf{B})).$$

In anticipation of applying Corollary 1, we observe that for sufficiently large n,

$$\sigma_1 = \frac{r}{4m} \ge \ln n \cdot \lambda_n(\mathcal{L}(\mathbf{B})) \ge \sqrt{\frac{\ln \left(2n \left(1 + 1/e^{-\ln^2(n)}\right)\right)}{2\pi^2}} \cdot \lambda_n(\mathcal{L}(\mathbf{B})).$$

Therefore, we can invoke Corollary 1 m times (once with $\mu = t$, m-1 times with $\mu = 0$) and the triangle inequality to see that

$$\Delta(\mathbf{U} \mod \mathcal{P}(\mathbf{B}), U(\mathcal{P}(\mathbf{B}))^m) \le m \cdot \frac{e^{-\ln^2(n)}}{2} = \mathsf{negl}(n).$$

Since V has columns that are uniform elements of $\mathcal{L}(B) \mod \mathcal{P}(S)$, and since $\mathcal{L}(S) \subseteq \mathcal{L}(B)$, it follows that

$$\Delta(\mathbf{U} + \mathbf{V} \mod \mathcal{P}(\mathbf{S}), U(\mathcal{P}(\mathbf{S}))^m) \leq \operatorname{negl}(n).$$

Multiplying on the left by S⁻¹ gives

$$\Delta\left(\widetilde{\mathbf{A}}, U(\mathcal{P}(\mathbb{Z}^n))^m\right) \leq \mathsf{negl}(n),$$

or equivalently, that $\widetilde{\mathbf{A}}$ is statistically close to uniform over $[0,1)^{n\times m}$.

For $j \in [m]$, let $\widetilde{\mathbf{a}}_j \in [0,1)^n$ denote the jth column of $\widetilde{\mathbf{A}}$. For all $j \in [m]$, sample $\mathbf{w}_j \sim D_{\mathbb{Z}^n + \widetilde{\mathbf{a}}_j, \sigma_2 \sqrt{2\pi}}$. Let $\mathbf{W} = [\mathbf{w}_1, \cdots, \mathbf{w}_m]$. Note that by construction, $\mathbf{w}_j \equiv \widetilde{\mathbf{a}}_j \mod \mathbb{Z}^n$, i.e., $\mathbf{W} \equiv \widetilde{\mathbf{A}} \mod \mathbb{Z}^n$. In anticipation of applying Lemma 8, we observe that for sufficiently large n,

$$\sigma_2 = \ln(n) \ge \sqrt{\frac{\ln\left(2n\left(1+1/e^{-\ln^2(n)}\right)\right)}{2\pi^2}}.$$

Therefore, by invoking Lemma 8 m times, the triangle inequality, and Proposition 1, we see that

$$\Delta\left(\mathbf{W}, \mathcal{N}\left(0, \sigma_2^2\right)^{n \times m}\right) \leq 4me^{-\ln^2(n)} + \mathsf{negl}(n) = \mathsf{negl}(n).$$

Let $A = \frac{1}{\sigma_2}W$. It follows that A is statistically close to $\mathcal{N}(0,1)^{n\times m}$.

Feed $\tilde{\mathbf{A}}$ into the SBP_{κ} solver to receive some $\mathbf{x} \in \{-1, 1\}^m$ such that $\|\mathbf{A}\mathbf{x}\|_{\infty} \le \kappa(m/n) \cdot \sqrt{m}$. Let $\mathbf{e} = -\mathbf{A}\mathbf{x} \in \mathbb{R}^n$ with $\|\mathbf{e}\|_{\infty} \le \kappa(m/n) \cdot \sqrt{m}$. Let $x_1 \in \{-1, 1\}$ be the first entry of \mathbf{x} . Let $\mathbf{e}' = \sigma_2 \mathbf{e} \in \mathbb{R}^n$. The reduction outputs $\mathbf{s} = x_1(\mathbf{U}\mathbf{x} + \mathbf{S}\mathbf{e}') \in \mathbb{R}^n$.

We first argue that $s \in \mathcal{L}(B)$. Since Ax + e = 0, by scaling up by σ_2 , we have Wx + e' = 0. Since $x \in \mathbb{Z}^m$,

$$0 = Wx + e' \equiv \widetilde{A}x + e' \equiv S^{-1}(V + U)x + e' \mod \mathbb{Z}^n$$

meaning that $S^{-1}(V + U)x + e' \in \mathbb{Z}^n$, and thus

$$(V + U)x + Se' \in \mathcal{L}(S) \subseteq \mathcal{L}(B)$$
.

Since V contains vectors in $\mathcal{L}(\mathbf{B})$ and $\mathbf{x} \in \mathbb{Z}^m$, we can subtract by $\mathbf{V}\mathbf{x}$ to get

$$Ux + Se' \in \mathcal{L}(B)$$
.

Multiplying by the sign $x_1 \in \{-1, 1\}$ gives

$$s = x_1(Ux + Se') \in \mathcal{L}(B),$$

as desired.

We next argue that the norm of $\mathbf{s} - \mathbf{t}$ is small. Decompose \mathbf{x} as $\mathbf{x}^{\top} = [x_1 || \mathbf{x}_{-1}^{\top}]$, and decompose \mathbf{U} as $\mathbf{U} = [\mathbf{u}_1 || \mathbf{U}_{-1}]$. We then have

$$s = x_1(Ux + Se') = x_1(u_1x_1 + U_{-1}x_{-1} + Se')$$

$$= x_1^2u_1 + x_1U_{-1}x_{-1} + x_1Se'$$

$$= u_1 + x_1U_{-1}x_{-1} + x_1Se'$$

$$= t + u'_1 + x_1U_{-1}x_{-1} + x_1Se',$$

where the distribution of \mathbf{u}_1' is $\mathcal{N}(\mathbf{0}, \sigma_1^2 \mathbf{I}_n)$. It follows that with all but negligible probability,

$$\|\mathbf{s} - \mathbf{t}\|_{2} = \|\mathbf{u}'_{1} + x_{1}\mathbf{U}_{-1}\mathbf{x}_{-1} + x_{1}\mathbf{Se'}\|_{2}$$

$$\leq \|\mathbf{u}'_{1}\|_{2} + \|\mathbf{U}_{-1}\mathbf{x}_{-1}\|_{2} + \|\mathbf{Se'}\|_{2}$$

$$\leq \sigma_{1}\sqrt{\frac{5n}{2}} + \sigma_{\max}(\mathbf{U}_{-1})\|\mathbf{x}_{-1}\|_{2} + \|\mathbf{Se'}\|_{2} \qquad \text{(by Lemma 10)}$$

$$\leq \sigma_{1}\sqrt{\frac{5n}{2}} + \sigma_{\max}(\mathbf{U}_{-1})\sqrt{m-1} + \|\mathbf{S}\|\|\mathbf{e'}\|_{1} \qquad \text{(by Lemma 1)}$$

$$\leq \sigma_{1}\sqrt{\frac{5n}{2}} + \sigma_{\max}(\mathbf{U}_{-1})\sqrt{m-1} + n\|\mathbf{S}\|\|\mathbf{e'}\|_{\infty} \qquad \text{(by Lemma 2)}$$

$$\leq \sigma_{1}\sqrt{\frac{5n}{2}} + \frac{3\sigma_{1}\sqrt{m}}{2} \cdot \sqrt{m-1} + n\|\mathbf{S}\|\|\mathbf{e'}\|_{\infty} \qquad \text{(by Lemma 9)}$$

$$\leq 4\sigma_{1}m + n\|\mathbf{S}\|\|\mathbf{e'}\|_{\infty}$$

$$\leq r + n\|\mathbf{S}\|\|\mathbf{e'}\|_{\infty}.$$

Therefore, it suffices to show that $\|\mathbf{e}'\|_{\infty} \leq 1/(8n)$. Recall that we have

$$\|\mathbf{e}'\|_{\infty} = \sigma_2 \|\mathbf{e}\|_{\infty} \le \sigma_2 \cdot \kappa(m/n) \cdot \sqrt{m} = \sigma_2 \cdot \left(\frac{n}{m}\right)^{1/2+\varepsilon} \sqrt{m} = \sigma_2 \cdot \frac{n^{1/2+\varepsilon}}{m^{\varepsilon}}.$$

Since $m \ge (8\sigma_2 n^{3/2+\varepsilon})^{1/\varepsilon}$, we have $\|\mathbf{e}\|_{\infty} \le 1/(8n)$, as desired.

Lastly, we note that the SBP_{κ} solver need only succeed with some non-negligible probability μ . As a result, we can repeat this whole process $O(1/\mu) = \text{poly}(n, m)$ times, and since we can efficiently verify whether the SBP_{κ} solver succeeded, the reduction will still go through.

3.2 Variants and Generalizations

We mention a few variants of SBP for which the reduction in Theorem 3 would also apply. As they are not critical to our main result, for simplicity, we only sketch the justifications.

- 1. **Uniform A**. Instead of having $\mathbf{A} \sim \mathcal{N}(0,1)^{n \times m}$, if we had a SBP solver that worked with $\mathbf{A} \sim U([0,1])^{n \times m}$, our reduction would actually be simpler and more direct. In particular, there would be no need for discrete Gaussian sampling.
- 2. **Zero entries in x**. Instead of requiring $\mathbf{x} \in \{\pm 1\}^m$ from the SBP solver, if we instead allowed $\mathbf{x} \in \{-1, 0, 1\}^m$ with $\mathbf{x} \neq \mathbf{0}$ from the SBP solver, a similar reduction to the one in Theorem 3 would work as well. The main difference is that the reduction would instead "guess" a coordinate $j \in [m]$ for which $x_j \neq 0$ (with success probability at least 1/m) and put the vector \mathbf{t} in the mean of that coordinate, instead of the first coordinate. (See [MR07] for more rigorous details.)

A priori, the version of SBP that allows zero-entries in **x** is in fact a lot easier. In particular, for $\kappa(x) = 1/\sqrt{x}$, setting $\mathbf{x} = (1, 0, ..., 0)^{\top}$ would get $\|\mathbf{A}\mathbf{x}\|_{\infty} \leq \widetilde{O}(1) \ll \kappa(m/n)\sqrt{m} = \sqrt{n}$

with high probability by standard Gaussian tail bounds. However, in our reduction, we have $\kappa(x) = 1/x^{1/2+\varepsilon}$. The bound on $\|\mathbf{A}\mathbf{x}\|_{\infty}$ is

$$\|\mathbf{A}\mathbf{x}\|_{\infty} \leq \kappa(m/n)\sqrt{m} = \frac{\sqrt{m}}{(m/n)^{1/2+\varepsilon}} = \frac{n^{1/2+\varepsilon}}{m^{\varepsilon}}.$$

In our reduction, we set m so that $m^{\varepsilon} \gg n^{1/2+\varepsilon}$, making $\|\mathbf{A}\mathbf{x}\|_{\infty} \ll 1$, in particular, a stronger requirement than $\|\mathbf{A}\mathbf{x}\|_{\infty} \leq \widetilde{O}(1)$.

3. **Larger** $\mathbf{x} \in \mathbb{Z}^m$. Instead of (in particular) requiring $\|\mathbf{x}\|_{\infty} \leq 1$ from the SBP solver, one could relax this requirement to $\|\mathbf{x}\|_{\infty} \leq B$ for some larger bound $B \in \mathbb{N}$. In addition to the modifications discussed in Item 2, we would put the vector \mathbf{t}/z in the mean of that coordinate, where $z \sim U(\{-B, -B+1, \dots, B-1, B\} \setminus \{0\})$. The runtime, γ , and κ would now worsen by a factor of $\Theta(B)$. (See [MR07] for more rigorous details.)

4 Reduction to Number Partitioning

Lemma 11 (Chinese Remainder Theorem). Let $p_1, ..., p_n \in \mathbb{N}$ be distinct positive prime numbers. For $q = \prod_{i \in [n]} p_i$, there is a group isomorphism

$$\varphi: \bigoplus_{i\in[n]} \mathbb{Z}/p_i\mathbb{Z} \longrightarrow \mathbb{Z}/q\mathbb{Z}.$$

Moreover, this map can be written as

$$\varphi: (y_1, \cdots, y_n) \mapsto \sum_{i \in [n]} c_i y_i$$

for $c_i \in \mathbb{Z}$ such that q/p_i divides c_i for all $i \in [n]$ (so that this map is well-defined). Furthermore, the inverse map $\varphi^{-1} : \mathbb{Z}/q\mathbb{Z} \to \bigoplus_{i \in [n]} \mathbb{Z}/p_i\mathbb{Z}$ can be written as

$$\varphi^{-1}: z \mapsto (z, z, \dots, z),$$

where for all $i \in [n]$, $z \in \mathbb{Z}/q\mathbb{Z} = \{0, ..., q-1\}$ is interpreted directly as an element of $\mathbb{Z}/p_i\mathbb{Z} = \{0, ..., p_i - 1\}$ by reduction modulo p_i .

Lemma 12 (Normalized CRT). Let $p_1, ..., p_n \in \mathbb{N}$ be distinct positive prime numbers. For $q = \prod_{i \in [n]} p_i$, there is a group isomorphism

$$\widetilde{\varphi}: \bigoplus_{i \in [n]} 1/p_i \cdot \mathbb{Z}/p_i \mathbb{Z} \longrightarrow 1/q \cdot \mathbb{Z}/q \mathbb{Z}.$$

Moreover, there exists an integer vector $\mathbf{c} \in \mathbb{Z}^n$ such that this map can be written as

$$\widetilde{\varphi}: (y_1, \cdots, y_n) \mapsto \mathbf{c}^{\mathsf{T}} \mathbf{y} = \sum_{i \in [n]} c_i y_i.$$

Furthermore, the inverse map $\widetilde{\varphi}^{-1}: 1/q \cdot \mathbb{Z}/q\mathbb{Z} \to \bigoplus_{i \in [n]} 1/p_i \cdot \mathbb{Z}/p_i\mathbb{Z}$ can be written as

$$\widetilde{\varphi}^{-1}: z \mapsto \left(\frac{q}{p_1}z, \frac{q}{p_2}z, \dots, \frac{q}{p_n}z\right).$$

Proof of Lemma 12. This follows directly by Lemma 11, by setting

$$\widetilde{\varphi}(y_1,\ldots,y_n)=1/q\cdot\varphi(p_1y_1,p_2y_2,\ldots,p_ny_n).$$

Letting $\mathbf{p} = (p_1, \dots, p_n) \in \mathbb{Z}^n$, we use the notation $[\cdot]_{\mathbf{p}} : [0, 1)^n \to \bigoplus_{i \in [n]} 1/p_i \cdot \mathbb{Z}/p_i \mathbb{Z}$ to denote the function

$$\left[\cdot\right]_{\mathbf{p}}:\mathbf{v}\mapsto\left(\frac{\left[v_{1}p_{1}\right]}{p_{1}},\ldots,\frac{\left[v_{1}p_{n}\right]}{p_{n}}\right).$$

More generally, for elements in $[0,1)^{n\times m}$, we extend $[\cdot]_p:[0,1)^{n\times m}\to \left(\bigoplus_{i\in[n]}1/p_i\cdot\mathbb{Z}/p_i\mathbb{Z}\right)^m$ to operate column-wise.

We will use the following basic fact.

Lemma 13. For any $A \in [0,1)^{n \times m}$ and $\mathbf{x} \in \mathbb{R}^m$,

$$\left\| \left(\mathbf{A} - \left\lfloor \mathbf{A} \right\rfloor_{\mathbf{p}} \right) \mathbf{x} \right\|_{1} \leq \frac{n}{\min_{i \in [n]} p_{i}} \left\| \mathbf{x} \right\|_{1}.$$

Proof. Let $p^* = \min_{i \in [n]} p_i$. Letting $\mathbf{M} = \mathbf{A} - [\mathbf{A}]_p$, by properties of the floor function, we have $\mathbf{M} \in \left[0, \frac{1}{p^*}\right)^{n \times m}$. Let $\mathbf{m}_1, \dots, \mathbf{m}_n \in \left[0, \frac{1}{p^*}\right)^m$ be the rows of \mathbf{M} . We have

$$\left\|\mathbf{M}\mathbf{x}\right\|_{1} = \sum_{i \in [n]} \left|\mathbf{m}_{i}^{\top}\mathbf{x}\right| \leq \sum_{i \in [n]} \left\|\mathbf{m}_{i}\right\|_{\infty} \left\|\mathbf{x}\right\|_{1} \leq \frac{n}{p^{*}} \left\|\mathbf{x}\right\|_{1},$$

as desired, where we have used Hölder's inequality to see that $|\mathbf{m}_i^{\mathsf{T}}\mathbf{x}| \leq ||\mathbf{m}_i||_{\infty} ||\mathbf{x}||_1$.

We also use the following fact about the density of prime numbers.

Lemma 14. For all sufficiently large $N \in \mathbb{N}$, there exist at least $N/\ln(N)$ distinct prime numbers in the interval [N, 10N].

Proof. For $N \in \mathbb{N}$, let $\pi(N) = |\{a \in [N] : a \text{ is prime}\}|$ denote the prime-counting function. By the prime number theorem, we know that for all sufficiently large N,

$$\frac{N}{2\ln N} \le \pi(N) \le \frac{2N}{\ln N}.$$

In particular,

$$\pi(N) \le \frac{2N}{\ln N}, \quad \pi(10N) \ge \frac{10N}{2\ln(10N)} = \frac{10N}{2\ln(N) + 2\ln(10)} > \frac{4N}{\ln N},$$

for sufficiently large N. Therefore, for sufficiently large N, by taking the difference of the two quantities, there are at least $N/\ln N$ primes in [N, 10N].

Theorem 4. Suppose $\kappa(m) = 2^{-\log^{2+\varepsilon} m}$ for some constant $\varepsilon > 0$. Then there exists $\gamma(n) = 2^{O\left(n^{\frac{1+\varepsilon}{1+\varepsilon}}\right)}$ such that there is a poly(m)-time reduction from IncGDD $_{\gamma}$ in dimension $n = \Omega((\log m)^{1+\varepsilon})$ to NPP $_{\kappa}$ (in dimension m).

We now state our main corollary for number partitioning.

Corollary 3. Suppose there is a polynomial time algorithm for NPP_{κ} (on average) for $\kappa(m)=2^{-\log^{2+\varepsilon} m}$ for some constant $\varepsilon>0$ that succeeds with non-negligible probability. Then, there are randomized $2^{O\left(n^{\frac{1}{1+\varepsilon}}\right)}$ -time algorithms for the (worst-case) lattice problems SIVP_{γ}, GapCRP_{γ}, and GDD_{γ} in dimension n for $\gamma(n)=2^{O\left(n^{\frac{1}{1+\varepsilon}}\right)}$. In particular, Assumption 2 implies that for all constant $\varepsilon>0$, there is no polynomial time algorithm for NPP_{κ} (on average) for $\kappa(m)=2^{-\log^{3+\varepsilon} m}$ that succeeds with non-negligible probability.

More generally, suppose there is a T(m)-time algorithm for NPP $_{\kappa}$ (on average) for $\kappa(m)=2^{-\log^{2+\varepsilon}m}$ for some constant $\varepsilon>0$ that succeeds with non-negligible probability. Then, there are randomized $T\left(2^{O\left(n^{\frac{1}{1+\varepsilon}}\right)}\right)$ -time algorithms for the (worst-case) lattice problems SIVP $_{\gamma}$, GapCRP $_{\gamma}$, and GDD $_{\gamma}$ in dimension n for $\gamma(n)=2^{O\left(n^{\frac{1}{1+\varepsilon}}\right)}$.

Proof of Corollary 3. This follows by directly composing Lemmas 3 to 5 and Theorem 4. □

We now prove Theorem 4.

Proof of Theorem 4. Let $(\mathbf{B} \in \mathbb{R}^{n \times n}, \mathbf{S} \in \mathbb{R}^{n \times n}, \mathbf{t} \in \mathbb{R}^n, r \in \mathbb{R})$ be the given IncGDD instance. Let

$$m = 2^{10n\frac{1}{1+arepsilon}},$$
 $\sigma_1 = \frac{r}{4m},$
 $\sigma_2 = \ln m,$
 $\gamma(n) = 4m \ln m = 40 \ln(2) 2^{10n\frac{1}{1+arepsilon}} n^{\frac{1}{1+arepsilon}}$

Let p_1, \ldots, p_n be n distinct prime numbers in the range [32nm, 320nm], which we know must exist for sufficiently large n, m by Lemma 14 (since $m \ge n$). Let $q = \prod_{i=1}^n p_i \le (320nm)^n$. Sample $\mathbf{u}_1 \sim \mathcal{N}(\mathbf{t}, \sigma_1^2 \mathbf{I}_n)$, $\mathbf{u}_2, \cdots, \mathbf{u}_m \sim_{\text{i.i.d.}} \mathcal{N}(\mathbf{0}, \sigma_1^2 \mathbf{I}_n)$. Let $\mathbf{U} = [\mathbf{u}_1, \mathbf{u}_2, \cdots, \mathbf{u}_m] \in \mathbb{R}^{n \times m}$. Sample m uniformly random lattice vectors $\mathbf{v}_i \in \mathcal{L}(\mathbf{B}) \mod \mathcal{P}(\mathbf{S})$ (see [Mic04, Proposition 2.9]), and let $\mathbf{V} = [\mathbf{v}_1, \mathbf{v}_2, \cdots, \mathbf{v}_m] \in \mathbb{R}^{n \times m}$. Define

$$A = S^{-1}(V+U) \mod \mathbb{Z}^n \in [0,1)^{n\times m}.$$

Proposition 2. We have the inequality

$$\Delta(\mathbf{A}, U([0,1)^{n\times m})) \leq \operatorname{negl}(m).$$

Proof. This follows from the same analysis as in Proposition 1. Specifically, since for sufficiently large n and m,

$$\sigma_1 = \frac{r}{4m} \ge \ln m \cdot \lambda_n(\mathcal{L}(\mathbf{B})) \ge \sqrt{\frac{\ln \left(2n \left(1 + 1/e^{-\ln^2(m)}\right)\right)}{2\pi^2}} \cdot \lambda_n(\mathcal{L}(\mathbf{B})),$$

by Corollary 1 and the triangle inequality, we get a total statistical distance of $m \cdot e^{-\ln^2(m)} = \text{negl}(m)$.

Let $\widetilde{\varphi}$ be the isomorphism guaranteed by Lemma 12, with $\mathbf{c} \in \mathbb{Z}^n$ being the coefficients defining the linear map $\widetilde{\varphi}$. Let $\mathbf{y} \in \mathbb{R}^m$ be defined by

$$y = \widetilde{\varphi}([A]_p) + f \mod \mathbb{Z}^m \in [0, 1)^m$$

where $\mathbf{f} \in \mathbb{R}^m$ is sampled as $\mathbf{f} \sim U([0, 1/q)^m)$.

We argue that $\Delta(\mathbf{y}, U([0,1)^m)) \leq \mathsf{negl}(m)$ as follows. Since \mathbf{A} is (close to) $U([0,1)^{n \times m})$, it follows that $[\mathbf{A}]_p$ is (close to) $U(\bigoplus_{i \in [n]} 1/p_i \cdot \mathbb{Z}/p_i\mathbb{Z})^m$. Moreover, since $\widetilde{\varphi}$ is a bijection, it follows that $\widetilde{\varphi}([\mathbf{A}]_p)$ is (close to) $U(1/q \cdot \mathbb{Z}/q\mathbb{Z})^m$. Since $\mathbf{f} \sim U([0,1/q)^m)$, we can see that $\Delta(\mathbf{y}, U([0,1)^m)) \leq \mathsf{negl}(m)$.

Sample $\mathbf{w} \sim D_{\mathbb{Z}^m + \mathbf{y}, \sigma_2 \sqrt{2\pi}}$. Note that by construction, $\mathbf{w} \equiv \mathbf{y} \mod \mathbb{Z}^m$. In anticipation of applying Lemma 8, we observe that for sufficiently large m,

$$\sigma_2 = \ln m \ge \sqrt{\frac{\ln\left(2m\left(1 + 1/e^{-\ln^2(m)}\right)\right)}{2\pi^2}}.$$

Therefore, by invoking Lemma 8, $\Delta(y, U([0, 1)^m)) \le \text{negl}(m)$, and the triangle inequality, we see that

$$\Delta\left(\mathcal{N}(\mathbf{0}, \sigma_2^2\mathbf{I}_m), \mathbf{w}\right) \leq e^{-\ln^2(m)} + \mathsf{negl}(m) = \mathsf{negl}(m).$$

Let $\mathbf{a} = \frac{1}{\sigma_2}\mathbf{w}$. It follows that \mathbf{a} is statistically close to $\mathcal{N}(\mathbf{0}, \mathbf{I}_m)$. Feed \mathbf{a} into the NPP solver to receive some $\mathbf{x} \in \{-1, 1\}^m$ such that $|\mathbf{a}^\top \mathbf{x}| \le \kappa(m) \sqrt{m}$. Let $e = -\mathbf{a}^\top \mathbf{x} \in \mathbb{R}$ with $|e| \le \kappa(m) \sqrt{m}$. Let $x_1 \in \{-1, 1\}$ be the first entry of \mathbf{x} . Let $e' = \sigma_2 e \in \mathbb{R}$, and let $e'' = \mathbf{f}^\top \mathbf{x} + e' \in \mathbb{R}$.

The reduction outputs

$$\mathbf{s} = x_1 \left(\mathbf{U} \mathbf{x} - \mathbf{S} (\mathbf{A} - [\mathbf{A}]_{\mathbf{p}}) \mathbf{x} + \mathbf{S} \widetilde{\varphi}^{-1} (e'') \right).$$

We first argue that $\mathbf{s} \in \mathcal{L}(\mathbf{B})$. Since $\mathbf{a}^{\top}\mathbf{x} + e = 0$, by scaling up, we have $\mathbf{w}^{\top}\mathbf{x} + e' = 0$. Since $\mathbf{x} \in \mathbb{Z}^m$,

$$0 = \mathbf{w}^{\top} \mathbf{x} + e' \equiv \mathbf{y}^{\top} \mathbf{x} + e' = (\widetilde{\varphi}([\mathbf{A}]_{\mathbf{p}}) + \mathbf{f})^{\top} \mathbf{x} + e' \mod 1,$$
$$\equiv \mathbf{c}^{\top} [\mathbf{A}]_{\mathbf{p}} \mathbf{x} + \mathbf{f}^{\top} \mathbf{x} + e' \mod 1,$$
$$\equiv \widetilde{\varphi}([\mathbf{A}]_{\mathbf{p}} \mathbf{x}) + e'' \mod 1.$$

By closure, we know $e'' \in 1/q \cdot \mathbb{Z}/q\mathbb{Z}$, so we have

$$0 \equiv \widetilde{\varphi}([\mathbf{A}]_{\mathbf{p}} \mathbf{x}) + e'' \equiv \widetilde{\varphi}([\mathbf{A}]_{\mathbf{p}} \mathbf{x}) + \widetilde{\varphi}\left(\widetilde{\varphi}^{-1}(e'')\right) \equiv \widetilde{\varphi}\left([\mathbf{A}]_{\mathbf{p}} \mathbf{x} + \widetilde{\varphi}^{-1}(e'')\right) \mod 1.$$

By applying $\widetilde{\varphi}^{-1}$, it follows that

$$\mathbf{A}\mathbf{x} - (\mathbf{A} - \lfloor \mathbf{A} \rfloor_{\mathbf{p}})\mathbf{x} + \widetilde{\varphi}^{-1}(e'') = \lfloor \mathbf{A} \rfloor_{\mathbf{p}} \mathbf{x} + \widetilde{\varphi}^{-1}(e'') \in \mathbb{Z}^n.$$

Plugging in the definition of **A** and since $\mathbf{x} \in \mathbb{Z}^m$,

$$S^{-1}(V+U)x-(A-\lfloor A\rfloor_p)x+\widetilde{\varphi}^{-1}(e'')\in\mathbb{Z}^n.$$

Multiplying by S on the left gives

$$(V + U)x - S(A - \lfloor A \rfloor_p)x + S\widetilde{\varphi}^{-1}(e'') \in \mathcal{L}(S) \subseteq \mathcal{L}(B).$$

Since V contains vectors in $\mathcal{L}(B)$ and $\mathbf{x} \in \mathbb{Z}^m$, we can subtract by $\mathbf{V}\mathbf{x}$ to get

$$Ux - S(A - [A]_p)x + S\widetilde{\varphi}^{-1}(e'') \in \mathcal{L}(B).$$

Multiplying by the sign $x_1 \in \{-1, 1\}$ gives

$$\mathbf{s} = x_1 \left(\mathbf{U} \mathbf{x} - \mathbf{S} (\mathbf{A} - [\mathbf{A}]_{\mathbf{p}}) \mathbf{x} + \mathbf{S} \widetilde{\varphi}^{-1} (e'') \right) \in \mathcal{L}(\mathbf{B}),$$

as desired.

We next argue that the norm of $\mathbf{s} - \mathbf{t}$ is small. Decompose \mathbf{x} as $\mathbf{x}^{\top} = [x_1 || \mathbf{x}_{-1}^{\top}]$, and decompose \mathbf{U} as $\mathbf{U} = [\mathbf{u}_1 || \mathbf{U}_{-1}]$. We then have

$$s = x_{1}Ux - x_{1}S(A - \lfloor A \rfloor_{p})x + x_{1}S\widetilde{\varphi}^{-1}(e'')$$

$$= x_{1}(\mathbf{u}_{1}x_{1} + \mathbf{U}_{-1}\mathbf{x}_{-1}) - x_{1}S(A - \lfloor A \rfloor_{p})x + x_{1}S\widetilde{\varphi}^{-1}(e'')$$

$$= \mathbf{u}_{1} + x_{1}\mathbf{U}_{-1}\mathbf{x}_{-1} - x_{1}S(A - \lfloor A \rfloor_{p})x + x_{1}S\widetilde{\varphi}^{-1}(e'')$$

$$= \mathbf{t} + \mathbf{u}'_{1} + x_{1}\mathbf{U}_{-1}\mathbf{x}_{-1} - x_{1}S(A - \lfloor A \rfloor_{p})x + x_{1}S\widetilde{\varphi}^{-1}(e''),$$

where the distribution of \mathbf{u}_1' is $\mathcal{N}(\mathbf{0}, \sigma_1^2 \mathbf{I}_n)$. It follows that

$$\begin{aligned} \|\mathbf{s} - \mathbf{t}\|_{2} &= \left\| \mathbf{u}_{1}' + x_{1} \mathbf{U}_{-1} \mathbf{x}_{-1} - x_{1} \mathbf{S} (\mathbf{A} - [\mathbf{A}]_{\mathbf{p}}) \mathbf{x} + x_{1} \mathbf{S} \widetilde{\varphi}^{-1} (e'') \right\|_{2} \\ &\leq \left\| \mathbf{u}_{1}' \right\|_{2} + \left\| \mathbf{U}_{-1} \mathbf{x}_{-1} \right\|_{2} + \left\| \mathbf{S} (\mathbf{A} - [\mathbf{A}]_{\mathbf{p}}) \mathbf{x} \right\|_{2} + \left\| \mathbf{S} \widetilde{\varphi}^{-1} (e'') \right\|_{2} \\ &\leq 4 \sigma_{1} m + \left\| \mathbf{S} (\mathbf{A} - [\mathbf{A}]_{\mathbf{p}}) \mathbf{x} \right\|_{2} + \left\| \mathbf{S} \widetilde{\varphi}^{-1} (e'') \right\|_{2} \\ &\leq 4 \sigma_{1} m + \left\| \mathbf{S} \right\| \left(\left\| (\mathbf{A} - [\mathbf{A}]_{\mathbf{p}}) \mathbf{x} \right\|_{1} + \left\| \widetilde{\varphi}^{-1} (e'') \right\|_{1} \right) \end{aligned} \qquad \text{(by Lemma 1)} \\ &\leq 4 \sigma_{1} m + \left\| \mathbf{S} \right\| \left(\frac{n}{\min_{i \in [n]} p_{i}} \left\| \mathbf{x} \right\|_{1} + \left\| \widetilde{\varphi}^{-1} (e'') \right\|_{1} \right) \\ &= r + \left\| \mathbf{S} \right\| \left(\frac{nm}{\min_{i \in [n]} p_{i}} + \left\| \widetilde{\varphi}^{-1} (e'') \right\|_{1} \right). \\ &\leq r + \left\| \mathbf{S} \right\| \left(\frac{1}{16} + \left\| \widetilde{\varphi}^{-1} (e'') \right\|_{1} \right). \end{aligned}$$

It suffices to upper bound $\|\widetilde{\varphi}^{-1}(e'')\|_1$ by 1/16. Recall from Lemma 12 that

$$\widetilde{\varphi}^{-1}(e^{\prime\prime}) = \left(\frac{q}{p_1}e^{\prime\prime}, \frac{q}{p_2}e^{\prime\prime}, \dots, \frac{q}{p_n}e^{\prime\prime}\right) \in \bigoplus_{i \in [n]} 1/p_i \cdot \mathbb{Z}/p_i\mathbb{Z}.$$

For of these entries to be small when viewed in \mathbb{R} , we want to ensure that there's no "wraparound." In particular, if

$$e'' \leq \frac{\min_{i \in [n]} p_i}{16qn},$$

then $\|\widetilde{\varphi}^{-1}(e'')\|_1 \le 1/16$, as desired. Since $\min_{i \in [n]} p_i \ge 32nm$, it suffices to show that

$$e'' \le \frac{2m}{q}.$$

Recall that $e'' = \mathbf{f}^{\top}\mathbf{x} + e' = \mathbf{f}^{\top}\mathbf{x} + \sigma_2 e$, where $\mathbf{f} \sim U([0, 1/q)^m)$ and $|e| \leq \kappa(m)\sqrt{m}$. It follows that

$$|e''| \le |\mathbf{f}^{\mathsf{T}}\mathbf{x}| + |e'| \le ||\mathbf{f}||_{\infty} ||\mathbf{x}||_{1} + \sigma_{2}|e| \le \frac{m}{q} + \sigma_{2} \cdot \kappa(m) \sqrt{m}.$$

Thus, it suffices to show $\sigma_2 \cdot \kappa(m) \sqrt{m} \le m/q$, or equivalently, $\kappa(m) \le \sqrt{m}/(q \ln m)$. We have

$$\frac{\sqrt{m}}{q \ln m} = \frac{2^{5n\frac{1}{1+\varepsilon}}}{q \log \ln(2)n^{\frac{1}{1+\varepsilon}}} \ge \frac{2^{5n\frac{1}{1+\varepsilon}}}{10 \ln(2)(320nm)^n n^{\frac{1}{1+\varepsilon}}}$$

$$= \frac{2^{5n\frac{1}{1+\varepsilon}}}{10 \ln(2)(320n)^n 2^{10n\frac{2+\varepsilon}{1+\varepsilon}} n^{\frac{1}{1+\varepsilon}}}$$

$$\ge \frac{1}{2^{11n^{\frac{2+\varepsilon}{1+\varepsilon}}}}$$

for sufficiently large *n*. On the other hand,

$$\kappa(m) = \frac{1}{2^{(\log m)^{2+\varepsilon}}} = \frac{1}{2^{\left(10n^{\frac{1}{1+\varepsilon}}\right)^{2+\varepsilon}}} = \frac{1}{2^{10^{2+\varepsilon}n^{\frac{2+\varepsilon}{1+\varepsilon}}}} \leq \frac{1}{2^{100n^{\frac{2+\varepsilon}{1+\varepsilon}}}}.$$

Therefore, $\kappa(m) \leq \sqrt{m}/(q \ln m)$, as desired.

Lastly, we note that the NPP_{κ} solver need only succeed with some non-negligible probability $\mu(m)$. As a result, we can repeat this whole process $O(1/\mu) = \text{poly}(m)$ times, and since we can efficiently verify whether the NPP_{κ} solver succeeded, the reduction will still go through.

Moreover, as NPP is a special case of SBP, all of the generalizations and variants discussed in Section 3.2 apply here to NPP as well.

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